

Asymptotic Behaviour of Homozygosity

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Definition

Let $\gamma(t)$ denote the gamma subordinator with Lévy measure

$$\Lambda(dx) = x^{-1}e^{-x}dx, \quad x > 0.$$

For any $\theta > 0$, let $J_1(\theta) \geq J_2(\theta) \geq \dots$ denote the jump sizes of $\gamma(t)$ over the interval $[0, \theta]$ in descending order. If we set $P_i(\theta) = J_i(\theta)/\gamma(\theta)$, $i \geq 1$, then the law of

$$\mathbf{P}(\theta) = (\mathbf{P}_1(\theta), \mathbf{P}_2(\theta), \dots)$$

is Kingman's Poisson-Dirichlet distribution $\mathbf{PD}(\theta)$. It is a probability on the infinite-dimensional simplex

$$\nabla_\infty = \{\mathbf{p} = (\mathbf{p}_1, \mathbf{p}_2, \dots) : \mathbf{p}_1 \geq \mathbf{p}_2 \geq \dots \geq 0, \sum_{i=1}^{\infty} \mathbf{p}_i \leq 1\}.$$

Definition

For any integer $\mathbf{m} \geq 2$, the function

$$\mathbf{H}(\mathbf{p}; \mathbf{m}) = \sum_{i=1}^{\infty} \mathbf{p}_i^{\mathbf{m}}, \quad \mathbf{p} \in \nabla_{\infty}$$

is loosely called the homozygosity of order \mathbf{m} . The name is taken from population genetics where the homozygosity corresponds to $\mathbf{m} = 2$. It represents the probability that all samples are of the same type when a random sample of size \mathbf{m} is selected from the population.

Definition

The function is closely associated with the Shannon entropy in communication, the Herfindahl-Hirschman index in economics, and the Gini-Simpson index in ecology. It provides a measure of concentration of the population in terms of individual types with large values corresponding to higher concentration.

Definition

If the proportions are random, then homozygosity becomes a random variable.

Assume that the proportions of individual types follow distribution $\mathbf{PD}(\theta)$. The focus of this talk will be the random homozygosity

$$\mathbf{H}(\mathbf{P}(\theta); \mathbf{m}).$$

LLN

Question: What is the asymptotic behaviour of $\mathbf{H}(\mathbf{P}(\theta); \mathbf{m})$ when θ tends to infinity?

LLN:

$\mathbf{H}(\mathbf{P}(\theta); \mathbf{m}) \rightarrow 0$ in probability, $\theta \rightarrow \infty$.

$\frac{\theta^{m-1}}{\Gamma(\mathbf{m})} \mathbf{H}(\mathbf{P}(\theta); \mathbf{m}) \rightarrow 1$ in probability, $\theta \rightarrow \infty$.

Gaussian Limit

Theorem (Joyce, Krone and Kurtz (02))

$$\sqrt{\theta} \left[\frac{\theta^{\mathbf{m}-1}}{\Gamma(\mathbf{m})} \mathbf{H}(\mathbf{P}(\theta); \mathbf{m}) - 1 \right] \Rightarrow \mathbf{Z}_{\mathbf{m}}$$

where $\mathbf{Z}_{\mathbf{m}}$ is a normal random variable with mean zero and variance

$$\frac{\Gamma(2\mathbf{m})}{\Gamma^2(\mathbf{m})} - \mathbf{m}^2.$$

Gaussian Limit

$$\mathbf{H}(\mathbf{P}(\theta); \mathbf{m}) \approx \frac{\Gamma(\mathbf{m})}{\theta^{\mathbf{m}-1}} + \frac{\Gamma(\mathbf{m})}{\theta^{\mathbf{m}-1/2}} \mathbf{Z}_{\mathbf{m}}$$

and

$$\frac{\theta^{\mathbf{m}-1}}{\Gamma(\mathbf{m})} \mathbf{H}(\mathbf{P}(\theta); \mathbf{m}) \approx \mathbf{1} + \theta^{-1/2} \mathbf{Z}_{\mathbf{m}}$$

It is natural to investigate more refined structures associated with the limits

$$\mathbf{H}(\mathbf{P}(\theta); \mathbf{m}) \rightarrow \mathbf{0}, \quad \theta \rightarrow \infty$$

and

$$\frac{\theta^{\mathbf{m}-1}}{\Gamma(\mathbf{m})} \mathbf{H}(\mathbf{P}(\theta); \mathbf{m}) \rightarrow \mathbf{1}, \quad \theta \rightarrow \infty.$$

Large Deviations From Zero

Theorem (Dawson and F (06))

The family $\{\mathbf{H}(\mathbf{P}(\theta); \mathbf{m}) : \theta > 0\}$ satisfies a LDP with speed θ and rate function

$$I(\mathbf{y}) = \begin{cases} \log \frac{1}{1-\mathbf{y}^{1/m}}, & \mathbf{y} \in [0, 1] \\ \infty, & \text{else.} \end{cases}$$

Moderate Deviations

Let $\mathbf{a}(\theta)$ satisfy

$$\lim_{\theta \rightarrow \infty} \mathbf{a}(\theta) = \infty, \quad \lim_{\theta \rightarrow \infty} \frac{\mathbf{a}(\theta)}{\sqrt{\theta}} = 0,$$

and

$$\liminf_{\theta \rightarrow \infty} \frac{\mathbf{a}^{1-\epsilon}(\theta)}{\theta^{(m-1)/(2m-1)}} > 0$$

for some ϵ in $(0, \frac{1}{2m-1})$.

Moderate Deviations

Theorem (Gao and F (08))

The family $\mathbf{a}(\theta) \left(\frac{\theta^{m-1}}{\Gamma(\mathbf{m})} \mathbf{H}(\mathbf{P}(\theta), \mathbf{m}) - 1 \right)$ satisfies a LDP with speed $\frac{\mathbf{a}^2(\theta)}{\theta}$ and rate function $\frac{\mathbf{x}^2}{2(\Gamma(2\mathbf{m})/\Gamma(\mathbf{m})^2 - \mathbf{m}^2)}$, $\mathbf{x} \in \mathbf{R}$.

Remark

Let $\mathbf{a}(\theta) = \theta^\delta$. Then moderate deviation holds for

$$\theta^\delta \left(\frac{\theta^{\mathbf{m}-1}}{\Gamma(\mathbf{m})} \mathbf{H}(\mathbf{P}(\theta), \mathbf{m}) - 1 \right)$$

if and only if $\delta \in \left(\frac{\mathbf{m}-1}{2\mathbf{m}-1}, \frac{1}{2} \right)$.

This indicates a significant departure from the Gaussian regime when δ is between 0 and $\frac{\mathbf{m}-1}{2\mathbf{m}-1}$.

Large Deviations From One

The case $\delta = 0$ corresponds to the large deviations of

$$\frac{\theta^{\mathbf{m}-1}}{\Gamma(\mathbf{m})} \mathbf{H}(\mathbf{P}(\theta), \mathbf{m})$$

from one.

Fundamental Differences From LDP for $\mathbf{H}(\mathbf{P}(\theta), \mathbf{m})$

- The state space is no longer compact
- Exponential tightness is not free
- Do not have exponential moment in the neighbourhood of zero

Large Deviations From One

Theorem (Dawson and F(16))

A large deviation principle holds for $\frac{\theta^{m-1}}{\Gamma(\mathbf{m})} \mathbf{H}(\mathbf{P}(\theta); \mathbf{m})$ as θ converges to infinity on space \mathbf{R} with speed $\theta^{1/m}$ and good rate function

$$\mathbf{S}(\mathbf{x}) = \begin{cases} [\Gamma(\mathbf{m})(\mathbf{x} - 1)]^{1/m}, & \mathbf{x} \geq 1, \\ +\infty, & \text{otherwise.} \end{cases}$$

Note: The scale of deviations for $\mathbf{x} < 1$ is different from that of $\mathbf{x} > 1$.

Link Between the Two Rate Functions

Question: Can one derive the LDP for $\frac{\theta^{m-1}}{\Gamma(\mathbf{m})} \mathbf{H}(\mathbf{P}(\theta); \mathbf{m})$ from the LDP for $\mathbf{H}(\mathbf{P}(\theta); \mathbf{m})$ or vice versa?

Answer: ??

Link Between the Two Rate Functions

Recall that the LDP for $\mathbf{H}(\mathbf{P}(\theta); \mathbf{m})$ has speed θ and rate function

$$\mathbf{I}(\mathbf{y}) = \begin{cases} -\log(1 - \mathbf{y}^{1/\mathbf{m}}), & \mathbf{y} \in [0, 1] \\ \infty, & \text{otherwise} \end{cases}$$

Since $\mathbf{H}(\mathbf{P}(\theta); \mathbf{m})$ and $\mathbf{H}(\mathbf{P}(\theta); \mathbf{m}) - \frac{\Gamma(\mathbf{m})}{\theta^{\mathbf{m}-1}}$ are exponentially equivalent, the same LDP holds for $\mathbf{H}(\mathbf{P}(\theta); \mathbf{m}) - \frac{\Gamma(\mathbf{m})}{\theta^{\mathbf{m}-1}}$.

Link Between the Two Rate Functions

Write $\frac{\theta^{m-1}}{\Gamma(\mathbf{m})} \mathbf{H}(\mathbf{P}(\theta); \mathbf{m})$ as

$$\frac{\theta^{m-1}}{\Gamma(\mathbf{m})} \left[\mathbf{H}(\mathbf{P}(\theta); \mathbf{m}) - \frac{\Gamma(\mathbf{m})}{\theta^{m-1}} \right] + 1.$$

For $\mathbf{x} \in [1, \infty)$ and $\frac{\theta^{m-1}}{\Gamma(\mathbf{m})} \mathbf{H}(\mathbf{P}(\theta); \mathbf{m}) = \mathbf{x}$, let $\mathbf{y} = \frac{\Gamma(\mathbf{m})}{\theta^{m-1}} (\mathbf{x} - 1)$. Then

$$\begin{aligned} \exp\{-\theta \mathbf{I}(\mathbf{y})\} &= \exp\left\{-\theta^{1/\mathbf{m} + \mathbf{m}/(\mathbf{m}-1)} \log \frac{1}{1 - \left(\frac{\Gamma(\mathbf{m})}{\theta^{m-1}} (\mathbf{x} - 1)\right)^{1/\mathbf{m}}}\right\} \\ &\approx \exp\{-\theta^{1/\mathbf{m}} \mathbf{S}(\mathbf{x})\}. \end{aligned}$$

Main Steps of Proof

Step 1

Showing that LDP for general θ is equivalent to θ being integers.

Step 2

For integer θ , find a new representation of $\frac{\theta^{m-1}}{\Gamma(\mathbf{m})} \mathbf{H}(\mathbf{P}(\theta); \mathbf{m})$ as

$$\frac{\theta^{m-1}}{\Gamma(\mathbf{m})} \mathbf{H}(\mathbf{P}(\theta); \mathbf{m}) = \frac{\theta^{m-1}}{\Gamma(\mathbf{m})} \left[\frac{1}{\gamma^{\mathbf{m}}(\theta)} \sum_{k=1}^{\theta} \mathbf{W}_k^{\mathbf{m}} \mathbf{H}_k \right]$$

where $\mathbf{W}_1, \dots, \mathbf{W}_\theta$ are independent copies of $\gamma(1)$, and **independently**, $\mathbf{H}_1, \dots, \mathbf{H}_\theta$ are independent copies of $\mathbf{H}(\mathbf{P}(1); \mathbf{m})$.

Main Steps of Proof

Step 3

Exploring the independence and the LDP for gamma distribution to verify that the LDP for $\frac{\theta^{m-1}}{\Gamma(\mathbf{m})} \mathbf{H}(\mathbf{P}(\theta); \mathbf{m})$ is equivalent to the LDP for

$$\frac{1}{\Gamma(\mathbf{m})\theta} \sum_{k=1}^{\theta} \mathbf{W}_k^m \mathbf{H}_k$$

Step 4

Applying Cramér's theorem for $\mathbf{x} < 1$.








Step 5

Applying Nagaev's result for $\mathbf{x} > 1$.

Generalizations

What about other random distributions?

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Thanks!