

Multivariate Regular Variation of In- and Out-Degree in a Network Growth Model

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Joint: [G. Samorodnitsky](#), [T. Wang](#), [P. Wan](#), [A. Willis](#), [D. Towsley](#),
[R. Davis](#)

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1. Growing preferential attachment networks

See Bollobás, Borgs, Chayes, and Riordan (2003) and Krapivsky and Redner (2001).

1.1. Model description

- Model parameters: $\alpha, \beta, \gamma, \delta_{\text{in}}, \delta_{\text{out}}$ with $\alpha + \beta + \gamma = 1$.
- $G(n)$ is a directed random graph with n edges, $N(n)$ nodes.
- Set of nodes of $G(n)$ is V_n ; so $|V_n| = N(n)$.
- Set of edges of $G(n)$ is $E_n = \{(u, v) \in V_n \times V_n : (u, v) \in E_n\}$.
- In-degree of v is $D_{\text{in}}(v)$; out-degree of v is D_{out} . Dependence on n is suppressed.
- Obtain graph $G(n)$ from $G(n-1)$ in a Markovian way as follows:



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1. With probability α , append to $G(n-1)$ a new node $v \notin V_{n-1}$ and create directed edge $v \mapsto w \in V_{n-1}$ with probability

$$\frac{D_{\text{in}}(w) + \delta_{\text{in}}}{n-1 + \delta_{\text{in}}N(n-1)}.$$

2. With probability γ , append to $G(n-1)$ a new node $v \notin V_{n-1}$ and create directed edge $w \in V_{n-1} \mapsto v \notin V_{n-1}$ with probability

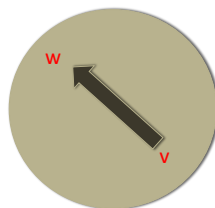
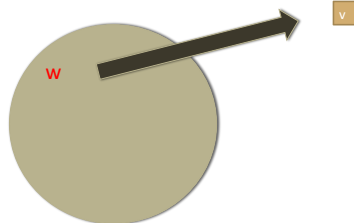
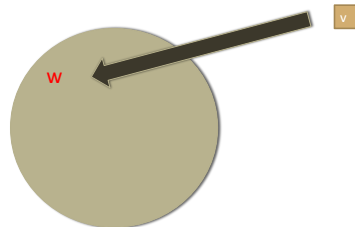
$$\frac{D_{\text{out}}(w) + \delta_{\text{out}}}{n-1 + \delta_{\text{out}}N(n-1)}.$$

3. With probability β , create new directed edge between existing nodes

$$v \in V_{n-1} \mapsto w \in V_{n-1}$$

with probability

$$\left(\frac{D_{\text{out}}(v) + \delta_{\text{out}}}{n-1 + \delta_{\text{out}}N(n-1)} \right) \left(\frac{D_{\text{in}}(w) + \delta_{\text{in}}}{n-1 + \delta_{\text{in}}N(n-1)} \right)$$





1.2. Background: What's known.

Notation:

$$N(n) = \# \text{ nodes in } V_n.$$

$$n = \# \text{ edges in } E_n.$$

$$N_{ij}(n) = \# \text{ nodes with in-degree}=i \text{ and out-degree}=j \text{ in } G(n).$$

Then (eg, [Bollobás, Borgs, Chayes, and Riordan \(2003\)](#)) the limiting proportion of nodes with in-degree= i and out-degree= j is

$$\lim_{n \rightarrow \infty} \frac{N_{ij}(n)}{N(n)} = p_{ij} = \text{a prob mass function.}$$

Marginally, the limiting degree frequency (p_{ij}) has power-law tails: For some finite positive constants C_{in} and C_{out} ,

$$p_i(\text{in}) := \sum_{j=0}^{\infty} p_{ij} \sim C_{in} i^{-\alpha_{in}} \quad \text{as } i \rightarrow \infty, \text{ as long as } \alpha \delta_{in} + \gamma > 0,$$

$$p_j(\text{out}) := \sum_{i=0}^{\infty} p_{ij} \sim C_{out} j^{-\alpha_{out}} \quad \text{as } j \rightarrow \infty, \text{ as long as } \gamma \delta_{out} + \alpha > 0,$$

where

$$\alpha_{in} = 1 + \frac{1 + \delta_{in}(\alpha + \gamma)}{\alpha + \beta}, \quad \alpha_{out} = 1 + \frac{1 + \delta_{out}(\alpha + \gamma)}{\gamma + \beta}.$$

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Conclude that:

- $\alpha_{in} > 1$, $\alpha_{out} > 1$, and
- Manufacture random pair

$$(I, 0) \sim \{p_{ij}\},$$

and then

$$P[I = i] \sim C_{in} i^{-\alpha_{in}}, \quad i \rightarrow \infty; \quad P[O = j] \sim C_{out} j^{-\alpha_{out}}, \quad j \rightarrow \infty.$$

- So, as $x \rightarrow \infty$,

$$P[I > x] \sim k_{in} x^{-(\alpha_{in}-1)}, \quad P[O > x] \sim k_{out} x^{-(\alpha_{out}-1)}.$$

- Statisticians note: First we let $n \rightarrow \infty$ and then $x \rightarrow \infty$. Data usually node based without letting $n \rightarrow \infty$.

Question: Is (I, O) jointly regularly varying? (Hint: Yes.)

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1.3. Methods of attack.

1. Full frontal assault: Prove directly $P[(I, O) \in \cdot]$ is a regularly varying measure. (Samorodnitsky, Resnick, Towsley, Davis, Willis, and Wan, 2016).
2. Sneaky Resnicki: Multivariate Tauberian theorem after computing the generating function. (Resnick and Samorodnitsky, 2015)
3. Analysis of the mass function $\{p_{ij}\}$. (Wang and Resnick, 2016)



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1.4. The Generating Function of $(I, O) \sim \{p_{ij}\}$.

- Because of the Markovian nature of the model construction, p_{ij} satisfies difference relation in (i, j) . (Bollobás, Borgs, Chayes, and Riordan, 2003).
- Solve the difference equations using pde methods (Samorodnitsky, Resnick, Towsley, Davis, Willis, and Wan, 2016) to get explicit formula for

$$\varphi(x, y) = E\left(x^I y^O\right) = \sum_{i=0}^{\infty} \sum_{j=0}^{\infty} x^i y^j p_{ij}, \quad 0 \leq x, y \leq 1, \quad (1)$$

where

$$\varphi(x, y) = \frac{\gamma}{\alpha + \gamma} x \varphi_1(x, y) + \frac{\alpha}{\alpha + \gamma} y \varphi_2(x, y), \quad (2)$$

with (constants c_1, c_2, a are functions of $\alpha_{in}, \alpha_{out}$)

$$\varphi_1(x, y) = c_1^{-1} \int_1^{\infty} z^{-(1+1/c_1)} (x + (1-x)z)^{-\delta_{in}+1} (y + (1-y)z^a)^{-\delta_{out}} dz, \quad (3)$$

$$\varphi_2(x, y) = c_1^{-1} \int_1^{\infty} z^{-(1+1/c_1)} (x + (1-x)z)^{-\delta_{in}} (y + (1-y)z^a)^{-\delta_{out}+1} dz \quad (4)$$

for $0 \leq x, y \leq 1$. Further:

$$c_1 = \frac{1}{\alpha_{in} - 1}, \quad c_2 = \frac{1}{\alpha_{out} - 1}, \quad a = c_2/c_1.$$

- Hmmmmmm... !
- Analyze the generating functions φ_1, φ_2 separately.



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2. Skip the blood and guts! What's the answer?

Theorem. The random vector (I, O) with joint mass function $\{p_{ij}\}$ satisfies as $t \rightarrow \infty$,

$$tP \left[\left(\frac{I}{t^{1/(\alpha_{in}-1)}}, \frac{O}{t^{1/(\alpha_{out}-1)}} \right) \in \cdot \right] \xrightarrow{v} \frac{\gamma}{\alpha + \gamma} \nu_1(\cdot) + \frac{\alpha}{\alpha + \gamma} \nu_2(\cdot),$$

vaguely in $M_+([0, \infty]^2 \setminus \{\mathbf{0}\})$ and ν_1 and ν_2 concentrate on $(0, \infty)^2$ and have Lebesgue densities f_1, f_2 given by,

$$f_1(x, y) = c_1^{-1} (\Gamma(\delta_{in} + 1) \Gamma(\delta_{out}))^{-1} x^{\delta_{in}} y^{\delta_{out}-1} \\ \times \int_0^\infty z^{-(2+1/c_1+\delta_{in}+a\delta_{out})} e^{-(x/z+y/z^a)} dz,$$

and

$$f_2(x, y) = c_1^{-1} (\Gamma(\delta_{in}) \Gamma(\delta_{out} + 1))^{-1} x^{\delta_{in}-1} y^{\delta_{out}} \\ \times \int_0^\infty z^{-(1+a+1/c_1+\delta_{in}+a\delta_{out})} e^{-(x/z+y/z^a)} dz.$$

3. Full frontal direct assault.

Samorodnitsky, Resnick, Towsley, Davis, Willis, and Wan (2016)

From the form of the generating function:

- The pair (I, O) has representation

$$(I, O) \stackrel{d}{=} B(1 + X_1, Y_1) + (1 - B)(X_2, 1 + Y_2),$$

where B is a Bernoulli switching variable independent of X_j, Y_j , $j = 1, 2$ with

$$P(B = 1) = 1 - P(B = 0) = \frac{\gamma}{\alpha + \gamma}.$$

and

$$(X_1, Y_1) \sim \varphi_1 \quad \text{and} \quad (X_2, Y_2) \sim \varphi_2.$$



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- Form of φ_i $i = 1, 2$:
 - The negative binomial distribution is id;
 - Suppose $\{T_\lambda(p), p \in (0, 1)\}$ and $\{\tilde{T}_\mu(p), p \in (0, 1)\}$ are two independent families of negative binomial random variables for any choice of λ, μ .
 - Write $X_j, Y_j, j = 1, 2$

$$(X_1, Y_1) = (T_{\delta_{in}+1}(Z^{-1}), \tilde{T}_{\delta_{out}}(Z^{-a})),$$

$$(X_2, Y_2) = (T_{\delta_{in}}(Z^{-1}), \tilde{T}_{\delta_{out}+1}(Z^{-a})),$$

where

- Z is Pareto on $[1, \infty)$ with index c_1^{-1} , independent of the negative binomial random variables.
- Prove multivariate regular variation of $(X_1, Y_i), i = 1, 2$ is inherited from the Pareto Z . [But the negative binomials smear the limit measure mass over whole first quadrant.]

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4. Sneaky: Tauberian method

Resnick and Samorodnitsky (2015)

- Abel-Tauberian theorems relate power law behavior of distributions and their transforms in \mathbb{R}_+^p .

Antecedents:

- $p = 1$, Bingham, Goldie, and Teugels (1987), Feller (1971), Karamata (1931).
- $p > 1$ for standard regular variation: Resnick (1991, 2007), Stadtmüller (1981), Stadtmüller and Trautner (1979, 1981), Stam (1977), Yakimiv (2005).
- Assume $U \in M_+(\mathbb{R}_+^p)$ with distribution function $U(\mathbf{x}) = U([\mathbf{0}, \mathbf{x}])$ and the Laplace transform $\hat{U}(\boldsymbol{\lambda})$ of U exists:

$$\hat{U}(\boldsymbol{\lambda}) = \int_{\mathbb{R}_+^p} \exp\left\{-\sum_{i=1}^p \lambda_i x_i\right\} U(d\mathbf{x}) < \infty, \quad \boldsymbol{\lambda} \in \mathbb{R}_+^p.$$



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4.1. First 1/2 Taub

Assume

1. $\exists b_i \in RV_{1/\gamma_i}$, $\gamma_i > 0$, $i = 1, \dots, p$; set

$$\mathbf{b}(t) = (b_1(t), \dots, b_p(t)).$$

2. U is regularly varying **infinite** measure with limit measure U_∞ on \mathbb{R}_+^p :

$$U_t(\cdot) := \frac{1}{t}U(\mathbf{b}(t)\cdot) \xrightarrow{v} U_\infty(\cdot).$$

3. Regularity condition:

$$\lim_{y \rightarrow \infty} \limsup_{t \rightarrow \infty} \int_{\cup_{i=1}^p [v_i > y]} e^{-\sum_{i=1}^p v_i/x_i} U_t(d\mathbf{v}) = 0. \quad ([UR])$$

Then

1. The Laplace transforms $\hat{U}(\mathbf{1}/\mathbf{x})$ and $\hat{U}_\infty(\mathbf{1}/\mathbf{x})$ are distribution functions of Radon measures on \mathbb{R}_+^p and
2. First 1/2 Taub: the measure corresponding to $\hat{U}(\mathbf{1}/\mathbf{x})$ inherits non-standard regular variation: for $\mathbf{x} > 0$

$$\frac{1}{t} \hat{U}\left(\frac{\mathbf{1}}{\mathbf{b}(t)\mathbf{x}}\right) \rightarrow \hat{U}_\infty\left(\frac{\mathbf{1}}{\mathbf{x}}\right), \quad (5)$$

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4.1.1. Proof sketch of First 1/2 Taub; map your way to happiness.

(i) Assume U is regularly varying:

$$U_t(\cdot) := \frac{1}{t}U(\mathbf{b}(t)\cdot) \xrightarrow{v} U_\infty(\cdot).$$

(ii) Set $\mathcal{F} = (\frac{1}{E_1}, \dots, \frac{1}{E_p})$, where (E_1, \dots, E_p) are iid exp rv's so $1/E_i$ is unit Frechet. In $M_+([0, \infty]^p \times \mathbb{R}_+^p)$

$$P[\mathcal{F} \in \cdot] \times U_t \xrightarrow{v} P[\mathcal{F} \in \cdot] \times U_\infty.$$

(iii) Define $h : [0, \infty]^p \times \mathbb{R}_+^p \mapsto [0, \infty]^p \times \mathbb{R}_+^p$ by

$$h(\mathbf{x}, \mathbf{y}) = (\mathbf{x}\mathbf{y}, \mathbf{y}).$$

(iv) By a continuity theorem for convergence of measures:

$$\begin{aligned} & (P[\mathcal{F} \in \cdot] \times U_t) \circ h^{-1}([\mathbf{0}, \mathbf{x}] \times [\mathbf{0}, \mathbf{y}\mathbf{1}]) \\ & \rightarrow (P[\mathcal{F} \in \cdot] \times U_\infty) \circ h^{-1}([\mathbf{0}, \mathbf{x}] \times [\mathbf{0}, \mathbf{y}\mathbf{1}]). \end{aligned}$$

(v) Unpack: As $t \rightarrow \infty$,

$$\int_{\mathbf{v} \leq \mathbf{y}\mathbf{1}} e^{-\sum_{i=1}^p v_i/x_i} U_t(d\mathbf{v}) \rightarrow \int_{\mathbf{v} \leq \mathbf{y}\mathbf{1}} e^{-\sum_{i=1}^p v_i/x_i} U_\infty(d\mathbf{v}).$$

(vi) Let $y \rightarrow \infty$ via condition [UR]. Done!

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4.2. Second 1/2 Taub.

Assume

1. The infinite measure $U \in M_+(\mathbb{R}_+^p)$ has distribution function $U(\mathbf{x}) = U([\mathbf{0}, \mathbf{x}])$ and the Laplace transform $\hat{U}(\boldsymbol{\lambda})$ of U exists:

$$\hat{U}(\boldsymbol{\lambda}) = \int_{\mathbb{R}_+^p} \exp\{-\boldsymbol{\lambda}'\mathbf{x}\} U(d\mathbf{x}) < \infty, \quad \boldsymbol{\lambda} \in \mathbb{R}_+^p.$$

2. $b_i \in RV_{1/\gamma_i}$, $\gamma_i > 0$ for $i = 1, \dots, p$.
3. Condition **[UR]** holds:

$$\lim_{y \rightarrow \infty} \limsup_{t \rightarrow \infty} \int_{\cup_{i=1}^p [v_i > y]} e^{-\sum_{i=1}^p v_i/x_i} U_t(d\mathbf{v}) = 0. \quad ([UR])$$

4. There exists a finite-valued function \hat{U}_∞ such that for $\mathbf{x} > \mathbf{0}$,

$$\frac{1}{t} \hat{U}\left(\frac{\mathbf{1}}{\mathbf{b}(t)\mathbf{x}}\right) = \frac{1}{t} \hat{U}\left(\frac{1}{b_1(t)x_1}, \dots, \frac{1}{b_p(t)x_p}\right) \rightarrow \hat{U}_\infty(\mathbf{1}/\mathbf{x}). \quad (6)$$

Then for some measure $U_\infty \in M_+(\mathbb{R}_+^p)$ whose Laplace transform is \hat{U}_∞ , we have as $t \rightarrow \infty$,

$$U_t(\cdot) = \frac{1}{t} U(\mathbf{b}(t)\cdot) \rightarrow U_\infty(\cdot), \quad \text{in } M_+(\mathbb{R}_+^p).$$

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4.3. Apply to random graph.

- To apply Tauberian theory, need infinite measure: So for φ_1 , fix $k > \alpha_{in} - 1$ and set

$$\psi(x, y) = \frac{\partial^k \varphi_1}{\partial x^k}(x, y) = \sum_{i=0}^{\infty} \sum_{j=0}^{\infty} x^i y^j m_{ij}^{(k)}$$

where the quantities $m_{ij}^{(k)}$ can be calculated in terms of p_{ij} 's.

- Define

$$U(\cdot) = \sum_{i,j} m_{ij}^{(k)} \epsilon_{(i,j)}(\cdot)$$

as an infinite Radon measure on \mathbb{R}_+^2 concentrating on $(\{0, 1, 2, \dots\})^2$ that puts mass $m_{ij}^{(k)}$ at (i, j) .

- Compute \hat{U} , verify [UR], verify $\hat{U}(\mathbf{1}/\mathbf{x})$ is regularly varying.
- Voila!

5. Analysis of the mass function $\{p_{ij}\}$.

Wang and Resnick (2016)

Suppose $U(\cdot)$ is a measure on \mathbb{R}_+^2 with [pdf, pmf] $[f(x, y), f(i, j)]$.

- If U is a regularly varying measure, is f a regularly varying [function, array]?

Not always true even in one dimension.

- If f is a regularly varying [function, array], is U a regularly varying measure.

Shock! Not necessarily and a general analogue of Karamata's theorem on integration fails.

- Definition of regularly varying array?



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5.1. Definitions: Regularly varying array.

- A doubly indexed function $f : \mathbb{Z}_+^2 \setminus \{\mathbf{0}\} \mapsto \mathbb{R}_+$ is regularly varying with scaling functions b_1 and b_2 and limit function $\lambda(x, y)$ if for some $h \in RV_\alpha$ for some $\alpha \in \mathbb{R}$, $b_i \in RV_{\beta_i}$, $\beta_i > 0$, we have

$$\lim_{n \rightarrow \infty} \frac{f([b_1(n)x], [b_2(n)y])}{h(n)} = \lambda(x, y) > 0, \quad \forall x, y > 0.$$

- A function $f : \mathbb{R}_+^2 \mapsto \mathbb{R}_+$ is regularly varying if the same limit holds without the square brackets $[\cdot], [\cdot]$.
- If $f(i, j)$ is regularly varying, it is embeddable if \exists regularly varying $g(x, y)$ and

$$g(x, y) = f([x], [y]).$$

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5.2. Typical result.

Suppose $u(i, j) > 0$ satisfies

1. $u(i, j)$ is regularly varying.
2. u satisfies some **extra condition**.

Then

1. The function

$$g(x, y) := u([x], [y])$$

is regularly varying as function of continuous variables and $u(i, j)$ is embeddable.

2. If $u(i, j) = p(i, j)$ is a pmf corresponding to (X, Y) , then

$$P[(X, Y) \in \cdot]$$

is a regularly varying measure.

Example of **extra condition**: Easiest is to suppose $u(i, j)$ is eventually decreasing.

BUT: This does not hold for preferential attachment problem.



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5.2.1. More flexible for the non-standard case.

Suppose

- $h(\cdot) \in RV_\rho$, $\rho < 0$, and $u : \mathbb{Z}_+^2 \mapsto \mathbb{R}_+$,
- Scaling functions are power laws: $b_i(t) = t^{1/\alpha_i}$, $i = 1, 2$. [The non-standard case is harder than the standard and we had to suppose b_i are power functions.]
- There exists a limit function $\lambda_0 > 0$ defined on

$$\mathcal{E}_0 := \{(x, y) : \|(x^{\alpha_1}, y^{\alpha_2})\| = 1\},$$

such that u satisfies

$$\lim_{t \rightarrow \infty} \frac{u([t^{1/\alpha_1}x], [t^{1/\alpha_2}y])}{h(t)} = \lambda_0(x, y), \quad \forall (x, y) \in \mathcal{E}_0. \quad (7)$$

Then

1. The doubly indexed function $u(i, j)$ is regularly varying: For all $x, y > 0$, define $\mathbf{w} = \mathbf{w}(x, y) := (x^{\alpha_1}, y^{\alpha_2})$ and

$$\lim_{n \rightarrow \infty} \frac{u([n^{1/\alpha_1}x], [n^{1/\alpha_2}y])}{h(n)} = \lambda(x, y) := \lambda_0\left(\frac{x}{\|\mathbf{w}\|^{1/\alpha_1}}, \frac{y}{\|\mathbf{w}\|^{1/\alpha_2}}\right) \|\mathbf{w}\|^\rho;$$

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- The doubly indexed function $u(i, j)$ is embeddable in a non-standard regularly varying function $f : \mathbb{R}_+^2 \mapsto \mathbb{R}$ with limit function $\lambda(\cdot)$ such that $f(x, y) = u([x], [y])$;
- If convergence in (7) is uniform on \mathcal{E}_0 , then also the measure corresponding to $u(i, j)$ is a (discretely supported) regularly varying measure.

5.3. Back to preferential attachment.

Recall the representation of

$$(I, O) \sim p_{i,j}.$$

$$(I, O) \stackrel{d}{=} B(1 + X_1, Y_1) + (1 - B)(X_2, 1 + Y_2),$$

where for Pareto Z :

$$(X_1, Y_1) = (T_{\delta_{in}+1}(Z^{-1}), \tilde{T}_{\delta_{out}}(Z^{-a})),$$

$$(X_2, Y_2) = (T_{\delta_{in}}(Z^{-1}), \tilde{T}_{\delta_{out}+1}(Z^{-a})),$$

From the representations:

$$\begin{aligned} \lim_{n \rightarrow \infty} \frac{p([n^{c_1}x], [n^{c_2}y])}{n^{-(1+c_1+c_2)}} &= \frac{\gamma}{\alpha + \gamma} \frac{c_1 \Gamma(\delta_{in} + 1) \Gamma(\delta_{out})}{c_1 \Gamma(\delta_{in} + 1) \Gamma(\delta_{out})} \int_0^\infty z^{-(2+1/c_1+\delta_{in}+a\delta_{out})} e^{-\left(\frac{x}{z} + \frac{y}{z^a}\right)} dz \\ &+ \frac{\alpha}{\alpha + \gamma} \frac{c_1 \Gamma(\delta_{in}) \Gamma(\delta_{out} + 1)}{c_1 \Gamma(\delta_{in}) \Gamma(\delta_{out} + 1)} \int_0^\infty z^{-(1+a+1/c_1+\lambda+a\delta_{out})} e^{-\left(\frac{x}{z} + \frac{y}{z^a}\right)} dz. \end{aligned}$$

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Remarks:

1. This convergence can be shown to be uniform on \mathcal{E}_0 .
2. Therefore, the uniform convergence implies

$$P[(I, O) \in \cdot]$$

is a regularly varying measure.

3. This closes the loops.



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6. Conclusion

- Reciprocity? Cliques? Neighborhoods?
- Inference based on
 - Tail empirical measure based on degree data—study extreme values.
 - Asymptotic normality of degree counts for undirected and directed graphs. Inference for central values. Progress:
 - * Undirected case: [Resnick and Samorodnitsky \(2016\)](#).
 - * Directed case: [Wang and Resnick \(2015\)](#).
- Embedding techniques in birth-death processes may illuminate properties of this model and other models.

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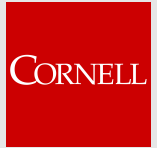
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