

# Conditional moment representations for dependent random variables

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## Abstract

The question considered in this paper is which sequences of  $p$ -integrable random variables can be represented as conditional expectations of a fixed random variable with respect to a given sequence of  $\sigma$ -fields. For finite families of  $\sigma$ -fields, explicit inequality equivalent to solvability is stated; sufficient conditions are given for finite and infinite families of  $\sigma$ -fields, and explicit expansions are presented.

## 1 Introduction

We analyze which sequences of random variables  $\{X_j\}$  can be represented as conditional expectations

$$E(Z|\mathcal{F}_j) = X_j. \quad (1)$$

of a  $p$ -integrable random variable  $Z$  with respect to a given sequence  $(\mathcal{F}_j)$  of  $\sigma$ -fields. The martingale theory answers this question for increasing  $\sigma$ -fields  $(\mathcal{F}_j)$ . We are more interested in other cases which include  $\sigma$ -fields generated by single independent, or say, Markov dependent, random variables. In particular, given a random sequence  $\xi_j$  and  $p$ -integrable random variables  $X_j = f_j(\xi_j)$ , we analyze when there is  $Z \in L_p$  such that

$$X_j = E(Z|\xi_j). \quad (2)$$

This is motivated by our previous results for independent random variables and by the alternating conditional expectations (ACE) algorithm of Breiman & Friedman [3]. In [3] the authors are interested in the  $L_2$ -best additive prediction  $Z$  of a random variable  $Y$  based on the finite number of the predictor variables  $\xi_1, \dots, \xi_d$ . The solution (ACE) is based on the fact that the best additive predictor  $Z = \phi_1(\xi_1) + \dots + \phi_d(\xi_d)$  satisfies the conditional moment constraints (2).

Relation (1) defines an inverse problem, and shares many characteristics of other inverse problems, c. f. Groetsch [8]. Accordingly, our methods partially rely on (non-constructive) functional analysis. We give sufficient conditions for the solvability of (1) in terms of *maximal correlations*. We also show that (2) has solution for finite  $d < \infty$ , if the joint density of  $\xi_1, \dots, \xi_d$  with respect to the product of marginals is bounded away from zero and  $EX_i = EX_j$ .

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We are interested in both finite, and infinite sequences, extending our previous results in [4, 5]. In this paper we concentrate on the  $p$ -integrable case with  $1 < p < \infty$ . The extremes  $p = 1$  or  $p = \infty$  seem to require different assumptions. For infinite sequences of independent r.v. all three cases  $1 < p < \infty$ ,  $p = 1$ , and  $p = \infty$  are completely solved in [5]. For finite sequences of dependent  $\sigma$ -fields, Kellerer [9] and Strassen [15] can be quoted in connection with conditional expectations problem with  $Z \geq 0$ , which can be modified to cover bounded random variables ( $p = \infty$ ) case. For pairs of  $\sigma$ -fields the case  $1 < p < \infty$  is solved in [4].

## 2 Notation and results

For  $2 \leq d \leq \infty$ , let  $\{\mathcal{F}_j\}_{1 \leq j \leq d}$  be a given family of  $\sigma$ -fields. By  $L_p^0(\mathcal{F})$  we denote the Banach space of all  $p$ -integrable  $\mathcal{F}$ -measurable centered random variables,  $1 \leq p \leq \infty$ . By  $\mathcal{E}_j$  we denote the conditional expectation with respect to  $\mathcal{F}_j$ . For  $d < \infty$  by  $\bigoplus_{j=1}^d L_p(\mathcal{F}_j)$  we denote the set of sums  $Z = Z_1 + \dots + Z_d$ , where  $Z_j \in L_p(\mathcal{F}_j)$ .

We shall analyze the following problems.

- For all *consistent*  $X_j \in L_p$  find  $Z \in L_p$  satisfying (1) and with minimal norm

$$E|Z|^p = \min. \quad (3)$$

- For all *consistent*  $X_j \in L_p$  find additive  $Z \in L_p$  satisfying (1); additive means that

$$Z = \sum_{j=1}^d Z_j, \text{ where } Z_j \in L_p(\mathcal{F}_j). \quad (4)$$

(for  $d = \infty$  the series in (4) is assumed to converge absolutely in  $L_p$ )

The above statements do not spell out the consistency conditions; these will be explicit in the theorems.

**Remark 2.1** *If (1) can be solved, then there is a minimal solution  $Z$ . This can be easily recovered from the Komlos law of large numbers [10].*

### 2.1 Maximal correlations

Maximal correlation coefficients play a prominent role below; for another use see also [3, Section 5]. The following maximal correlation coefficient is defined in [4].

Let

$$\tilde{\rho}(\mathcal{F}, \mathcal{G}) = \sup \{ \text{corr}(X, Y) : X \in L_2(\mathcal{F}), Y \in L_2(\mathcal{G}), E(X|\mathcal{F} \cap \mathcal{G}) = 0 \}.$$

Notice that  $\tilde{\rho}(\mathcal{F}, \mathcal{G}) = 0$  for independent  $\mathcal{F}, \mathcal{G}$  but also for increasing  $\sigma$ -fields  $\mathcal{F} \subset \mathcal{G}$ . If the intersection  $\mathcal{F} \cap \mathcal{G}$  is trivial,  $\tilde{\rho}$  coincides with the usual *maximal correlation coefficient*, defined in general by

$$\rho(\mathcal{F}, \mathcal{G}) = \sup_{X \in L_2(\mathcal{F}), Y \in L_2(\mathcal{G})} \text{corr}(X, Y). \quad (5)$$

Note that if  $\rho(\mathcal{F}, \mathcal{G}) < 1$  then  $\mathcal{F} \cap \mathcal{G}$  is trivial.

Given  $d \leq \infty$ ,  $\sigma$ -fields  $\{\mathcal{F}_j\}_{j \leq d}$ , and a finite subset  $T \subset I := \{1, 2, \dots, d\}$  put

$$\mathcal{F}_T = \sigma(\mathcal{F}_j : j \in T).$$

Define *pairwise maximal correlation*  $r$  by

$$r = \sup_{i \neq j} \rho(\mathcal{F}_i, \mathcal{F}_j)$$

and global maximal correlation

$$R = \sup_{T \cap S = \emptyset} \rho(\mathcal{F}_T, \mathcal{F}_S).$$

For  $p = 2$  a version of  $R$  based on additive r.v. will also play a role. Let

$$R_\ell = \sup \left\{ \text{corr}(U, V) : U = \sum_{j \in T} X_j, V = \sum_{j \in S} X_j, X_j \in L_2(\mathcal{F}_j), T \cap S = \emptyset \right\}.$$

Clearly,  $r \leq R_\ell \leq R$ . All three coefficients coincide for two  $\sigma$ -fields  $d = 2$  case. One can easily see that  $R_\ell = 0$  and  $R = 1$  can happen already for  $d = 3$ .

## 2.2 Main results

In Section 2.4 we present complete solution of (1) for two  $\sigma$ -fields case  $d = 2$ . For general families of  $\sigma$ -fields, there seems to be little hope to get existence and uniqueness results as precise as for  $d = 2$ . As Logan & Shepp [11] point out, complications arise even in relatively simple situations. One possible source of such complications is linear dependence between  $L_p^0(\mathcal{F}_j)$ . Suitable assumptions on maximal correlation coefficients exclude the latter.

The following result extends [5, Corollary 1] to infinite sequences of dependent families of  $\sigma$ -fields.

**Theorem 2.1** (i) Fix  $1 < p < \infty$  and suppose  $R < 1$ . Then equation (1) is solvable for  $Z$  for all  $X_j \in L_p^0(\mathcal{F}_j)$  such that  $E(\sum_j |X_j|^2)^{p/2} < \infty$ , and the solution is unique.

(ii) If  $R_\ell < 1$  then for all  $X_j \in L_2^0(\mathcal{F}_j)$  such that  $\sum_j EX_j^2 < \infty$  there is the unique additive solution  $Z$  to (1), and it satisfies  $E|Z|^2 \leq \frac{1+R_\ell}{1-R_\ell} \sum_j E|X_j|^2$ . Moreover, if  $d < \infty$  then  $Z$  is given by ACE formula [3].

If one isn't interested in sharp moment estimates for  $Z$ , and only finite families  $d < \infty$  are of interest, then one can iterate Theorem 2.11 for a pair of  $\sigma$ -fields, relaxing  $R < 1$ . By Lemma 3.2, this yields the following.

**Theorem 2.2** If  $d < \infty$ ,

$$\rho_\star = \max_{1 \leq j \leq d} \rho(\mathcal{F}_{\{1, \dots, j\}}, \mathcal{F}_{j+1}) < 1, \quad (6)$$

and  $1 < p < \infty$ , then equation (1) has an additive (4) solution  $Z$  for all  $X_j \in L_p^0(\mathcal{F}_j)$ .

Moreover, inequality (12) holds with  $q = 2$  and  $\delta = (1 - \rho_\star)^{d/2}$ .

The following criterion for solvability of the additive version of (1) uses the pairwise maximal correlation  $r$  and gives explicit alternative to ACE. For  $d = 2$  the assumptions are close to [4], except that we assume  $p = 2$  and (implicitly) linear independence.

**Theorem 2.3** If  $d < \infty$ ,  $r < \frac{1}{d-1}$ , and  $p = 2$ , then for all  $X_j \in L_2(\mathcal{F}_j)$  with  $EX_i = EX_j$  there is unique  $Z \in L_2$  such that (1) and (4) hold.

Moreover, the solution is given by the explicit series expansion

$$\begin{aligned} Z &= EX_1 \\ &+ \sum_{k=0}^{\infty} (-1)^k \sum_{i_1 \in I} \sum_{i_2 \in I \setminus i_1} \dots \sum_{i_k \in I \setminus i_{k-1}} \mathcal{E}_{i_1} \dots \mathcal{E}_{i_k} \sum_{j \in I \setminus i_k} (X_j - EX_j) \end{aligned} \quad (7)$$

(with the convention  $\sum_{j \in \emptyset} X_j = 0$ ).

Furthermore,  $\text{Var}(Z) \leq \frac{1}{1-r(d-1)} \sum_{j=1}^d \text{Var}(X_j)$ .

For finite families of  $\sigma$ -fields, inequality (12) is equivalent to solvability of (1). Lemma 3.3 gives a pairwise condition for (12) and was motivated by Breiman & Friedman [3]. This implies [3, Proposition 5.2].

**Corollary 2.4 ([3])** *If  $d < \infty$ , vector spaces  $L_2^0(\mathcal{F}_j)$  are linearly independent, and for all  $1 \leq j \leq d, k \neq j$ , the operators  $\mathcal{E}_j \mathcal{E}_k : L_2^0(\mathcal{F}_k) \rightarrow L_2^0(\mathcal{F}_j)$  are compact, then for all square integrable  $X_1, \dots, X_d$  with equal means  $EX_i = EX_j$ , there is the unique additive solution  $Z$  of (1).*

### 2.3 Conditioning with respect to random variables

We now state sufficient conditions for solvability of (1) in terms of joint distributions for finite families  $d < \infty$  of  $\sigma$ -fields generated by random variables  $\mathcal{F}_j = \sigma(\xi_j)$ , where  $\xi_1, \dots, \xi_d$  is a given random sequence.

We begin with the density criterion that gives explicit estimate for  $R$ , and was motivated by [14]. By Lemma 3.2, it implies that (1) has unique additive solution  $Z$  for all  $1 < p < \infty$ . Although it applies both to discrete and continuous distributions (typically, the density in the statement is with respect to the product of marginals), it is clear that the result is far from being optimal.

**Theorem 2.5** *Suppose there is a product probability measure  $\mu = \mu_1 \otimes \dots \otimes \mu_d$  such that the distribution of  $\xi_1, \dots, \xi_d$  on  $\mathbb{R}^d$  is absolutely continuous with respect to  $\mu$  and its density  $f$  is bounded away from zero and infinity,  $0 < b \leq f(x_1, \dots, x_d) \leq B < \infty$ . Then  $R \leq 1 - \frac{b}{B^2}$ .*

The next result follows from Corollary 2.4 by [13, page 106, Exercise 15]; it is stated for completeness.

**Proposition 2.6 ([3])** *Suppose  $d < \infty$  and for every pair of  $i \neq j$  the density  $f_{i,j}$  of the distribution of  $(\xi_i, \xi_j)$  with respect to the product measure  $\mu_{i,j} = \mu_i \otimes \mu_j$  of the marginals exists and*

$$\max_{i \neq j} \int \int f_{i,j}^2(x, y) d\mu_i(x) d\mu_j(y) < \infty. \quad (8)$$

*If vector spaces  $L_2^0(\mathcal{F}_j)$  are linearly independent  $p = 2$ , then  $R_\ell < 1$ . In particular, (1) has unique additive solution  $Z$  for all square integrable  $X_1, \dots, X_d$  with equal means  $EX_i = EX_j$ .*

In general, linear independence is difficult to verify (vide [11], where it fails). The following consequence of Proposition 2.6 gives a relevant “density criterion”.

**Corollary 2.7** *Suppose the density  $f$  of the distribution of  $\xi_1, \dots, \xi_d$  ( $d < \infty$ ) with respect to the product of marginals  $\mu = \mu_1 \otimes \dots \otimes \mu_d$  exists. If  $f$  is strictly positive, i.e.,  $\mu(\{(x_1, \dots, x_d) : f(x_1, \dots, x_d) = 0\}) = 0$  and  $\int f^2 d\mu < \infty$  then there is an additive solution to (1) for all  $X_j \in L_2(\mathcal{F}_j)$  such that  $EX_i = EX_j$ .*

In relation to Theorem 2.5, one should note that the lower bound on the density is of more relevance. (On the other hand, in Theorem 2.5 we use the density with respect to arbitrary product measure rather than the product of marginals.)

**Proposition 2.8** *Let  $f$  be the density of the absolute continuous part of the distribution of  $\xi_1, \dots, \xi_d$  ( $d < \infty$ ) with respect to the product of marginals  $\mu = \mu_1 \otimes \dots \otimes \mu_d$ . If  $f$  is bounded away from zero, i.e., there is  $b > 0$  such that  $\mu(\{(x_1, \dots, x_d) : f(x_1, \dots, x_d) \geq b\}) = 1$ , then (12) holds for all  $1 < q < \infty$ . In particular, for  $1 < p < \infty$*

*for  $X_j \in L_p(\mathcal{F}_j)$  such that  $EX_i = EX_j$  there is an additive solution to (1).*

## 2.4 Results for two $\sigma$ -fields

This case is rather completely settled. Most of the results occurred in various guises in the literature. They are collected below for completeness, and to point out what to aim for in the more general case.

The following shows that for  $d = 2$  there is at most one solution of (1) and (4). (Clearly, there is no  $Z$  if  $X_1, X_2$  are not consistent, e.g., if  $EX_1 \neq EX_2$ .)

**Proposition 2.9** *Given  $X_j \in L_p(\mathcal{F}_j)$ ,  $1 \leq p \leq \infty$ , there is at most one  $Z = Z_1 + Z_2 + Z' \in L_1$  such that (1) holds and  $E(Z_j|\mathcal{F}_1 \cap \mathcal{F}_2) = 0$ .*

Since best additive approximations satisfy (1), uniqueness allows to consider the inverse problem (1) instead. This is well known, c.f., [7].

**Corollary 2.10** *If  $p = 2$  and the best additive approximation  $Z = Z_1 + Z_2 + Z'$  of  $Y \in L_2$  (i.e.,  $Z$  minimizing  $E(Y - E(Y|\mathcal{F}_1 \cap \mathcal{F}_2) - (Z_1 + Z_2))^2$ ) exists, then it is given by the solution to Problem SA (if solvable).*

The following result points out the role of maximal correlation and comes from [4].

**Theorem 2.11** ([4]) *Suppose  $1 < p < \infty$  is fixed. The following conditions are equivalent:*

1. *There is a minimal solution to (1) for all consistent  $X_1, X_2$  in  $L_p(\mathcal{F}_1), L_p(\mathcal{F}_2)$  respectively;*
2. *here is an additive solution to (1) for all consistent  $X_1, X_2$  in  $L_p(\mathcal{F}_1), L_p(\mathcal{F}_2)$  respectively;*
3.  *$\tilde{\rho} < 1$ .*

Moreover, the explicit consistency condition is

$$E\{X_1|\mathcal{F}_1 \cap \mathcal{F}_2\} = E\{X_2|\mathcal{F}_1 \cap \mathcal{F}_2\}.$$

Furthermore, if  $E(Z|\mathcal{F}_1 \cap \mathcal{F}_2) = 0$ , the minimum norm in (3) is bounded by  $E|Z|^2 \leq \frac{1}{1-\tilde{\rho}}(EX_1^2 + EX_2^2)$  and the bound is sharp.

The solution is explicit:

$$Z = E(X_1|\mathcal{F}_1 \cap \mathcal{F}_2) + \sum_{k=0}^{\infty} (\mathcal{E}_2 \mathcal{E}_1)^k (X_2 - \mathcal{E}_2 X_1) + \sum_{k=0}^{\infty} (\mathcal{E}_1 \mathcal{E}_2)^k (X_1 - \mathcal{E}_1 X_2) \quad (9)$$

and both series converge in  $L_p$ .

**Remark 2.2** *For  $p = 2$  formula (9) resembles the following explicit expansion for the orthogonal projection  $L_2 \rightarrow cl(L_2(\mathcal{F}_1) \oplus L_2(\mathcal{F}_2))$  (see [1]).*

$$Z = E\{Y|\mathcal{F}_1 \cap \mathcal{F}_2\} + \sum_{k=1}^{\infty} (\mathcal{E}_1(\mathcal{E}_2 \mathcal{E}_1)^k + \mathcal{E}_2(\mathcal{E}_1 \mathcal{E}_2)^k - (\mathcal{E}_1 \mathcal{E}_2)^k - (\mathcal{E}_2 \mathcal{E}_1)^k) Y; \quad (10)$$

(the series converges in  $L_2$ ).

### 3 Proofs

The following uniqueness result is proved in [3] for the square-integrable case  $p = 2$  (the new part is  $1 < p < 2$ ).

**Lemma 3.1** (i) If  $L_p^0(\mathcal{F}_j)$  are linearly independent and  $d < \infty$ , then for every  $\{X_j\}$  in  $L_p(\mathcal{F}_j)$ ,  $p \geq 2$ , there is at most one solution of (1) in the additive class (4).  
(ii) If (12) holds with  $q = 2$ , then for every  $\{X_j\}$  in  $L_p(\mathcal{F}_j)$ ,  $p \geq 2$ , there is at most one solution of (1) in the additive class (4).  
(iii) Fix  $1 < p < \infty$ . If there are constants  $c, C$  such that for all centered  $\{X_j\} \in L_q(\mathcal{F}_j)$

$$cE\left(\sum_{j=1}^d X_j^2\right)^{q/2} \leq E\left|\sum_{j=1}^d X_j\right|^q \leq CE\left(\sum_{j=1}^d X_j^2\right)^{q/2} \quad (11)$$

holds for  $q = p$  and for the conjugate exponent  $q = \frac{p}{p-1}$ , then for every  $\{X_j\}$  in  $L_p(\mathcal{F}_j)$  there is at most one solution of (1) in the additive class (4).

**Proof of Lemma 3.1.** The case  $p = 2$  goes as follows. Suppose  $Z = Z_1 + Z_2 + \dots$  has  $\mathcal{E}_j(Z) = 0$  for all  $j$ . Then  $EZ^2 = \sum_j EZZ_j = \sum_j E(Z_j\mathcal{E}_j(Z)) = 0$ . This implies that  $Z_j = 0$  for all  $j$  either by linear independence, or by (12).

The second part uses the existence part of the proof of Theorem 2.1. Take  $Z = \sum_j Z_j$  ( $L_p$ -convergent series) such that  $\mathcal{E}_j(Z) = 0$ . Then by (11)

$$\|Z\|_p \leq C(E(\sum_j Z_j^2)^{p/2})^{1/p} = C \sum_j E(Z_j X_j),$$

where  $E(\sum_j X_j^2)^{q/2} = 1$ ,  $1/p + 1/q = 1$  and  $X_j \in L_q^0(\mathcal{F}_j)$ . The latter holds because the conjugate space to  $L_q^0(\ell_2(\mathcal{F}_j))$  is  $L_p^0(\ell_2(\mathcal{F}_j))$ . The existence part of the proof of Theorem 2.1 implies that there is  $\tilde{Z} \in L_q$  such that  $\mathcal{E}_j(\tilde{Z}) = X_j$  and  $\tilde{Z} = \sum_j \tilde{Z}_j$  with  $\tilde{Z}_j \in L_q^0(\mathcal{F}_j)$ . Therefore

$$\sum_j E(Z_j X_j) = \sum_j E(Z_j \tilde{Z}) = E(Z \tilde{Z}) = E(Z \sum_j \tilde{Z}_j) = \sum_j E(\tilde{Z}_j \mathcal{E}_j(Z)) = 0.$$

This shows  $E|Z|^p = 0$  and by the left hand side of (11) we have  $Z_j = 0$  a.s. for all  $j$ .  $\square$

**Proof of Proposition 2.9.** Clearly  $Z' = E\{Z|\mathcal{F}_1 \cap \mathcal{F}_2\}$  is uniquely determined by  $Z$  and without loosing generality we may assume  $Z' = 0$ . Suppose that a particular  $Z = Z_1 + Z_2$  has  $\mathcal{E}_j Z = 0$ . Then  $Z_1 = \mathcal{E}_1 \mathcal{E}_2 Z_1$ . Using this iteratively, by ‘‘alternierende Verfahren’’ (see [12]) we get  $Z_1 = (\mathcal{E}_1 \mathcal{E}_2)^k Z_1 \rightarrow E(Z_1|\mathcal{F}_1 \cap \mathcal{F}_2) = 0$ . By symmetry,  $Z_2 = 0$  and uniqueness follows.  $\square$

**Proof of Corollary 2.10.** Without loss of generality we may assume  $E(Y|\mathcal{F}_1 \cap \mathcal{F}_2) = 0$ . For optimal  $Z = Z_1 + Z_2$  we have

$$\begin{aligned} \min &= E(Y - (Z_1 + Z_2))^2 \\ &= E(Y - (\mathcal{E}_1(Y) - \mathcal{E}_1(Z_2) + Z_2))^2 + E(\mathcal{E}_1(Y) - \mathcal{E}_1(Z))^2 \geq \min + E(\mathcal{E}_1(Y) - \mathcal{E}_1(Z))^2. \end{aligned}$$

Since the same analysis applies to  $\mathcal{E}_2$ , the optimal  $Z$  has to satisfy (1). By Theorem 2.9, there is only one such  $Z$ , so this one has to be the optimal one.  $\square$

**Proof of Theorem 2.11.** For  $E(Y_j|\mathcal{F}_1 \cap \mathcal{F}_2) = 0$  we have

$$E(Y_1 + Y_2)^2 \geq EY_1^2 + EY_2^2 - 2\tilde{\rho}(EY_1^2 EY_2^2)^{1/2} \geq (1 - \tilde{\rho})(EY_1^2 + EY_2^2)$$

Therefore the linear operator  $T : \mathcal{L}_2^0 \rightarrow \mathcal{L}_2^0(\mathcal{F}_1) \times \mathcal{L}_2^0(\mathcal{F}_2)$  given by  $T(Y) = (\mathcal{E}_1(Y), \mathcal{E}_2(Y))$  is onto and the norm of its left inverse is bounded by  $(1 - \bar{\rho})^{-1/2}$  (here  $\mathcal{L}_p^0$  denotes the null space of the linear operator  $E(\cdot|\mathcal{F}_1 \cap \mathcal{F}_2)$  on  $L_p$ ). This proves the bound  $E|Z|^2 \leq (EX_1^2 + EX_2^2)/(1 - \bar{\rho})$ .

Because of the explicit formula for  $Z$ , it is clear that  $3 \Rightarrow 2$ ; implication  $2 \Rightarrow 1$  holds by general principles (see Remark 2.1). The equivalence  $1 \Leftrightarrow 3$  is in [4].  $\square$

It is easy to see that if for given  $\{X_j\}$  (1) is solvable in the additive class (4), then there is also a minimal solution (3), see Remark 2.1. The following shows that for finite families of  $\sigma$ -fields the solvability of both problems is actually equivalent, at least when trivial constraints  $EX_i = EX_j$  are the only ones to be used.

**Lemma 3.2** *Fix  $1 < p < \infty$  and suppose  $d < \infty$ . The following conditions are equivalent*

- (i) Equation (1) has an additive (4) solution  $Z$  for all  $X_j \in L_p^0(\mathcal{F}_j)$ ;
- (ii) Equation (1) has a minimal (3) solution  $Z$  for all  $X_j \in L_p^0(\mathcal{F}_j)$ ;
- (iii) There is  $\delta = \delta(q) > 0$  such that for all  $X_j \in L_q^0(\mathcal{F}_j)$

$$E \left| \sum_j X_j \right|^q \geq \delta^q \sum_j E |X_j|^q, \quad (12)$$

where  $1/p + 1/q = 1$ .

Moreover, if inequality (12) holds, then there is an additive solution  $Z$  to (1) with  $E|Z|^p \leq \frac{1}{\delta^p} \sum_j E|X_j|^p$ .

**Remark 3.1** *In addition, if  $L_q^0(\mathcal{F}_j)$  are linearly independent then the following equivalent condition can be added:*

- (iv)  $L_q^0(\mathcal{F}_1) \oplus \dots \oplus L_q^0(\mathcal{F}_d)$  is a closed subspace of  $L_q(\mathcal{F}_I)$ .

**Proof of Lemma 3.2.** (iii) $\Rightarrow$ (i) Consider the linear bounded operator  $T : L_p \rightarrow \ell_p(L_p^0(\mathcal{F}_j))$  defined by  $Z \mapsto (E(Z|\mathcal{F}_j) : j = 1, \dots, d)$ . The conjugate operator  $T^* : \ell_q \rightarrow L_q$  is given by  $(X_j) \mapsto \sum_{j=1}^d X_j$ .

Coercivity criterion for  $T$  being onto is  $\|T^*(X_j)\|_{L_q} \geq \delta \|(X_j)\|_{\ell_q}$ , which is (12), see [13, Theorem 4.15]. Therefore (i) follows.

The left-inverse operator has  $\ell_p \rightarrow L_p$  operator norm  $\|T^{-1}\| \leq 1/\delta$ , which gives the estimate of the norm of  $Z$  as claimed.

(i) $\Rightarrow$ (ii) If there is additive solution, then  $X_j$  are consistent and Remark 2.1 implies that there is the minimal norm solution.

(ii) $\Rightarrow$ (iii) This is a simple operator coercivity of linear operators analysis, see [13, Theorem 4.15]. Namely, if for all  $X_j$  there is  $Z$  such that (1) holds, then the linear operator  $T : L_p^0 \rightarrow L_p^0(\mathcal{F}_1) \times \dots \times L_p^0(\mathcal{F}_d)$  given by  $Z \mapsto (\mathcal{E}_j(Z))$  is onto. Therefore the conjugate operator satisfies

$$\|T^*(X_1, \dots, X_d)\|_q \geq \delta \|(X_1, \dots, X_d)\|_{\ell_q(L_q(\mathcal{F}_j))}$$

and inequality (12) follows.

$\square$

**Proof of Remark 3.1.** (iv) $\Rightarrow$ (iii) If  $L_q^0(\mathcal{F}_1) \oplus \dots \oplus L_q^0(\mathcal{F}_d)$  is a closed subspace of  $L_q(\mathcal{F}_I)$  then (12) holds. Indeed, by linear independence, the linear operator  $(X_1 + \dots + X_d) \mapsto (X_1, \dots, X_d)$  is an injection of the Banach space  $L_q^0(\mathcal{F}_1) \oplus \dots \oplus L_q^0(\mathcal{F}_d)$  into  $L_q^0(\mathcal{F}_1) \times \dots \times L_q^0(\mathcal{F}_d)$  with  $\ell_q$  norm. Since the range is closed, open mapping theorem ([13, Theorem 2.11]) implies (12).

(iii) $\Rightarrow$ (iv) is trivial.  $\square$

**Proof of Theorem 2.1.** From the proof of Bryc & Smolenski [6, (7)] we have the left hand side of the inequality

$$cE \left| \sum_{j=1}^d \epsilon_j X_j \right|^q \leq E \left| \sum_{j=1}^d X_j \right|^q \leq CE \left| \sum_{j=1}^d \epsilon_j X_j \right|^q.$$

(The right hand side is [6, (7)].)

By the Khinchin inequality this implies (11). Note, that a more careful analysis gives explicit estimates for the constants involved.

For  $q = 2$  the above is replaced by

$$\frac{1 - R_\ell}{1 + R_\ell} \sum_{j=1}^d EX_j^2 \leq E \left| \sum_{j=1}^d X_j \right|^2 \leq \frac{1 + R_\ell}{1 - R_\ell} \sum_{j=1}^d EX_j^2,$$

which is given in [2, Lemma 1].

Existence of the solution follows now from functional analysis: Consider the linear bounded (cf (11)) operator  $T : L_p^0 \rightarrow L_p^0(\ell_2(\mathcal{F}_j))$  defined by  $Z \mapsto (E(Z|\mathcal{F}_j) : j = 1, 2, \dots)$ . The conjugate operator  $T^* : L_q^0(\ell_2) \rightarrow L_q^0$  is given by  $(X_j) \mapsto \sum_{j=1}^\infty X_j$ .

Coercivity criterion for  $T$  being onto, see [13, Theorem 4.15], is  $\|T^*(X_j)\|_{L_q} \geq \delta \|(X_j)\|_{L_q(\ell_2)}$ , which follows from (11). Therefore the existence of a solution to (1) follows and the minimal solution exists by Remark 2.1.

For  $p = 2$  inequalities (11) show that  $L_2^0(\ell_2) = \ell_2(L_2^0(\mathcal{F}_j)) \ni (X_j)$  generates the  $L_2$  convergent series  $\sum_j X_j$ . Denote by  $H$  the set of random variables represented by such series. By (11)  $H$  is closed and since the orthogonal projection onto  $H$  shrinks the norm, the minimal solution to (1) has to be in  $H$ , thus it is additive (4). The left-inverse operator has  $\ell_2 \rightarrow L_2$  operator norm  $\|T^{-1}\| \leq (\frac{1-R_\ell}{1+R_\ell})^{1/2}$ , which gives the estimate for the norm of  $Z$  as claimed.

The uniqueness follows from (11) by Lemma 3.1.  $\square$

**Proof of Theorem 2.2.** Use Theorem 2.11 to produce recurrently

$\mathcal{F}_{1,2}$ -measurable  $Z^1$  such that  $\mathcal{E}_1(Z^1) = X_1, \mathcal{E}_2(Z^1) = X_2$ ;

$\mathcal{F}_{1,2,3}$ -measurable  $Z^2$  such that  $\mathcal{E}_{1,2}(Z^2) = Z^1, \mathcal{E}_3(Z^2) = X_3$ ;

$\vdots$

$\mathcal{F}_{1,\dots,d}$ -measurable  $Z^d$  such that  $\mathcal{E}_{1,\dots,d-1}(Z^d) = Z^{d-1}, \mathcal{E}_d(Z^d) = X_d$ .

This shows that for all  $d < \infty$  there is a solution to (1), and hence there is a minimal solution.

Therefore, by Lemma 3.2 there is an additive solution (4), and (12) holds.

Notice that for  $d < \infty$  inequality (12) implies (11), which by Lemma 3.1 implies uniqueness.

The inequality (12) for  $q = 2$  follows recurrently from  $E(\sum_{j=1}^k X_j + X_{k+1})^2 \geq (1 - \rho_*)(E(\sum_{j=1}^k X_j)^2 + EX_{k+1}^2)$ .  $\square$

**Proof of Theorem 2.3.**

To verify that the series in (7) converge, notice that for  $j \neq i_k$

$$\|\mathcal{E}_{i_1} \dots \mathcal{E}_{i_k}\|_{L_2^0(\mathcal{F}_j) \rightarrow L_2^0} \leq r^k.$$

Therefore

$$\begin{aligned} & \left\| \sum_k (-1)^k \sum_{i_1 \in I} \sum_{i_2 \in I \setminus i_1} \dots \sum_{i_k \in I \setminus i_{k-1}} \mathcal{E}_{i_1} \dots \mathcal{E}_{i_k} \sum_{j \in I \setminus i_k} (X_j - EX_j) \right\|_2 \\ & \leq 2d \sum_k (d-1)^k r^k \max_j \|X_j\|_2. \end{aligned}$$

Clearly, (4) holds true. We check now that  $Z$  defined by (7) satisfies (1). To this end, without loss of generality we assume  $EX_j = 0$  and we verify (1) for  $j = 1$  only. Splitting the sum (7) we get

$$\mathcal{E}_1(Z) = \sum_{k=0}^{\infty} (-1)^k \sum_{i_2 \in I \setminus 1} \dots \sum_{i_k \in I \setminus i_{k-1}} \mathcal{E}_1 \mathcal{E}_{i_2} \dots \mathcal{E}_{i_k} \sum_{j \in I \setminus i_k} (X_j - EX_j)$$

$$+ \sum_{k=0}^{\infty} (-1)^k \sum_{i_1 \in I \setminus 1} \sum_{i_2 \in I \setminus i_1} \dots \sum_{i_k \in I \setminus i_{k-1}} \mathcal{E}_1 \mathcal{E}_{i_1} \dots \mathcal{E}_{i_k} \sum_{j \in I \setminus i_k} (X_j - EX_j).$$

The 0-th term of the sum on the left is  $X_1$  and the  $k$ -th term of the sum on the left cancels the  $(k-1)$  term of the sum on the right. Therefore  $\mathcal{E}_1(Z) = X_1$ .

To prove the uniqueness, it suffices to notice that  $r < \frac{1}{d-1}$  implies linear independence. Alternatively, suppose that both  $Z = Z_1 + \dots + Z_d$  and  $Z' = Z_1' + \dots + Z_d'$  have the same conditional moments  $\mathcal{E}_1$ . Then  $\|Z_1 - Z_1'\|_2 = \|\mathcal{E}_1(Z_2 + \dots + Z_d) - \mathcal{E}_1(Z_2' + \dots + Z_d')\| \leq r \sum_{j=2}^d \|Z_j - Z_j'\|_2$ , and the similar estimate holds for all other components. Therefore  $\sum_{j=1}^d \|Z_j - Z_j'\|_2 \leq r(d-1) \sum_{j=1}^d \|Z_j - Z_j'\|_2$ . Since  $r < 1/(d-1)$ , this implies the sum vanishes, proving uniqueness.

To prove the variance estimate notice that  $r < 1/(d-1)$  implies (12) with  $p = 2$  and  $\delta^2 = 1 - r(d-1)$ . Indeed,

$$E \left| \sum_{j=1}^d X_j \right|^2 \geq \sum_{j=1}^d EX_j^2 - r \sum_{j \neq k} (EX_j^2 EX_k^2)^{1/2}.$$

The estimate now follows from the elementary inequality

$$\frac{1}{d-1} \sum_{j=1}^d \sum_{k=1, k \neq j}^d a_k a_j \leq \frac{1}{d-1} \sum_{j=1}^d \sum_{k=1, k \neq j}^d \frac{1}{2} (a_k^2 + a_j^2) = \sum_{j=1}^d a_j^2,$$

valid for arbitrary numbers  $a_1, \dots, a_d$ .  $\square$

**Lemma 3.3** *If  $d < \infty$ , vector spaces  $L_2^0(\mathcal{F}_j)$  are linearly independent, and for all  $1 \leq j \leq d, k \neq j$ , the operators  $\mathcal{E}_j \mathcal{E}_k : L_2^0(\mathcal{F}_k) \rightarrow L_2^0(\mathcal{F}_j)$  are compact, then  $R_\ell < 1$ ; hence inequality (12) holds for  $q = 2$ .*

**Proof of Lemma 3.3.** The proof is similar to the proof of Proposition 2.9 with  $T = P_S P_Q$ , where  $S, Q$  are disjoint and  $P_Q$  denotes the orthogonal projection onto the  $L_2$ -closure of  $\bigoplus_{j \in Q} L_2^0(\mathcal{F}_j)$ ; operator  $T$  is compact, compare [3, Proposition 5.3]. Details are omitted.  $\square$

**Proof of Theorem 2.5.** Take  $U \in L_2(\mathcal{F}_S), V \in L_2(\mathcal{F}_T)$  with disjoint  $S, T \subset I$  and such that  $EU = EV = 0, EU^2 = EV^2 = 1, EUV = \rho$ . Then

$$E(U - V)^2 = 2 - 2\rho \geq 2 \frac{b}{B^2}.$$

Indeed, we have

$$\begin{aligned} E(U - V)^2 &= \int_{\mathbb{R}^d} (U(\mathbf{x}) - V(\mathbf{x}))^2 f(\mathbf{x}) d\mu(\mathbf{x}) \geq b \int_{\mathbb{R}^d} (U(\mathbf{x}) - V(\mathbf{x}))^2 d\mu(\mathbf{x}) \\ &= b \int_{\mathbb{R}^d} \int_{\mathbb{R}^d} (U(\mathbf{x}) - V(\mathbf{y}))^2 d\mu(\mathbf{y}) d\mu(\mathbf{x}) \\ &\geq \frac{b}{B^2} \int_{\mathbb{R}^d} \int_{\mathbb{R}^d} (U(\mathbf{x}) - V(\mathbf{y}))^2 f(\mathbf{y}) d\mu(\mathbf{y}) f(\mathbf{x}) d\mu(\mathbf{x}) = \frac{2b}{B^2}. \end{aligned}$$

$\square$

Since the above analysis can also be carried through for  $E(U + V)^2$ , we get the following.

**Corollary 3.4 (c.f.[14] Lemma 1)** *Under the assumption of Theorem 2.5, for  $V_j \in L_2^0(\mathcal{F}_j)$  we have*

$$E|V_1 + \dots + V_d|^2 \geq \frac{b}{2B^2 - b} (EV_1^2 + \dots + EV_d^2)$$

**Lemma 3.5** *Let  $f$  be the density of the absolute continuous part of the distribution of  $\xi_1, \dots, \xi_d$  ( $d < \infty$ ) with respect to the product of marginals  $\mu = \mu_1 \otimes \dots \otimes \mu_d$ . If  $f$  is strictly positive, i.e.,  $\mu(\{(x_1, \dots, x_d) : f(x_1, \dots, x_d) = 0\}) = 0$ , then vector spaces  $L_2^0(\mathcal{F}_j)$  are linearly independent.*

**Proof of Lemma 3.5.** Suppose  $X_1 = X_1(\xi_1), \dots, X_d = X_d(\xi_d) \in L_2^0$  are non-zero. Denote by  $\mu$  the product of marginal measures on  $\mathbb{R}^d$  and let  $A_\epsilon = \{(x_1, \dots, x_d) : f(x_1, \dots, x_d) > \epsilon\}$ . Choose  $\epsilon > 0$  such that

$$\int_{A_\epsilon} |\sum X_j|^2 d\mu < \frac{1}{2} \sum EX_j^2.$$

Then

$$\begin{aligned} E|\sum X_j|^2 &\geq \int_{A_\epsilon} |\sum X_j(x_j)|^2 f(x_1, \dots, x_d) d\mu \\ &\geq \epsilon \int_{A_\epsilon} |\sum X_j|^2 d\mu \geq \frac{\epsilon}{2} \sum EX_j^2 > 0. \end{aligned}$$

This proves linear independence of  $X_j$ .  $\square$

**Proof of Proposition 2.8.** This follows the proof of Proposition 3.5 with  $\epsilon = b$ . Namely,

$$E|\sum X_j|^q \geq b \int_{\mathbb{R}^d} |\sum X_j|^q d\mu \geq cE(\sum X_j^2)^{q/2}.$$

The last inequality holds by the Marcinkiewicz-Zygmund inequality, because under  $\mu$  random variables  $X_j$  are independent and centered.  $\square$

## 4 Example

The following simple example illustrates sharpness of some moment estimates.

**Example 4.1** *Let  $d < \infty$ . Suppose  $X_1, \dots, X_d$  are in  $L_2$ , centered and have linear regressions, i.e., there are constants  $a_{i,j}$  such that  $E(X_i|X_j) = a_{i,j}X_j$  for all  $i, j$  (for instance, this holds true for  $(X_1, \dots, X_d)$  with elliptically contoured distributions, or when each  $X_j$  is two-valued). Let  $C = [C_{i,j}]$  be the covariance matrix. Clearly, if either  $R_\ell < 1$  or  $r < 1/(d-1)$ , then  $C$  is non degenerate.*

*Explicit solutions illustrating Theorems 2.11, 2.1, and 2.3 are then possible:*

*It is easy to check that  $Z = \sum_{j=1}^d \theta_j X_j$ , where  $\begin{bmatrix} \theta_1 \\ \vdots \\ \theta_d \end{bmatrix} = C^{-1} \begin{bmatrix} 1 \\ \vdots \\ 1 \end{bmatrix}$ , satisfies  $E(Z|X_j) = X_j$  for*

*all  $j$ , and it clearly is additive. Moreover, if  $\text{corr}(X_i, X_j) = \rho$  doesn't depend on  $i \neq j$ , then*

$$Z = \frac{1}{1 + (d-1)\rho} (X_1 + \dots + X_d).$$

*In particular, for  $d = 2$  we have  $Z = \frac{1}{1+\rho}(X_1 + X_2)$ . For negative  $\rho$ , those point out the sharpness of estimates for  $EZ^2$ .*

*(It is easy to see directly from  $E(X_1 + \dots + X_d)^2 = d + d(d-1)\rho \geq \frac{1-R}{1+R}$  that  $(d-1)\rho \geq -\frac{2R}{1+R} > -1$ , provided  $R_\ell < 1$ . Therefore  $Z$  is well defined.)*

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## References

- [1] N. Aronszajn, Theory of reproducing kernels. *Trans. Amer. Math. Soc.* 68 (1950), 337–404.
- [2] R. Bradley, On the spectral density and asymptotic normality of weakly dependent random fields. *Journ. Theoret. Probab.* 5 (1992), 355–373.
- [3] L. Breiman & J. H. Friedman, Estimating Optimal Transformations for Multiple Regression and Correlation. *JASA* 80 (1985), 580–598.
- [4] W. Bryc, Conditional expectation with respect to dependent  $\sigma$ -fields. In: *Proceedings of the Seventh Conference on Probability Theory (Brasov)*, Ed. M. Iosifescu, Bucuresti, 1984, 409–411.
- [5] W. Bryc & S. Kwapien, On the conditional expectation with respect to a sequence of independent  $\sigma$ -fields. *Z. Wahrsch. Verw. Gebiete* 46 (1979), 221–225.
- [6] W. Bryc & W. Smoleński, Moment conditions for almost sure convergence of weakly correlated random variables. *Proc. Amer. Math. Soc.* 119 (1993), 629–635.
- [7] A. Buja & T. Hastie & R. Tibshirani, Linear smoothers and additive models. *Ann. Statist.* 17 (1989), 453–510.
- [8] C. Groetsch, *Inverse Problems in the Mathematical Sciences*. Vieweg, Braunschweig (Germany), 1993.
- [9] H. G. Kellerer, Masstheoretische Marginalproblem, *Math. Ann.* 153 (1964), 168–198.
- [10] J. Komlos, A generalization of a problem of Steinhaus. *Acta Math. Hung.* 18 (1967), 217–229.
- [11] B. F. Logan & L. A. Shepp, Optimal reconstruction of a function from its projections. *Duke Math. Journ.* 42 (1975), 645–659.
- [12] G. C. Rota, An “alterninde Verfahren” for general positive operators. *Bull. Amer. Math. Soc.* 68 (1962), 95–102.
- [13] W. Rudin, *Functional Analysis*. McGraw, New York 1973.
- [14] C. J. Stone, Additive regression and other nonparametric models. *Ann. Statist.* 13 (1985), 689–705.
- [15] V. Strassen, The existence of probability measures with given marginals, *Ann. Math. Statist.* 36 (1965), 423–439.