

# Large deviations for quadratic functionals of Gaussian processes

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## Abstract

The Large Deviation Principle is derived for several unbounded additive functionals of centered stationary Gaussian processes. For example, the rate function corresponding to  $\frac{1}{T} \int_0^T X_t^2 dt$  is the Fenchel-Legendre transform of  $L(y) = -\frac{1}{4\pi} \int_{-\infty}^{\infty} \log(1 - 4\pi y f(s)) ds$ , where  $X_t$  is a continuous time process with the bounded spectral density  $f(s)$ . Similar results in the discrete-time version are obtained for the energy of multivariate Gaussian processes and for the sums of  $p < 2$  powers. Explicit rate functions are obtained in several instances.

## 1 Introduction

Let  $E$  be a separable Banach space. Throughout most of the paper  $E = R$ , except in Proposition 2.4, where  $E = R^2$ , in Proposition 2.5, where  $E = R^{d+1}$ , and in Proposition 2.2, where the general case is considered.

Suppose  $\mathbf{S}_n, n > 0$ , are  $E$ -valued random variables. We shall say that  $\{n^{-1}\mathbf{S}_n\}$  satisfies the Large Deviation Principle (LDP), if there is a lower semicontinuous rate function  $I : E \rightarrow [0, \infty]$ , with compact level sets  $I^{-1}([0, a])$  for all  $a > 0$ , and such that

$$\liminf_{n \rightarrow \infty} n^{-1} \log P(n^{-1}\mathbf{S}_n \in A) \geq - \inf_{x \in A} I(x)$$

for all open subsets  $A \subset E$ ;

$$\limsup_{n \rightarrow \infty} n^{-1} \log P(n^{-1}\mathbf{S}_n \in A) \leq - \inf_{x \in A} I(x)$$

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for all closed subsets  $A \subset E$ .

We shall work with the continuous indices  $n$  (which below are denoted by  $T$  rather than  $n$ ) as well as with the discrete  $n = 1, 2, \dots$ ; in Section 2.3 we shall also consider other normalizations.

For a general stationary process  $X_j$ , the Large Deviation Principle for the empirical measures (i.e., in the discrete time setup corresponding to  $\mathbf{S}_n = \sum_{j=1}^n \delta_{X_j}$ ) and the related question for bounded additive functionals (i.e.,  $\mathbf{S}_n = \sum_{j=1}^n F(X_j)$ , with bounded  $F(\cdot)$ ) have been studied by a number of authors under some restriction on dependence; see [13, Section 6.4] for a sample of results, and [13, Section 6.9 page 280] for relevant references. Gaussian processes were studied in [15], LDP for Gaussian fields is given in [24], see also [14] for an interesting case.

Large deviations for general unbounded additive functionals of Markov chains under minimal assumptions were studied e.g. in [22].

Quadratic forms in Gaussian random variables have been studied by various asymptotic methods e.g. in statistical and electrical engineering literature; for an early paper using the saddle point method to approximate the distribution for a fixed number of variables, see [21], see also [20]. There is also a number of papers on the CLT, see e.g. [1], [25] and the references therein. Several results directly pertinent to the Large Deviation Principle have appeared: [11] gives a version of the Large Deviation Principle restricted to certain sets and obtained using the Grenander-Szegö method as employed below (and also in [2] and [8]). Their results however deal with quadratic forms in implicit way and without explicit expressions for the rate function; [2] presents the heuristic reasoning that motivated and facilitated much of this paper; in [9], the LDP given as Corollary 2.1 below is stated under an additional technical assumption; in [7] explicit rate function is found for autoregressive AR(1) processes.

In this paper, the Large Deviation Principle is derived for several unbounded additive functionals of stationary centered Gaussian processes that possess spectral density. Of those, quadratic functionals received most attention - for electrical engineering motivation the reader is referred to [8]; motivation from control theory is presented in the introduction to [7]; statistical motivation can be read out from [11].

The following describes the contents of the paper. In Theorem 2.1 we show that  $\frac{1}{T} \int_0^T X_t^2 dt$ , where  $X_t$  is a continuous time process with the bounded spectral density  $f(s)$ , satisfies the Large Deviation Principle and the rate function is given by the Fenchel-Legendre transform of  $L(y) = -\frac{1}{4\pi} \int_{-\infty}^{\infty} \log(1 - 4\pi y f(s)) ds$ . In Theorem 2.2 we show the corresponding multivariate discrete-time result. The LDP with normalization of  $o(n)$  and the quadratic rate function (corresponding to more moderate deviations) is derived in Theorem 2.3 for unbounded spectral densities (see [13, Section 3.7] for such results in the context of Cramér's theorem). In Corollary 2.2 we analyze  $\frac{1}{n} \sum_{j=1}^n |X_j|^p$  for  $p < 2$ ; the truncation lemma used in the proof allows also to prove the Large Deviation Principle for H-mixing sequences, in particular extending a result from [5]. In Section 2.5 we specify the univariate version of Theorem 2.2 (Corollary 2.1) to ARMA(p,q) processes. Proposition 2.3 points out the relevance to the CLT. In Section 2.7 we incorporate a non-zero mean in the univariate version of Theorem 2.2, thus deriving the Large Deviation Principle for the empirical variance. In Section 2.8 the Large Deviation Principle is derived for the empirical autocorrelation vector of an i.i.d. process  $X_j$  and some counter intuitive results concerning the validity of this LDP when  $\{X_j\}$  is an AR(1) process are presented. An approach to higher order expansions is sketched in Section 2.9. Examples with explicit rate function are collected in Section 3.

## 2 Results

This section contains statements of our main results. The proofs are given in Section 4, except for those results that are marked as immediate consequences of other theorems.

## 2.1 Continuous time

Let  $\{X_t\}$  be a real-valued, centered, separable stationary Gaussian process with the covariance  $R(t) = E(X_0 X_t)$  and spectral density  $f(s)$ , i.e.,  $R(t) = \int_{-\infty}^{\infty} e^{its} f(s) ds$ .

Denote  $S_T = \int_0^T X_t^2 dt$ ,  $M = \text{ess sup } f(s)$ .

**Theorem 2.1** *Suppose that  $\{X_t\}_{t \geq 0}$  has bounded spectral density function  $f(s) \in L_1(R, ds)$ . Then  $\{\frac{1}{T} S_T\}$  satisfies the Large Deviation Principle with the rate function*

$$I(x) = \sup_{-\infty < y < 1/(4\pi M)} \{xy - L(y)\},$$

where for  $y < 1/(4\pi M)$

$$L(y) = -\frac{1}{4\pi} \int_{-\infty}^{\infty} \log(1 - 4\pi y f(s)) ds. \quad (1)$$

## 2.2 Discrete time

The following result is the finite-dimensional discrete time version of Theorem 2.1.

**Theorem 2.2** *Let  $\{\mathbf{X}_k\}_{k=1,2,\dots}$  be a centered, stationary Gaussian  $R^d$ -valued sequence with the spectral density  $\mathbf{F}(s) = [F_{i,j}(s)]$  such that  $\text{ess sup } \|\mathbf{F}(s)\| < \infty$  (where  $\|\mathbf{F}\|$  denotes the operator norm associated with the matrix  $\mathbf{F}$ , c.f. (28) below). Then for every nonnegative definite symmetric real matrix  $\mathbf{W}$ ,  $\{n^{-1} \sum_{j=1}^n \langle \mathbf{X}_j | \mathbf{W} \mathbf{X}_j \rangle\}$  satisfies the Large Deviation Principle with the rate function*

$$I(x) = \sup_{-\infty < y < 1/(2M)} \{xy - L(y)\},$$

where  $M = \text{ess sup } \|\mathbf{W}^{1/2} \mathbf{F}(s) \mathbf{W}^{1/2}\|$  and for  $y < 1/(2M)$

$$L(y) = -\frac{1}{4\pi} \int_0^{2\pi} \log \det(I - 2y \mathbf{W} \mathbf{F}(s)) ds. \quad (2)$$

**Remark 2.1** *Clearly, Theorem 2.2 implies that the Large Deviation Principle holds also when  $\mathbf{W}$  is a nonpositive definite symmetric real matrix. However, in Section 2.8 we give an example of  $\mathbf{W}$  that is neither positive definite nor negative definite for which  $L(y) = \infty$  even when all eigenvalues of  $2y \mathbf{W} \mathbf{F}(s)$  are uniformly (in  $s$ ) strictly less than 1.*

The following special case of Theorem 2.2 is of interest.

**Corollary 2.1** *Let  $\{X_k\}_{k=1,2,\dots}$  be a real-valued, centered, stationary Gaussian process with bounded spectral density function  $f(s)$  and  $M = \text{ess sup } f(s)$ . Then  $\{\frac{1}{n} \sum_{j=1}^n X_j^2\}$  satisfies the Large Deviation Principle with the rate function*

$$I(x) = \sup_{-\infty < y < 1/(2M)} \{xy - L(y)\}, \quad (3)$$

where  $M = \text{ess sup } f(s)$  and for  $y < 1/(2M)$

$$L(y) = -\frac{1}{4\pi} \int_0^{2\pi} \log(1 - 2yf(s)) ds. \quad (4)$$

The following result deals with additive functionals that have finite all exponential moments.

**Proposition 2.1** *Suppose  $\{X_k\}_{k=1,2,\dots}$  is a centered, real-valued stationary Gaussian process with the continuous spectral density  $f(s)$  satisfying  $\int_0^{2\pi} \log f(s) ds > -\infty$ . Let  $F : R \rightarrow R$  be a continuous function such that  $\lim_{r \rightarrow \infty} \sup_{\{x: |F(x)| \geq r\}} x^{-2} |F(x)| = 0$ . Then  $\{\frac{1}{n} \sum_{j=1}^n F(X_j)\}$  satisfies the Large Deviation Principle.*

The following corollary follows from Corollary 2.1 if  $p = 2$  and from Proposition 2.1 if  $p < 2$ .

**Corollary 2.2** *Suppose that  $\{X_k\}_{k=1,2,\dots}$  has continuous spectral density satisfying  $\int_0^{2\pi} \log f(s) ds > -\infty$ . If  $p \leq 2$  then  $\{\frac{1}{n} \sum_{j=1}^n |X_j|^p\}$  satisfies the Large Deviation Principle.*

**Remark 2.2** *Theorems 2.1 and 2.2 can be also extended to the multivariate index case (Gaussian random fields on  $R^k$  or  $Z^k$ ). Indeed, [18, Chapter 8] develops the relevant abstract results.*

### 2.3 Unbounded spectral density

A suitably modified variant of the Large Deviation Principle holds true also when the spectral density is unbounded. Namely, taking  $S_n = \sum_{j=1}^n X_j^2$  we shall show that for a certain sequence  $m_n \rightarrow \infty$  random variables  $\{m_n(\frac{1}{n}S_n - E(X_1^2))\}$  satisfy the upper and lower bounds with exponent  $m_n^2/n$ , i.e.,

$$\begin{aligned} -\inf_{x \in A^\circ} I(x) &\leq \liminf_{n \rightarrow \infty} \frac{m_n^2}{n} \log P(m_n(\frac{1}{n}S_n - E(X_1^2)) \in A) \\ &\leq \limsup_{n \rightarrow \infty} \frac{m_n^2}{n} \log P(m_n(\frac{1}{n}S_n - E(X_1^2)) \in A) \leq -\inf_{x \in \bar{A}} I(x), \end{aligned} \quad (5)$$

where  $A^\circ$  and  $\bar{A}$  denote the interior and the closure of a measurable set  $A$  respectively.

**Theorem 2.3** *Suppose that real-valued, centered stationary Gaussian process  $\{X_j\}_{j \geq 1}$  has spectral density function  $f(s) \in L_q(ds)$ , where  $2 < q \leq \infty$ . Let  $\{m_n\}$  be such that  $n^{-1/q}m_n \rightarrow \infty$  (if  $q = \infty$ , assume  $m_n \rightarrow \infty$ ), and  $n^{-1/2}m_n \rightarrow 0$ . Then  $\{m_n(\frac{1}{n}S_n - E(X_1^2))\}$  satisfies the Large Deviation Principle (5) with the rate function*

$$I(x) = \frac{x^2}{2\sigma^2},$$

where

$$\sigma^2 = \frac{1}{\pi} \int_0^{2\pi} f^2(s) ds. \quad (6)$$

**Remark 2.3** *With minor changes in the statement and in the proof, Theorem 2.3 holds true both in the multivariate setup of Theorem 2.2 and in the continuous time setup of Theorem 2.1 with the same  $I(x)$ , but with (6) replaced by*

$$\sigma^2 = \pi^{-1} \int_0^{2\pi} \text{tr}(\mathbf{F}(s))^2 ds \quad (7)$$

in the former case (taking  $\mathbf{W} = \mathbf{I}$ ) and

$$\sigma^2 = 4\pi \int_{-\infty}^{\infty} f^2(s) ds \quad (8)$$

in the latter.

### 2.4 Mixing

The proof of Proposition 2.1 gives also the Large Deviation Principle for unbounded functionals (of not necessarily Gaussian processes) under mixing conditions.

For a subset  $C \subset N$  let  $\mathcal{F}_C = \sigma\{X_j : j \in C\}$ . The following are variants of H-mixing, c.f. [10], [13].

**(H-1)** There are  $C, \ell < \infty$  and  $\alpha > 1$  such that for  $k \geq 1$  and all  $j \leq k$  if  $Y_j \geq 0$  are bounded  $\mathcal{F}_{[a_j, b_j]}$ -measurable and  $a_1 \leq b_1 \leq a_2 \leq \dots$  are such that  $a_{j+1} - b_j \geq \ell$  then

$$|E(\prod_{j=1}^k Y_j)| \leq C^k \prod_{j=1}^k \|Y_j\|_\alpha. \quad (9)$$

(H-2) There are  $C < \infty, 0 \leq \gamma < 1, \delta > 0$  and a non-negative sequence  $\beta(n)$  such that for all  $n$  large enough  $\beta(n) \leq \frac{C}{n \log^{1+\delta} n}$  and for all  $X \in L_\infty(\mathcal{F}_{[0,k]})$  and  $Y \in L_\infty(\mathcal{F}_{[k+\ell, \infty)})$

$$|E(XY) - E(X)E(Y)| \leq \gamma \|X\|_{1+\beta(\ell)} \|Y\|_{1+\beta(\ell)}. \quad (10)$$

Let  $E$  be a separable Banach space with norm  $\|\cdot\|$  and  $\mathbf{S}_n = \sum_{j=1}^n \mathbf{X}_j$ .

**Proposition 2.2** Suppose  $\{\mathbf{X}_k\}_{k=1,2,\dots}$  is stationary  $E$ -valued, satisfies conditions (H-1) and (H-2) above, and

$$E(\exp(\theta \|\mathbf{X}_1\|)) < \infty$$

for all  $\theta > 0$ . Then  $\{\frac{1}{n}\mathbf{S}_n\}$  satisfies the Large Deviation Principle.

Recall that  $\psi$ -mixing coefficients are defined by

$$\psi(n) = \sup \left\{ \frac{|E(XY) - E(X)E(Y)|}{E(X)E(Y)} : X \in L_1(\mathcal{F}_{[0,k]}), Y \in L_1(\mathcal{F}_{[k+n, \infty)}), X > 0, Y > 0, k \geq 1 \right\}$$

(this is equivalent to the usual definition that uses indicator functions for  $X$  and  $Y$ ).

**Corollary 2.3 (compare [5, Theorem 2])** Suppose  $\{\mathbf{X}_k\}_{k=1,2,\dots}$  is a stationary sequence of  $E$ -valued random variables with  $E(\exp(\theta \|\mathbf{X}_1\|)) < \infty$  for all  $\theta > 0$  and  $\psi(n) \rightarrow 0$ . Then  $\{\frac{1}{n}\mathbf{S}_n\}$  satisfies the Large Deviation Principle.

Indeed, it is easily seen that under  $\psi$ -mixing both (H-1) and (H-2) are satisfied with  $\alpha = 1, C = 1 + \psi(\ell), \beta(n) = 0$  and with arbitrary  $\gamma > 0$ .

**Remark 2.4** It is known that if  $\psi(N) < \infty$  for some  $N$  and  $\{X_j\}$  is ergodic-mixing (or in the terminology of [5]  $\psi_-(M) > 0$ ), then  $\psi(n) \rightarrow 0$ , see [3].

## 2.5 Application to ARMA(p,q)

Suppose  $\{X_k\}_{k=1,2,\dots}$  is an ARMA(p, q) sequence, i.e.,  $\{X_k\}$  is the stationary solution of

$$\sum_{i=0}^p \alpha_i X_{n-i} = \sum_{j=0}^q \beta_j \gamma_{n-j}, \quad (11)$$

where  $\{\gamma_j\}$  are i.i.d.  $N(0,1)$  r.v. (Note that sequences  $\{\gamma_j\}$  and  $\{X_j\}$  are dependent; in particular, each  $\gamma_n$  might depend on the whole trajectory of  $\{X_j\}$ )

For  $z \in C$  define polynomials  $p(z) = \sum_{i=0}^p \alpha_i z^i$  and  $q(z) = \sum_{j=0}^q \beta_j z^j$ , where without loss of generality  $\alpha_0 \alpha_p \beta_0 \beta_q \neq 0$ . It is well known (c.f. [23, page 42, Theorem 3]) that if  $p(\cdot)$  has no roots of modulus one, then the stationary solution of (11) exists.

For fixed  $y \in R$ , such that  $y \neq 0$  if  $q > p$ , and  $y \neq \frac{\alpha_0 \alpha_q}{2\beta_0 \beta_q}$  if  $q = p$ , let

$$g_y(z) = p(z)p(1/z) - 2yq(z)q(1/z). \quad (12)$$

Denote by  $U_1(y), \dots, U_k(y)$  the (complex) roots of the equation  $g_y(z) = 0$  (in variable  $z$ ) that have modulus larger than 1. Here, multiple roots are listed separately and it is easy to see that  $k = \max\{p, q\}$ .

Define

$$U_0(y) = 1 \text{ if } q < p,$$

$$U_0(y) = \alpha_0 \alpha_q - 2y\beta_0 \beta_q \text{ if } q = p,$$

$U_0(y) = y$  if  $q > p$ .

Let  $M = \sup_s \left| \frac{q(e^{is})}{p(e^{is})} \right|$ . For  $y < 1/(2M)$ ,  $y \neq 0$  (if  $p < q$ ),  $y \neq \frac{\alpha_0 \alpha_q}{2\beta_0 \beta_q}$  (if  $q = p$ ), let

$$\Phi(y) = \sum_{j=0}^k \log |U_j(y)|. \quad (13)$$

**Theorem 2.4** *Suppose that  $\{X_k\}_{k=1,2,\dots}$  is the stationary solution of (11) and that polynomial  $p(z)$  has no zeros of modulus one. Then  $\{\frac{1}{n} \sum_{j=1}^n X_j^2\}$  satisfies the Large Deviation Principle. Moreover,  $\Phi(y)$  extends continuously to all  $y < 1/(2M)$ , and the rate function is given by the Fenchel-Legendre transform of  $L(y) = \frac{1}{2}\Phi(0) - \frac{1}{2}\Phi(y)$ , i.e.,*

$$I(x) = \sup_{y < 1/(2M)} \{xy + \frac{1}{2}\Phi(y) - \frac{1}{2}\Phi(0)\}.$$

**Remark 2.5** *If  $U_j(y) > 0$  are real, one can write*

$$I(x) = -\frac{1}{2}\Phi(0) - \frac{1}{2} \sum_{j=0}^k \left( y \frac{U_j'(y)}{U_j(y)} - \log U_j(y) \right),$$

where  $y = y(x)$  is the (unique, since  $L'(\cdot)$  is increasing) solution of the equation

$$\sum_{j=0}^k \frac{U_j'(y)}{U_j(y)} = -2x.$$

Thus  $I(x) = \sum_j I_j(x_j)$ , where  $I_j(\cdot)$  are the rate functions corresponding to suitable ARMA(1,1) and  $\sum_j x_j = x$  is the "equal energy" (i.e.,  $I_k'(x_k) = I_j'(x_j)$ ) decomposition of  $x$ .

## 2.6 Normal convergence

Lemmas 4.3 and 4.6 from the proof of the LDP yield the following Central Limit Theorem. At least in the univariate discrete time setup this result is known, see [1, Theorem 2], [17, Theorem 2] for a direct proof (for non-normal convergence, see [25]). Related results for more general processes are given in [4, Theorem 5] and the references therein, c.f. also [23, page 58, Theorem 3].

**Proposition 2.3** (i) *If  $\{X_t\}$  is a real-valued, centered, separable stationary Gaussian process with the spectral density  $f(s) \in L_2(\mathbb{R}, ds) \cap L_1(\mathbb{R}, ds)$ , then  $\frac{1}{\sqrt{T}} \int_0^T (X_t^2 - E(X_0^2)) dt$  is asymptotically normal  $N(0, \sigma)$  as  $T \rightarrow \infty$  with  $\sigma^2$  given by (8).*

(ii) *If  $\{\mathbf{X}_k\}_{k=1,2,\dots}$  is a centered, stationary Gaussian  $\mathbb{R}^d$ -valued sequence with the spectral density  $\mathbf{F}(s) = [F_{i,j}(s)]$ , such that  $\text{tr}(\mathbf{F}(s))^2$  is integrable, then*

$$\frac{1}{\sqrt{n}} \sum_{i=1}^n (\langle \mathbf{X}_i | \mathbf{X}_i \rangle - E(\langle \mathbf{X}_1 | \mathbf{X}_1 \rangle))$$

*is asymptotically normal  $N(0, \sigma)$  as  $n \rightarrow \infty$  with  $\sigma^2$  given by (7).*

## 2.7 Non-centered processes and the LDP for the empirical variance

Many of the results presented above carry over to the case of non-centered stationary Gaussian processes by application of the contraction principle. For concreteness, consider the setup of Corollary 2.1, i.e. let  $\{X_j\}$  be a real-valued centered stationary Gaussian process.

The next proposition deals with the Large Deviation Principle (in  $R^2$ ) for the sequence  $\{n^{-1}\mathbf{S}_n = n^{-1}[\sum_{j=1}^n X_j, \sum_{j=1}^n X_j^2]'\}$ .

**Proposition 2.4** *Suppose that spectral density  $f(\cdot)$  is differentiable. Then  $\{n^{-1}\mathbf{S}_n\}$  satisfies the Large Deviation Principle with the rate function*

$$J(x_1, x_2) = I(x_2 - x_1^2) + \frac{x_1^2}{2f(0)}, \quad (14)$$

where  $I(\cdot)$  is the rate function given by (3), and if  $f(0) = 0$  then  $J(x_1, x_2) = \infty$  for  $x_1 \neq 0$  while  $J(0, x_2) = I(x_2)$ .

Applying the contraction principle (see [13, Theorem 4.2.1]) with respect to the continuous function  $g(x_1, x_2) = x_2 + 2x_1\mu + \mu^2 : R^2 \rightarrow R$ , we see that for a non-centered process  $Y_j = X_j + \mu$ , the sequence  $\{n^{-1}\sum_{j=1}^n Y_j^2 = g(n^{-1}\mathbf{S}_n)\}$  satisfies the Large Deviation Principle (in  $R$ ) with rate function

$$J'(z) = \inf_{\{(x_1, x_2): z=g(x_1, x_2)\}} J(x_1, x_2) = \sup_{y < 1/(2M)} \left\{ zy - \frac{\mu^2 y}{1 - 2yf(0)} - L(y) \right\},$$

where  $M = \text{ess sup } f(s)$  and  $L(y)$  given by (4), compare also [2, page 361]. Similarly, applying the contraction principle with respect to the continuous function  $h(x_1, x_2) = x_2 - x_1^2$  results with the empirical variance of  $\{X_j\}_{j=1}^n$  satisfying the Large Deviation Principle with the rate function  $I(\cdot)$  given by (3) (i.e. the *same* rate as for  $\{n^{-1}\sum_{j=1}^n X_j^2\}$ ).

## 2.8 The empirical autocorrelation vector

For  $j \geq 0$ , let  $S_n^{(j)} = \sum_{k=1}^{n-j} X_k X_{k+j}$ . Then  $n^{-1}S_n^{(j)}$  is the  $j$ -th empirical autocorrelation based on the sample of size  $n$ . For fixed  $d \geq 1$  let  $\mathbf{S}_n = [S_n^{(0)}, \dots, S_n^{(d)}] \in R^{d+1}$ . If  $f(\cdot)$  is the spectral density of  $\{X_j\}$ , denote

$$\mathbf{f}(s) = [f(s), f(s) \cos s, \dots, f(s) \cos sd]' \in R^{d+1}.$$

**Proposition 2.5** *Suppose that  $\{X_k\}_{k=1,2,\dots}$  are i.i.d.  $N(0,1)$  random variables. Then  $\{\frac{1}{n}\mathbf{S}_n\}$  satisfies the Large Deviation Principle with the rate function*

$$I(\mathbf{x}) = \sup\{\langle \mathbf{x} | \mathbf{y} \rangle - L(\mathbf{y}) : \mathbf{y} \in D\},$$

where

$$D = \{\mathbf{y} \in R^{d+1} : \sup_{0 \leq s \leq 2\pi} \langle \mathbf{y} | \mathbf{f}(s) \rangle < 1/2\},$$

and for  $\mathbf{y} \in D$

$$L(\mathbf{y}) = -\frac{1}{4\pi} \int_0^{2\pi} \log(1 - 2\langle \mathbf{y} | \mathbf{f}(s) \rangle) ds.$$

**Remark 2.6** *The proof of Proposition 2.5 (with the same formula for the rate function) extends to any differentiable spectral density  $f(s)$  provided that for all  $\mathbf{y} \in D$*

$$\limsup_{n \rightarrow \infty} n^{-1} \log E(\exp(\langle \mathbf{y} | \mathbf{S}_n \rangle)) < \infty. \quad (15)$$

However, the example below shows that for  $d = 1$  and for every AR(1) process with  $0 < |a| < 1$ , (15) is false for some  $\mathbf{y} \in D$ . Hence, in these cases even if  $\{\frac{1}{n}\mathbf{S}_n\}$  satisfies the Large Deviation Principle, the rate function cannot be given by the expression as in Proposition 2.5.

**Example 2.1** Let  $X_k$  be an AR(1) process (with  $\beta_0 = 1$ ,  $\beta_1 = 0$  and  $0 < |a| < 1$ ) corresponding to  $r_i = E[X_0 X_i] = a^i / (1 - a^2)$  for  $i = 0, 1, \dots$  and  $f(s) = 1 / (1 + a^2 - 2a \cos s)$ . Therefore  $\mathbf{y} = \lambda[1 + a^2, -2a] \in D$  for every  $\lambda < 1/2$ . Let  $\mathbf{R}_n$  denote the covariance matrix of  $\mathbf{X} = [X_1, \dots, X_n]'$  and let  $\mathbf{Y}_n$  be the  $n \times n$  symmetric Toeplitz matrix corresponding to  $y_0 = \lambda(1 + a^2)$ ,  $y_1 = -\lambda a$  and  $y_i = 0$  for all  $1 < i \leq n - 1$ . Since  $\mathbf{R}_n^{-1}[r_0, \dots, r_{n-1}]' = [1, 0, \dots, 0]'$ , we have for  $\lambda > (1 - a^2)/2$  and all  $n$  large enough

$$\langle [r_0, \dots, r_{n-1}] | (\mathbf{R}_n^{-1} - 2\mathbf{Y}_n)[r_0, \dots, r_{n-1}]' \rangle = r_0 - 2\lambda(1 + a^2) \sum_{i=0}^{n-1} r_i^2 + 4\lambda a \sum_{i=0}^{n-2} r_i r_{i+1} < 0,$$

implying that  $E(\exp(\lambda(1 + a^2)S_n^{(0)} - 2\lambda a S_n^{(1)})) = \infty$  (see Lemma 4.1).

Note that the above expression is related to Theorem 2.2. Indeed,

$$\lambda(1 + a^2)(S_n^{(0)} - \gamma X_n^2 - (1 - \gamma)X_1^2) - 2\lambda a S_n^{(1)} = \sum_{j=1}^{n-1} \langle \mathbf{X}_j | \mathbf{W}_\gamma \mathbf{X}_j \rangle$$

where  $\mathbf{X}_j = [X_j, X_{j+1}]' \in \mathbb{R}^2$  and

$$\mathbf{W}_\gamma = \lambda \begin{bmatrix} \gamma(1 + a^2) & -a \\ -a & (1 - \gamma)(1 + a^2) \end{bmatrix}.$$

Considering  $\lambda \geq 0$ ,  $\mathbf{W}_\gamma$  is nonnegative definite iff  $\gamma \in [a^2 / (1 + a^2), 1 / (1 + a^2)]$ . For this range of  $\gamma$  it follows by applying Lemma 4.6 to  $\mathbf{Y}_j = \mathbf{W}_\gamma^{1/2} \mathbf{X}_j$  that for all  $\lambda < 1/2$ ,

$$\lim_{n \rightarrow \infty} n^{-1} \log E(\exp(\lambda(1 + a^2)(S_n^{(0)} - \gamma X_n^2 - (1 - \gamma)X_1^2) - 2\lambda a S_n^{(1)})) = -\frac{1}{2} \log(1 - 2\lambda). \quad (16)$$

It can also be verified that for every  $\gamma > 1 / (1 + a^2)$  the left side of (16) is infinite for some  $\lambda \in (0, 1/2)$ , while the eigenvalues of  $\mathbf{W}_\gamma \mathbf{F}(s)$  (which are 0 and  $\lambda$ ) are independent of  $\gamma$ .

**Remark 2.7** The example shows that the large deviations of the empirical autocorrelation vector are sensitive to boundary effects (the choice of  $\gamma$ ), and that Theorem 2.2 does not extend to matrices  $\mathbf{W}$  which are neither nonnegative definite nor nonpositive definite.

## 2.9 Exact asymptotic

The following result comes essentially from [18, page 76]. Together with saddle point approximation, it can be used to find higher order asymptotic expansions for probabilities of "regular enough" sets in Corollary 2.1. We do not pursue this possibility here.

**Proposition 2.6** Suppose  $\{X_k\}_{k \geq 1}$  is a centered, real-valued stationary Gaussian sequence with bounded spectral density  $f(s)$  and  $M = \text{ess sup } f(s)$ . Let  $S_n = \sum_{k=1}^n X_k^2$  and  $L(y)$  be defined by (4). Then for all  $y < 1 / (2M)$  the sequence  $\{\exp(-nL(y))E(\exp(yS_n))\}$  is monotonically nonincreasing. If in addition  $f(s)$  is differentiable and for some  $\alpha > 0$  the function  $f'(s)$  is uniformly Lipschitz continuous with exponent  $\alpha$  then

$$\lim_{n \rightarrow \infty} \exp(-nL(y))E(\exp(yS_n)) = \exp(L(y) - \frac{1}{2\pi} \int \int_{|z| \leq 1} |h'_y(z)|^2 d\sigma),$$

where

$$h_y(z) = \frac{1}{4\pi} \int_0^{2\pi} \log(1 - 2yf(s)) \frac{1 + ze^{-is}}{1 - ze^{-is}} ds,$$

and  $\sigma(dz)$  is the surface measure on the unit disc in  $\mathbb{C}$ .



### 3 Examples

Below we collect several examples with explicit rate function.

**Example 3.1 (Ornstein-Uhlenbeck process)** Suppose  $X_t$  is the stationary solution to  $dX_t = -aX_t + \sqrt{a}dW_t$ ,  $a > 0$ . The spectral density is  $f(s) = \frac{1}{\pi} \frac{a}{a^2 + s^2}$ . Expression (1) from Theorem 2.1 can be integrated, giving for  $y < a/4$

$$L(y) = \frac{1}{2}a - \frac{1}{2}\sqrt{a^2 - 4ay}.$$

Therefore for  $x > 0$

$$I(x) = \frac{a}{4} \left( \sqrt{x} - \frac{1}{\sqrt{x}} \right)^2$$

and  $I(x) = \infty$  otherwise.

**Example 3.2 (Narrow-band noise; continuous time)** Suppose  $X_t$  has spectral density  $f(s) = \frac{\sigma^2}{W}$  for  $s$  in a (symmetric about 0) set of Lebesgue measure  $W$  and 0 otherwise. By Theorem 2.1 the Large Deviation Principle holds for  $T^{-1}S_T$  and for  $y < W/(4\pi\sigma^2)$

$$L(y) = -\frac{W}{4\pi} \log(1 - 4\pi\sigma^2 y/W). \quad (17)$$

Therefore, for  $x > 0$

$$I(x) = \frac{W}{4\pi} \left( \frac{x}{\sigma^2} - 1 - \log\left(\frac{x}{\sigma^2}\right) \right) \quad (18)$$

and  $I(x) = \infty$  otherwise.

**Example 3.3 (Narrow-band noise; discrete time)** Suppose  $X_k$  has spectral density  $f(s) = \frac{2\pi\sigma^2}{W}$  for  $s$  in a (symmetric about 0) set of Lebesgue measure  $W$  and 0 otherwise. Then by Corollary 2.1 the LDP holds, with  $L(y)$  and the rate function  $I(x)$  given by (17) and (18), respectively.

**Example 3.4 (ARMA(1,1))** Suppose  $\{X_k\}$  solves the recurrence

$$X_{n+1} = aX_n + \beta_1\gamma_n + \beta_0\gamma_{n+1},$$

where  $\{\gamma_j\}$  are i.i.d.  $N(0,1)$  r.v. and  $|a| \neq 1$ .

Then  $\{X_k\}$  is ARMA(1,1) with  $p(z) = 1 - az$ ,  $q(z) = \beta_0 + \beta_1z$  and

$$g_y(z) = (1 + a^2 - 2y(\beta_0^2 + \beta_1^2)) - (a + 2y\beta_0\beta_1)(z + 1/z).$$

Here  $U_0(y) = a + 2y\beta_0\beta_1$  and

$$U_1(y) = \frac{1 + a^2 - 2y(\beta_0^2 + \beta_1^2)}{2(a + 2y\beta_0\beta_1)} + \frac{\sqrt{((1+a)^2 - 2y(\beta_0 - \beta_1)^2)((1-a)^2 - 2y(\beta_0 + \beta_1)^2)}}{2(a + 2y\beta_0\beta_1)}.$$

Therefore

$$\Phi(y) = \log\left(\frac{1 + a^2 - 2y(\beta_0^2 + \beta_1^2)}{2}\right) + \frac{\sqrt{((1+a)^2 - 2y(\beta_0 - \beta_1)^2)((1-a)^2 - 2y(\beta_0 + \beta_1)^2)}}{2}$$

Since  $\Phi(0) = 0 \vee \log a^2$ , we get

$$L(y) = -\frac{1}{2} \log\left(\frac{1 + a^2 - 2y(\beta_0^2 + \beta_1^2) + \sqrt{((1+a)^2 - 2y(\beta_0 - \beta_1)^2)((1-a)^2 - 2y(\beta_0 + \beta_1)^2)}}{2(1 \vee a^2)}\right).$$

To find the rate function explicitly, one needs to solve the resulting quartic equation and choose its correct root. Therefore, below we list only special cases when this can be avoided.

(i) Explicit rate function for AR(1), compare [7] (choose  $\beta_0 = 1, \beta_1 = 0, |a| \neq 1$ ). Then

$$L(y) = -\frac{1}{2} \log\left(\frac{1 + a^2 - 2y + \sqrt{(1 + a^2 - 2y)^2 - 4a^2}}{2 \vee (2a^2)}\right).$$

and for  $x > 0$

$$I(x) = \frac{1}{2} \log \frac{1 + \sqrt{4a^2x^2 + 1}}{2x(1 \vee a^2)} + \frac{1}{2}(a^2 + 1)x - \frac{1}{2}\sqrt{4a^2x^2 + 1}.$$

(ii) Explicit rate function for the Moving Average of two r.v.(choose  $a = 0$ ). Then

$$L(y) = -\frac{1}{2} \log\left(\frac{1 - 2y(\beta_0^2 + \beta_1^2) + \sqrt{(1 - 2y(\beta_0 - \beta_1)^2)(1 - 2y(\beta_0 + \beta_1)^2)}}{2}\right)$$

In particular, if  $\beta_1 = \beta_0 = 1$ , we get

$$L(y) = -\frac{1}{2} \log \frac{1 - 4y + \sqrt{1 - 8y}}{2} = \log \frac{2}{1 + \sqrt{1 - 8y}}$$

and for  $x > 0$

$$I(x) = \frac{x}{16} \sqrt{1 + \frac{16}{x}} + \frac{x - 8}{16} + \log \frac{1 + \sqrt{1 + \frac{16}{x}}}{4}.$$

**Example 3.5 (i.i.d. example for Proposition 2.5)** Suppose  $X_i$  are  $N(0,1)$  i.i.d. (i.e.,  $f(s) = 1$ ). Then by Proposition 2.5 the LDP holds for  $[\frac{1}{n} \sum_{k=1}^n X_k^2, \frac{1}{n} \sum_{k=1}^{n-1} X_k X_{k+1}]'$ . The calculations done in Example 3.4 give

$$L(y_0, y_1) = -\frac{1}{2} \log\left(\frac{1 - 2y_0 + \sqrt{(1 - 2y_0)^2 - 4y_1^2}}{2}\right)$$

where

$$D = \{(y_0, y_1) : |y_1| < 1/2 - y_0\}.$$

Therefore one gets the rate function

$$I(x_0, x_1) = \frac{x_0 - 1}{2} + \frac{1}{2} \log\left(\frac{x_0}{x_0^2 - x_1^2}\right)$$

if  $x_0 > 0$  and  $|x_1| < x_0$  ( $I(x_0, x_1) = \infty$  otherwise).

## 4 Proofs

We shall need the following well known elementary result.

**Lemma 4.1** Suppose  $\mathbf{X} = [X_1, \dots, X_n]'$  is a real valued centered Gaussian vector with the covariance matrix  $\mathbf{R}$  and let  $\mathbf{M}$  be a symmetric real valued  $n \times n$ -matrix. Then with  $\lambda_1, \dots, \lambda_n$  the eigenvalues of the matrix  $\mathbf{MR}$

$$\log E \exp(z(\mathbf{X}|\mathbf{MX})) = -\frac{1}{2} \sum_{j=1}^n \log(1 - 2z\lambda_j)$$

for  $z \in \mathbb{C}$  such that  $\max_j \{\operatorname{Re}(z)\lambda_j\} < 1/2$ . Furthermore,  $\log E \exp(y(\mathbf{X}|\mathbf{MX})) = \infty$  for  $y \in \mathbb{R}$  such that  $\max_j \{y\lambda_j\} \geq 1/2$ .

With  $\mathbf{X} = \mathbf{R}^{1/2}\mathbf{Z}$  and  $\mathbf{Z}$  a standard multivariate normal, Lemma 4.1 follows by direct integration of the density of  $\mathbf{Z}$ .

**Lemma 4.2** If  $\{Y_j\}$  are i.i.d. r.v. with mean zero, finite second moment and positive probability density function at 0, then for each  $\theta > 0$  there is  $\delta > 0$  such that

$$\inf\{P(|\sum_{i=1}^{\infty} k_i Y_i| < \theta) : \sum_i |k_i| \leq 1\} \geq \delta.$$

*Proof:* Denote  $\sigma^2 = E(Y^2)$  and fix the sequence  $\{k_i\}$ . Without loss of generality, we may assume that  $|k_i| \geq |k_{i+1}|$  for all  $i \geq 1$ . Note that then the condition  $\sum_j |k_j| \leq 1$  implies that  $|k_j| \leq 1/j$  for all  $j \geq 1$ . Consequently, for every  $r \geq 1$  by Chebyshev's inequality we have

$$P(|\sum_{i=r}^{\infty} k_i Y_i| < \theta) \geq 1 - \frac{\sigma^2}{\theta^2} \sum_{j=r}^{\infty} \frac{1}{j^2}. \quad (19)$$

Note that one can find  $r_0 = r_0(\theta)$  such that the right hand side of (19) is strictly positive. Choose now such  $r_0(\theta/2)$ . By independence we have

$$P(|\sum_{i=1}^{\infty} k_i Y_i| < \theta) \geq P(|\sum_{i=1}^{r_0} k_i Y_i| < \theta/2) P(|\sum_{i=r_0}^{\infty} k_i Y_i| < \theta/2)$$

and, since  $|k_i| \leq 1$ , using (19) we get

$$\begin{aligned} P(|\sum_{i=1}^{\infty} k_i Y_i| < \theta) &\geq P(\max_{1 \leq i \leq r_0} |Y_i| < \theta/(2r_0)) P(|\sum_{i=r_0}^{\infty} k_i Y_i| < \theta/2) \\ &\geq P(|Y_1| < \theta/(2r_0))^{r_0} \left(1 - \frac{4\sigma^2}{\theta^2} \sum_{j=r_0}^{\infty} \frac{1}{j^2}\right) =: \delta. \end{aligned}$$

This ends the proof with  $\delta > 0$  as defined above.  $\square$

## 4.1 Proof of Theorem 2.1

For complex  $z$  with  $\operatorname{Re}(z) < \frac{1}{4\pi M}$ , let  $L_T(z) = \log E(\exp(zS_T))$ .

The following Lemma was motivated by a heuristic argument in [2].

**Lemma 4.3** Under the assumptions of Theorem 2.1, for  $\operatorname{Re}(z) < \frac{1}{4\pi M}$  we have

$$\lim_{T \rightarrow \infty} \frac{1}{T} L_T(z) = -\frac{1}{4\pi} \int_{-\infty}^{\infty} \log(1 - 4\pi z f(s)) ds.$$

*Proof:* For  $T > 0$ , denote by  $\lambda_j = \lambda_j(T)$  the eigenvalues of

$$\int_0^T R(t-s)g(s)ds = \lambda g(t) \in L_2([0, T]) \quad (20)$$

and let  $e_j = e_j(t) \in L_2([0, T], dt)$  be the corresponding orthonormal eigenfunctions. Since by Mercer's theorem,  $R(t-s) = \sum_j \lambda_j e_j(t)e_j(s)$  with positive and summable eigenvalues  $\{\lambda_j\}$ , we have the Karhunen-Loève expansion  $X_t = \sum_j \sqrt{\lambda_j} \gamma_j e_j(t)$ , where  $\gamma_j$  are i.i.d.  $N(0,1)$ . Note that

$$\sup_j \lambda_j = \sup_{g \in L_2, \|g\|=1} \int_0^T g(t)dt \int_0^T g(u)du \int_{-\infty}^{\infty} e^{i(t-u)s} f(s)ds.$$

Since for  $T < \infty$  each square-integrable  $g(\cdot)$  is integrable, we may switch the order of integration, which gives

$$\sup_j \lambda_j \leq M \int_{-\infty}^{\infty} \left| \int_0^T g(t)e^{its} dt \right|^2 ds = 2\pi M, \quad (21)$$

where the last equality is by Plancherel's theorem. Therefore  $Re(z) < 1/(4\pi M) \leq 1/(2\lambda_j)$  and

$$1/T \log E[\exp(zS_T)] = -1/(2T) \sum_{j=1}^{\infty} \log(1 - 2z\lambda_j). \quad (22)$$

Let  $\mu_T(dx) = 1/T \sum_j \delta_{\lambda_j}(dx)$  be the distribution of the eigenvalues on  $[0, 2\pi M]$ . Fix  $z$  and choose  $\delta > 0$  such that  $2|z|\delta < 1$  and such that  $\{s : 2\pi f(s) = \delta\}$  is of Lebesgue measure zero. By [18, page 139] for  $k = 1, 2, \dots$  we have

$$\lim_{T \rightarrow \infty} \int_0^{2\pi M} x^k \mu_T(dx) = (2\pi)^{k-1} \int_{-\infty}^{\infty} f^k(s)ds, \quad (23)$$

and also for every bounded continuous  $F(\cdot)$

$$\lim_{T \rightarrow \infty} \int_{\delta}^{2\pi M} F(x) \mu_T(dx) = \frac{1}{2\pi} \int_{\{s: 2\pi f(s) \geq \delta\}} F(2\pi f(s)) ds. \quad (24)$$

Let  $P_k(x)$  be the  $k$ -th Taylor polynomial for  $x \mapsto \log(1 - 2zx)$ . Notice that from (23) and (24), for each fixed  $k$  we get

$$\int_0^{\delta} P_k(x) \mu_T(dx) \rightarrow \frac{1}{2\pi} \int_{\{s: 2\pi f(s) \leq \delta\}} P_k(2\pi f(s)) ds. \quad (25)$$

Clearly, for  $0 \leq x \leq \delta$  we have

$$|P_k(x) - \log(1 - 2zx)| = \left| \sum_{j=k+1}^{\infty} (2zx)^j / j \right| < \frac{1}{k} \frac{(2x|z|)^{k+1}}{1 - 2|z|\delta} \leq \frac{1}{k} \frac{2x|z|}{1 - 2|z|\delta}.$$

Given  $\epsilon > 0$  choose  $k > 2|z|(1 - 2|z|\delta)^{-1}\epsilon^{-1}$ . Then by (25) choose  $T_0 = T_0(k)$  such that for all  $T > T_0$  we have

$$\left| \int_0^{\delta} P_k(x) \mu_T(dx) - \frac{1}{2\pi} \int_{\{s: 2\pi f(s) \leq \delta\}} P_k(2\pi f(s)) ds \right| < \epsilon$$

and by (23) (with  $k=1$ )

$$\int_0^{2\pi M} x \mu_T(dx) < 2R(0).$$

Enlarging  $T_0$  if necessary, by (24) we may also ensure

$$\left| \int_{\delta}^{2\pi M} \log(1 - 2zx) \mu_T(dx) - \frac{1}{2\pi} \int_{\{s: 2\pi f(s) \geq \delta\}} \log(1 - 4\pi z f(s)) ds \right| < \epsilon$$

for all  $T > T_0$ . Therefore for all  $T > T_0$  we have

$$\begin{aligned} & \left| \int_0^{2\pi M} \log(1 - 2zx) \mu_T(dx) - \frac{1}{2\pi} \int_{-\infty}^{\infty} \log(1 - 4\pi z f(s)) ds \right| \\ & \leq \left| \int_{\delta}^{2\pi M} \log(1 - 2zx) \mu_T(dx) - \frac{1}{2\pi} \int_{\{s: 2\pi f(s) \geq \delta\}} \log(1 - 4\pi z f(s)) ds \right| \\ & \quad + \left| \int_0^{\delta} P_k(x) \mu_T(dx) - \frac{1}{2\pi} \int_{\{s: 2\pi f(s) \leq \delta\}} P_k(2\pi f(s)) ds \right| \\ & \quad + \epsilon \int_0^{2\pi M} x \mu_T(dx) + \epsilon \int_{-\infty}^{\infty} f(s) ds < (2 + 3R(0))\epsilon. \end{aligned}$$

□

**Remark 4.1** By the induced convergence for analytic functions, from Lemma 4.3 it follows that for  $y < \frac{1}{4\pi M}$

$$T^{-1} \frac{d}{dy} L_T(y) \rightarrow \frac{d}{dy} L(y) = \int_{-\infty}^{\infty} \frac{f(s)}{1 - 4\pi y f(s)} ds$$

(this can be also verified directly using [18, page 139]).

**Remark 4.2** Let  $\lambda_1(T)$  be the maximal eigenvalue of (20). Then  $\lambda_1(T) \leq 2\pi M$  by (21), and therefore by [18, page 139] one has  $\lambda_1(T) \rightarrow 2\pi M$  as  $T \rightarrow \infty$ .

*Proof of Theorem 2.1:* By Remark 4.2 and Lemma 4.1 it follows that  $L(y) = \lim_{T \rightarrow \infty} T^{-1} L_T(y)$  is infinite for  $y > 1/(4\pi M)$ , and by Lemma 4.3  $L(y)$  exists and given by (1) for all  $y < 1/(4\pi M)$ . Define  $L(1/(4\pi M)) = \lim_{y \nearrow 1/(4\pi M)} L(y)$  (which by monotone convergence coincides with  $L(1/(4\pi M))$  of (1)), and note that by the monotonicity of  $L_T(y)$  with respect to  $y$

$$\liminf_{T \rightarrow \infty, y_T \rightarrow 1/(4\pi M)} T^{-1} L_T(y_T) \geq L(1/(4\pi M)). \quad (26)$$

By [13, Theorem 2.3.6], the result follows immediately if  $L(1/(4\pi M)) = \infty$ , for then (26) holds with equality, and  $L(\cdot)$  is steep, i.e.,  $\lim_{y \nearrow 1/(4\pi M)} \frac{d}{dy} L(y) = \infty$ .

Checking the proof of Gärtner-Ellis Theorem in [13, Theorem 2.3.6] (see also [13, Theorem 4.5.1]) we have the following (even if  $T^{-1} L_T(1/(4\pi M))$  fails to converge).

(a) The upper bound holds on  $(-\infty, \infty)$  with the rate function

$$I(x) = \sup_{y < 1/(4\pi M)} \{xy - L(y)\},$$

(b) Excluding the trivial case of zero spectral density, since  $L'(y) > 0$  is non-decreasing, there is  $c > 0$  such that  $L'(y) \rightarrow c$  as  $y \nearrow 1/(4\pi M)$ . Then, the lower bound holds on  $(-\infty, c)$  with same rate function  $I(x)$ .

Consequently, if  $L(\cdot)$  is steep (as is the case for example when  $f(s)$  is differentiable), then the proof of the Large Deviation Principle is complete (even for  $L(1/(4\pi M)) < \infty$ ). In case  $L(\cdot)$  is not steep, i.e.  $c < \infty$ , it is simple to check that for  $x \geq c$  the rate function is given by

$$I(x) = \frac{x}{4\pi M} - L\left(\frac{1}{4\pi M}\right).$$

Therefore to prove the Large Deviation Principle it suffices to establish the lower bound

$$\liminf_{T \rightarrow \infty} \frac{1}{T} \log P(|T^{-1}S_T - x| < \epsilon) \geq -I(x + \epsilon) \quad (27)$$

for all  $x > c$  and all  $\epsilon > 0$  small enough. Indeed, (27) gives the large deviations lower bound for small enough open balls centered at  $x > c$ . Since  $I(x)$  is convex and finite on  $(0, \infty)$ , it is continuous at  $x = c$  and hence the lower bound extends to open balls centered at  $x = c$ .

The proof of (27) follows the strategy of  $T$ -dependent change of measure as in [13, Exercise 2.3.24].

Let  $\lambda_1(T) \geq \lambda_2(T) \geq \dots \geq \lambda_n(T) \geq \dots$  be the eigenvalues of (20) and let

$$k_j = \frac{\lambda_j}{T(1 - 2y_T\lambda_j)}$$

where  $y_T \rightarrow 1/(4\pi M)$  is such that  $\sum_{j=1}^{\infty} k_j = x$ .

To see why such  $y_T$  exist note that  $T^{-1} \frac{d}{dy} L_T(y) = T^{-1} \sum_j \lambda_j / (1 - 2y\lambda_j)$  is monotone in  $y$  and approaches  $\infty$  as  $y$  approaches  $1/(2\lambda_1)$ . Thus  $\sum_{j=1}^{\infty} k_j = x$  has the unique solution  $y_T < 1/(2\lambda_1(T))$  and  $\limsup_T y_T \leq 1/(4\pi M)$  by Remark 4.2. Moreover, for each fixed  $y < 1/(4\pi M)$ , by Remark 4.1  $\lim_T T^{-1} \frac{d}{dy} L_T(y) = \frac{d}{dy} L(y) \leq c < x$ ; hence  $y_T \rightarrow 1/(4\pi M)$  as claimed.

We take the above sequence  $y_T \rightarrow 1/(4\pi M)$  and do a change of measure via the Radon-Nikodym derivative

$$\frac{dQ_T}{dP} = \exp(y_T S_T - L_T(y_T)).$$

For large enough  $T$ , we have  $y_T \geq 0$ . Therefore

$$\begin{aligned} T^{-1} \log P(|T^{-1}S_T - x| < \epsilon) &= T^{-1} \log \left( \int \frac{dP}{dQ_T} 1_{|T^{-1}S_T - x| < \epsilon} dQ_T \right) \\ &\geq T^{-1} \log Q_T(|T^{-1}S_T - x| < \epsilon) - (y_T(x + \epsilon) - T^{-1}L_T(y_T)) \end{aligned}$$

By (26) to end the proof we only need a uniform in  $T$  estimate from below on

$$Q_T(|T^{-1}S_T - x| < \epsilon)$$

for all  $\epsilon > 0$ .

Let  $V_T$  denote the r.v.  $(T^{-1}S_T - x)$  under measure  $Q_T$ . Note that by (22) the Laplace transform of  $V_T$  for our choice of  $y_T$  is given by

$$E[e^{sV_T}] = \prod_{i=1}^{\infty} \exp(-sk_i) / \sqrt{1 - sk_i}.$$

Therefore  $V_T$  has the representation

$$V_T = \sum_{j=1}^{\infty} k_j (Z_j^2 - 1)$$

with  $Z_j$  i.i.d. normal  $N(0,1)$ .

The theorem now follows from Lemma 4.2, which we use with  $Y_j = x(Z_j^2 - 1)$  and  $\theta = \epsilon$ .  $\square$

**Remark 4.3** *Formally, one might expect that the value of  $L(1/(4\pi M))$  given by (1) equals the limit of  $T^{-1} \log E(\exp(S_T/(4\pi M)))$  as  $T \rightarrow \infty$ . Indeed, this is true when  $L(1/(4\pi M)) = \infty$ , but otherwise it is not clear that the limit even exists and our proof of Theorem 2.1 circumvents this point.*

## 4.2 Proof of Theorem 2.2

Throughout this proof we consider  $R^n$ ,  $n \geq 1$  as Hilbert subspaces of  $\ell_2$  with the inherited norms. For an  $n \times n$ -matrix  $\mathbf{A}$ , we consider the usual operator norm

$$\|\mathbf{A}\| = \sup_{\mathbf{y} \in R^n \setminus \mathbf{0}} \frac{\|\mathbf{A}\mathbf{y}\|}{\|\mathbf{y}\|}, \quad (28)$$

and the Hilbert-Schmidt norm

$$|\mathbf{A}| = \sqrt{\text{tr}(\mathbf{A}\mathbf{A}')}$$

(with the usual convention that  $\mathbf{A}'$  is the conjugate transpose of the matrix  $\mathbf{A}$ ). It is well known that  $|\mathbf{ABC}| \leq \|\mathbf{A}\| \cdot |\mathbf{B}| \cdot \|\mathbf{C}\|$ , and that  $\|\mathbf{A}\| \leq |\mathbf{A}|$ , see e.g. [16, Section XI.6]. We shall also use the elementary bound  $\text{tr} \mathbf{A} \leq n^{1/2}|\mathbf{A}| \leq n\|\mathbf{A}\|$ .

The distribution of the eigenvalues  $\{\lambda_1, \dots, \lambda_n\}$  of  $\mathbf{A}$  is the discrete probability measure

$$\mu_n(dx) = n^{-1} \sum_{j=1}^n \delta_{\lambda_j}(dx)$$

(either on  $R$  or on  $C$ , depending on whether  $\mathbf{A}$  is symmetric, or not).

Consider now two sequences of matrices  $\{\mathbf{A}_n\}$  and  $\{\mathbf{B}_n\}$ . The following result is known and a short proof is enclosed for completeness.

**Lemma 4.4** ([18, p 105]) *Suppose the  $n \times n$  matrices  $\mathbf{A}_n$  and  $\mathbf{B}_n$  have the distribution of the eigenvalues  $\mu_n$  and  $\nu_n$  respectively and assume that*

$$\sup_n (\|\mathbf{A}_n\| + \|\mathbf{B}_n\|) < \infty, \quad (29)$$

and

$$\lim_{n \rightarrow \infty} n^{-1} |\mathbf{A}_n - \mathbf{B}_n|^2 = 0. \quad (30)$$

Then  $\lim_{n \rightarrow \infty} |\int x^k \mu_n(dx) - \int x^k \nu_n(dx)| = 0$  for every  $k = 1, 2, \dots$

*Proof:*

$$\begin{aligned} |\int x^k \mu_n(dx) - \int x^k \nu_n(dx)| &= n^{-1} |\text{tr}(\mathbf{A}_n^k - \mathbf{B}_n^k)| \leq n^{-1/2} |\mathbf{A}_n^k - \mathbf{B}_n^k| \\ &= n^{-1/2} \left| \sum_{j=1}^k \mathbf{A}_n^{k-j} (\mathbf{A}_n - \mathbf{B}_n) \mathbf{B}_n^{j-1} \right| \\ &\leq n^{-1/2} |\mathbf{A}_n - \mathbf{B}_n| k \max\{\|\mathbf{A}_n\|^{k-1}, \|\mathbf{B}_n\|^{k-1}\}. \end{aligned}$$

□

Let  $\mathbf{R}_n = \text{cov}(\mathbf{X}_0, \mathbf{X}_n)$  be the  $d \times d$ -covariance matrices, and consider the block-Toeplitz  $nd \times nd$  matrix

$$\mathbf{A}_n = \begin{bmatrix} \mathbf{R}_0 & \mathbf{R}_1 & \dots & \mathbf{R}_{n-1} \\ \mathbf{R}'_1 & \mathbf{R}_0 & \dots & \mathbf{R}_{n-2} \\ \vdots & & \ddots & \vdots \\ \mathbf{R}'_{n-1} & \mathbf{R}'_{n-2} & \dots & \mathbf{R}_0 \end{bmatrix}.$$

Let  $\mu_n$  be the distribution of the eigenvalues of  $\mathbf{A}_n$ . The asymptotic of  $\mu_n$  follows by extending the argument of [18, page 113] to the  $d$ -dimensional matrix case as follows.

**Lemma 4.5** *If  $M = \text{ess sup} \|\mathbf{F}(s)\| < \infty$  then  $\sup_n \|\mathbf{A}_n\| \leq M$ . Moreover, for any  $a < b$  such that  $m(s : \lambda_j(s) = a) = m(s : \lambda_j(s) = b) = 0$  for  $j = 1, \dots, d$ ,*

$$\lim_{n \rightarrow \infty} \mu_n([a, b]) = (2\pi d)^{-1} \sum_{j=1}^d m(s : a < \lambda_j(s) < b), \quad (31)$$

where  $m$  is Lebesgue measure on  $[0, 2\pi]$  and  $\lambda_1(s) \geq \lambda_2(s) \geq \dots \geq \lambda_d(s) \geq 0$  are the eigenvalues of  $\mathbf{F}(s)$  (recall that  $\mathbf{F}(s), 0 \leq s \leq 2\pi$ , are Hermitian, nonnegative definite matrices).

*Proof:* For  $(n-1)/2 \geq A \geq 1$  let

$$\mathbf{B}_{n,A} = \begin{bmatrix} \widehat{\mathbf{R}}_0 & \widehat{\mathbf{R}}_1 & \dots & \widehat{\mathbf{R}}_A & 0 & \dots & 0 \\ \widehat{\mathbf{R}}_1' & \widehat{\mathbf{R}}_0 & \dots & \widehat{\mathbf{R}}_{A-1}' & \widehat{\mathbf{R}}_A & \ddots & 0 \\ \vdots & \ddots & \ddots & & & \ddots & 0 \\ \widehat{\mathbf{R}}_A' & & & & & & \widehat{\mathbf{R}}_A \\ 0 & \ddots & & & \ddots & & \vdots \\ \vdots & & \ddots & & \ddots & \widehat{\mathbf{R}}_0 & \widehat{\mathbf{R}}_1 \\ 0 & \dots & 0 & \widehat{\mathbf{R}}_A' & \dots & \widehat{\mathbf{R}}_1' & \widehat{\mathbf{R}}_0 \end{bmatrix}$$

be an  $nd \times nd$ -matrix, where  $\widehat{\mathbf{R}}_k = (1 - k/A)\mathbf{R}_k$  for  $k = 0, \dots, A$  and  $\widehat{\mathbf{R}}_k = 0$  for  $k > A$  (with  $\widehat{\mathbf{R}}_{-k} = \widehat{\mathbf{R}}_k'$ ). Let  $\mathbf{C}_{n,A}$  be the block-circulant matrix associated with  $\mathbf{B}_{n,A}$ , defined as follows.

$$\mathbf{C}_{n,A} = \begin{bmatrix} \widehat{\mathbf{R}}_0 & \widehat{\mathbf{R}}_1 & \dots & \widehat{\mathbf{R}}_A & 0 & \dots & 0 & \widehat{\mathbf{R}}_A' & \dots & \widehat{\mathbf{R}}_1' \\ \widehat{\mathbf{R}}_1' & \widehat{\mathbf{R}}_0 & \dots & \widehat{\mathbf{R}}_{A-1}' & \widehat{\mathbf{R}}_A & 0 & & \ddots & & \vdots \\ \vdots & & \ddots & & \ddots & \ddots & & & \ddots & \widehat{\mathbf{R}}_A' \\ \widehat{\mathbf{R}}_A' & & & \ddots & & & & & & 0 \\ 0 & \ddots & & & & & & & & \vdots \\ \vdots & & & & & & & & \ddots & 0 \\ 0 & & & & & & & & & \widehat{\mathbf{R}}_A \\ \widehat{\mathbf{R}}_A & \ddots & & & & & & \ddots & & \vdots \\ \vdots & & \ddots & & & \ddots & & \ddots & & \widehat{\mathbf{R}}_1 \\ \widehat{\mathbf{R}}_1 & \dots & \widehat{\mathbf{R}}_A & 0 & \dots & 0 & \widehat{\mathbf{R}}_A' & \dots & \widehat{\mathbf{R}}_1' & \widehat{\mathbf{R}}_0 \end{bmatrix}.$$

Let  $\mathbf{F}_A(s) = \sum_{k=-A}^A e^{-iks} \widehat{\mathbf{R}}_k$ , with  $\{\lambda_{j,k}\}_{j=1,\dots,d}$  denoting the eigenvalues of  $\mathbf{F}_A(2\pi k/n)$ ,  $k = 0, \dots, n-1$  and  $\mathbf{v}_{j,k} \in \mathbb{R}^d$  the corresponding eigenvectors. The usual argument for circulant matrices shows that for  $j = 1, \dots, d, k = 0, \dots, n-1$  the  $nd$ -dimensional vectors

$$(\mathbf{v}_{j,k}, e^{2\pi ik/n} \mathbf{v}_{j,k}, \dots, e^{2\pi ik(n-1)/n} \mathbf{v}_{j,k})$$

are the linearly independent eigenvectors of  $\mathbf{C}_{n,A}$  corresponding to the eigenvalues  $\lambda_{j,k}$ ; therefore those are all the eigenvalues of  $\mathbf{C}_{n,A}$ . Consequently,  $\|\mathbf{C}_{n,A}\| \leq \sup_s \|\mathbf{F}_A(s)\|$  and since

$$\mathbf{F}_A(s) = \frac{1}{2\pi} \int_0^{2\pi} \frac{\sin^2(A(s-t)/2)}{A \sin^2((s-t)/2)} \mathbf{F}(t) dt,$$



clearly,

$$\sup_s \|\mathbf{F}_A(s)\| \leq \sup_s \|\mathbf{F}(s)\| \cdot \frac{1}{2\pi} \int_0^{2\pi} \frac{\sin^2(At/2)}{A \sin^2(t/2)} dt = M. \quad (32)$$

We turn now to prove that  $\|\mathbf{A}_n\| \leq M$  and  $\|\mathbf{B}_{n,A}\| \leq M$ . To this end, fix  $n$ , pick  $\mathbf{x}_j \in R^d$  and write  $\mathbf{x} = (\mathbf{x}_j)$  as a column vector. Then,

$$\begin{aligned} \langle \mathbf{x} | \mathbf{A}_n \mathbf{x} \rangle &= (2\pi)^{-1} \int_0^{2\pi} \left\langle \sum_{k=1}^n e^{-iks} \mathbf{x}_k | \mathbf{F}(s) \sum_{m=1}^n e^{ims} \mathbf{x}_m \right\rangle ds \\ &\leq (2\pi)^{-1} \int_0^{2\pi} \left\| \sum_{k=1}^n e^{-iks} \mathbf{x}_k \right\|^2 \|\mathbf{F}(s)\| ds \\ &\leq \sup_{0 \leq s \leq 2\pi} \|\mathbf{F}(s)\| \left( \frac{1}{2\pi} \int_0^{2\pi} \left\| \sum_{k=1}^n e^{-iks} \mathbf{x}_k \right\|^2 ds \right) = M \|\mathbf{x}\|^2. \end{aligned}$$

By a similar argument we have for  $n > A$

$$\begin{aligned} \langle \mathbf{x} | \mathbf{B}_{n,A} \mathbf{x} \rangle &= (2\pi)^{-1} \int_0^{2\pi} \left\langle \sum_{k=1}^n e^{-iks} \mathbf{x}_k | \mathbf{F}_A(s) \sum_{m=1}^n e^{ims} \mathbf{x}_m \right\rangle ds \\ &\leq \|\mathbf{x}\|^2 \sup_s \|\mathbf{F}_A(s)\| \leq M \|\mathbf{x}\|^2. \end{aligned}$$

This shows that matrices  $\mathbf{A}_n$  and  $\mathbf{B}_{n,A}$  and  $\mathbf{C}_{n,A}$  satisfy (29) for every choice of  $A \leq (n-1)/2$ . By applying Parseval's relation elementwise one has

$$\sum_{j=-\infty}^{\infty} |\mathbf{R}_j|^2 = (2\pi)^{-1} \int_0^{2\pi} |\mathbf{F}(s)|^2 ds \leq d M^2.$$

Since for every  $n > A$  we have

$$n^{-1} |\mathbf{A}_n - \mathbf{B}_{n,A}|^2 \leq 2 \sum_{j=1}^A (j/A)^2 |\mathbf{R}_j|^2 + 2 \sum_{j=A+1}^{\infty} |\mathbf{R}_j|^2,$$

by Kronecker's Lemma it follows that  $n^{-1} |\mathbf{A}_n - \mathbf{B}_{n,A}|^2$  can be made arbitrarily small (uniformly in  $n > A$ ) by choosing  $A$  large enough. Therefore, by choosing first  $A$  large and then  $n$  large enough, we can make sure that (30) holds both for  $|\mathbf{A}_n - \mathbf{B}_{n,A}|$  and for  $|\mathbf{B}_{n,A} - \mathbf{C}_{n,A}|$  since

$$|\mathbf{B}_{n,A} - \mathbf{C}_{n,A}|^2 \leq 2A \sum_{j=1}^A |\mathbf{R}_j|^2 \leq Ad M^2.$$

Consequently, by Lemma 4.4 the asymptotic of  $\mu_n$  is the same as the asymptotic of the distribution of the eigenvalues of  $\mathbf{C}_{n,A}$  provided we let  $n \rightarrow \infty$  first and then take  $A \rightarrow \infty$ .

Fix a positive integer  $\ell$ . In view of the continuity of  $\mathbf{F}_A(s)$  we have for any fixed  $A \geq 1$

$$\lim_{n \rightarrow \infty} n^{-1} \sum_{k=1}^n \text{tr} (\mathbf{F}_A(2\pi k/n)^\ell) = \frac{1}{2\pi} \int_0^{2\pi} \text{tr} (\mathbf{F}_A(s)^\ell) ds.$$

Also

$$\left| (2\pi)^{-1} \int_0^{2\pi} \text{tr} (\mathbf{F}_A(s)^\ell - \mathbf{F}(s)^\ell) ds \right|^2 \leq d (2\pi)^{-1} \int_0^{2\pi} |\mathbf{F}_A(s)^\ell - \mathbf{F}(s)^\ell|^2 ds$$

$$\leq d\ell^2 M^{2(\ell-1)} (2\pi)^{-1} \int_0^{2\pi} |\mathbf{F}_A(s) - \mathbf{F}(s)|^2 ds ,$$

and since,

$$(2\pi)^{-1} \int_0^{2\pi} |\mathbf{F}_A(s) - \mathbf{F}(s)|^2 ds = 2 \sum_{j=1}^A (j/A)^2 |\mathbf{R}_j|^2 + 2 \sum_{j=A+1}^{\infty} |\mathbf{R}_j|^2 ,$$

we have for  $A \rightarrow \infty$  that  $\int_0^{2\pi} \text{tr} (\mathbf{F}_A(s)^\ell - \mathbf{F}(s)^\ell) ds \rightarrow 0$ , leading to

$$\lim_{A \rightarrow \infty} \lim_{n \rightarrow \infty} n^{-1} \sum_{k=1}^n \text{tr} (\mathbf{F}_A(2\pi k/n)^\ell) = \frac{1}{2\pi} \int_0^{2\pi} \text{tr} (\mathbf{F}(s)^\ell) ds .$$

With the above limit holding for every positive integer  $\ell$ , the limit (31) follows by [18, page 105].  $\square$

Let  $S_n = \sum_{j=1}^n \langle \mathbf{X}_j | \mathbf{X}_j \rangle$  and for complex  $z$ , let  $L_n(z) = \log E(\exp(zS_n))$ .

**Lemma 4.6** *If  $\sup_s \|\mathbf{F}(s)\| = M < \infty$ , then the limit  $\lim_{n \rightarrow \infty} \frac{1}{n} L_n(z)$  exists for every  $z$  in the half-plane  $\text{Re } z < \frac{1}{2M}$  and*

$$\lim_{n \rightarrow \infty} \frac{1}{n} L_n(z) = -\frac{1}{4\pi} \int_0^{2\pi} \log \det(I - 2z\mathbf{F}(s)) ds . \quad (33)$$

**Remark 4.4** *For  $d = 1$  this lemma is known, see [8, page 105], or [9, Example 3.1 a)].*

*Proof:* Clearly,

$$S_n = [\mathbf{X}_1, \dots, \mathbf{X}_n][\mathbf{X}_1, \dots, \mathbf{X}_n]' .$$

Therefore by Lemma 4.1, for  $\text{Re}(z) < 1/(2 \max_j \lambda_j)$

$$n^{-1} L_n(z) = -1/(2n) \sum_{j=1}^{nd} \log(1 - 2z\lambda_j),$$

where  $\{\lambda_j\}$  are the eigenvalues of the symmetric nonnegative definite matrix  $\mathbf{A}_n$ .

Lemma 4.5 implies that  $\max_j \lambda_j = \|\mathbf{A}_n\| \leq M$  for all  $n$ , and by (31) actually  $\|\mathbf{A}_n\| \rightarrow M$  as  $n \rightarrow \infty$ . Consequently, (33) follows by applying (31) and observing that

$$n^{-1} L_n(z) = -\frac{d}{2} \int_0^M \log(1 - 2zx) \mu_n(dx) .$$

$\square$

**Remark 4.5** *By the induced convergence for analytic functions, from Lemma 4.6 it follows that for  $y < 1/(2M)$*

$$n^{-1} \frac{d}{dy} L_n(y) \rightarrow \frac{d}{dy} L(y) = \frac{1}{2\pi} \sum_{j=1}^d \int_0^{2\pi} \frac{\lambda_j(s)}{1 - 2y\lambda_j(s)} ds,$$

where  $\lambda_j(s), j = 1, \dots, d$  are the (nonnegative) eigenvalues of  $\mathbf{F}(s)$ . (This claim can also be verified directly from (31).)

*Proof of Theorem 2.2:* For  $\mathbf{W}$  an identity matrix, the proof repeats the reasoning from the proof of Theorem 2.1. Indeed, by Lemma 4.6,  $n^{-1}L_n(y)$  converges to  $L(y)$  of (2) for  $y < 1/(2M)$ , while by Lemmas 4.1 and 4.5, for  $y > 1/(2M)$

$$L(y) = \lim_{n \rightarrow \infty} n^{-1}L_n(y) = \infty .$$

Excluding the trivial case of zero spectral density, notice that  $L'(y) > 0$  is monotonically increasing for  $y < 1/(2M)$ , and let  $c > 0$  be such that  $L'(y) \rightarrow c$  as  $y \nearrow 1/(2M)$ . Defining  $L(1/(2M)) = \lim_{y \nearrow 1/(2M)} L(y)$ , if  $L(y)$  is steep, i.e.  $c = \infty$ , then the proof of Gärtner-Ellis Theorem in [13, Theorem 2.3.6] (see also [13, Theorem 4.5.1]) yields the Large Deviation Principle even if  $n^{-1}L_n(1/(2M))$  fails to converge. If  $L(\cdot)$  is not steep then for  $x \geq c$ , the rate function is given by  $I(x) = \frac{x}{2M} - L(\frac{1}{2M})$ . Letting  $\{\lambda_j\}$  denote the nonnegative eigenvalues of the matrix  $\mathbf{A}_n$ , the  $n$ -dependent change of measure via  $\frac{dQ_n}{dP} = \exp(y_n S_n - L_n(y_n))$  results with  $n^{-1}S_n - x$  (under  $Q_n$ ) having the representation  $\sum_{j=1}^{nd} k_j(Z_j^2 - 1)$  with  $Z_j$  i.i.d. normal  $N(0, 1)$  and  $k_j = \lambda_j/(n(1 - 2y_n\lambda_j))$ , where  $y_n < 1/(2 \max_j \lambda_j)$  chosen such that  $\sum_{j=1}^{nd} k_j = x$ . Since  $\max_j \{\lambda_j\} = \|\mathbf{A}_n\| \rightarrow M$  as  $n \rightarrow \infty$  it follows by Remark 4.5, that  $\lim_n y_n = 1/(2M)$  and the proof of the large deviations lower bound for  $x > c$  is completed by applying Lemma 4.2 (note that  $\liminf_n n^{-1}L_n(y_n) \geq L(1/(2M))$ ). For any  $\mathbf{W}$  nonnegative definite symmetric real matrix, we have  $\mathbf{W} = \mathbf{W}^{1/2}\mathbf{W}^{1/2}$  with  $\mathbf{W}^{1/2}$  also nonnegative symmetric real matrix. Hence  $\langle \mathbf{X}_j | \mathbf{W} \mathbf{X}_j \rangle = \langle \mathbf{Y}_j | \mathbf{Y}_j \rangle$  for  $j = 1, 2, \dots$ , where  $\mathbf{Y}_j = \mathbf{W}^{1/2} \mathbf{X}_j$  is a stationary process of bounded spectral density  $\mathbf{W}^{1/2} \mathbf{F}(s) \mathbf{W}^{1/2}$ . Therefore, the general case follows by applying the above proof to the process  $\{\mathbf{Y}_j\}$ .  $\square$

**Remark 4.6** For  $d = 1$ , by Lemma 4.1 and [18, pages 38, 44],  $n^{-1} \log E(\exp((2M)^{-1} \sum_{j=1}^n X_j^2))$  converges as  $n \rightarrow \infty$  to  $L(1/(2M))$  of (4). The validity of this result in the general context of Theorem 2.2 is not addressed here.

### 4.3 Proof of Theorem 2.3

The proof is based on the Gärtner-Ellis Theorem (c.f. [13, Theorem 2.3.6 and Remark (a)]) used with the normalization  $a_n = m_n^2/n \rightarrow 0$ .

We shall need the following estimate for the maximal eigenvalue of the covariance matrices.

**Lemma 4.7** *If  $1 \leq q \leq \infty$  then there is  $C < \infty$  such that for all  $n > 1$  if  $\mathbf{A}_n$  is the covariance matrix of  $[X_1, \dots, X_n]'$  then  $\|\mathbf{A}_n\| \leq Cn^{1/q}$ .*

*Proof:* Let  $\mathbf{x} = [x_1, \dots, x_n]'$  be such that  $\|\mathbf{x}\| = 1$  and  $\|\mathbf{A}_n\| = \langle \mathbf{x} | \mathbf{A}_n \mathbf{x} \rangle$ . Then, denoting  $1/p + 1/q = 1$ , we have  $\|\mathbf{A}_n\| = \frac{1}{2\pi} \int_0^{2\pi} f(s) |\sum x_j e^{ijs}|^2 ds \leq \|f\|_q (\frac{1}{2\pi} \int_0^{2\pi} |\sum x_j e^{ijs}|^{2p} ds)^{1/p} \leq C(\sum |x_j|)^{(2p-2)/p} \leq Cn^{1/q}$ .  $\square$

*Proof of Theorem 2.3:* Denote  $T_n = m_n(\frac{1}{n}S_n - EX_1^2)$  and as previously, let  $\lambda_j = \lambda_j(n)$ ,  $1 \leq j \leq n$ , be the eigenvalues of the covariance of  $X_1, \dots, X_n$ . Since by Lemma 4.7 and the choice of  $m_n$   $\max_j \lambda_j/m_n \rightarrow 0$ , for every  $y \in R$  and for all  $n \geq n_0(y)$  we have

$$\log E \exp(nm_n^{-2}yT_n) = -ynm_n^{-1}EX_1^2 - \frac{1}{2} \sum_{j=1}^n \log(1 - 2y\lambda_j/m_n) .$$

Notice that by Taylor's Theorem for  $|w| < 1$

$$\log(1 - w) = -w - (1/2)w^2(1 - tw)^{-2},$$

where  $t = t(w) \in [0, 1]$ . This is applied here to  $w_j = 2y\lambda_j/m_n$  which by Lemma 4.7 satisfies  $\sup_j |w_j| \rightarrow 0$  as  $n \rightarrow \infty$ , and hence,  $|1 - t(w_j)w_j| \rightarrow 1$  uniformly in  $1 \leq j \leq n$ . This shows that the limit of

$$m_n^2 n^{-1} \log E \exp(nm_n^{-2} y T_n)$$

is the same as that of

$$ym_n(n^{-1} \sum_{j=1}^n \lambda_j - E(X_1^2)) + y^2 n^{-1} \sum_{j=1}^n \lambda_j^2$$

Clearly,  $\sum_{j=1}^n \lambda_j = \text{tr } \mathbf{A}_n = nE(X_1^2)$ , and

$$n^{-1} \sum_{j=1}^n \lambda_j^2 = n^{-1} \text{tr } \mathbf{A}_n^2 = \sum_{k=-(n-1)}^{n-1} (1 - |k|/n) r_k^2 = \sum_{k=-(n-1)}^{n-1} r_k^2 - 2 \sum_{k=1}^{n-1} (k/n) r_k^2$$

Notice that by Parseval's identity  $\sum_{k=-(n-1)}^{n-1} r_k^2 \rightarrow \sum_{k=-\infty}^{\infty} r_k^2 = \sigma^2/2$  as  $n \rightarrow \infty$ . On the other hand, by Kronecker's Lemma  $\sum_{k=1}^{n-1} (k/n) r_k^2 \rightarrow 0$  as  $n \rightarrow \infty$  leading to

$$\lim_{n \rightarrow \infty} m_n^2 n^{-1} \log E \exp(nm_n^{-2} y T_n) = \frac{1}{2} y^2 \sigma^2.$$

This ends the proof by the Gärtner-Ellis Theorem.  $\square$

#### 4.4 Proof of Propositions 2.1 and 2.2

Let  $E$  be a separable Banach space.

**Lemma 4.8** *Suppose  $\{\mathbf{X}_k\}$  are  $E$ -valued such that*

(a) *for every  $M > 0$  there exists an  $E$ -valued sequence  $\mathbf{Y}_k^M$  such that  $\{n^{-1} \sum_{k=1}^n \mathbf{Y}_k^M\}$  satisfies the Large Deviation Principle*

(b) *for each  $\theta > 0$*

$$\limsup_{M \rightarrow \infty} \limsup_{n \rightarrow \infty} \frac{1}{n} \log E(\exp(\theta \sum_{k=1}^n \|\mathbf{X}_k - \mathbf{Y}_k^M\|)) \leq K(\theta), \quad (34)$$

*and  $K(\theta)/\theta \rightarrow 0$  as  $\theta \rightarrow \infty$ .*

*Then  $\{n^{-1} \sum_{k=1}^n \mathbf{X}_k\}$  satisfies the Large Deviation Principle if this sequence is exponentially tight.*

**Remark 4.7** *If  $\|\mathbf{Y}_k^M\| \leq M$  for all  $M$ , then inequality (34) implies that for all  $\theta > 0$*

$$\limsup_{n \rightarrow \infty} \frac{1}{n} \log E(\exp(\theta \sum_{k=1}^n \|\mathbf{X}_k\|)) < \infty,$$

*which in finite dimensional case implies that the sequence  $\{n^{-1} \sum_{k=1}^n \mathbf{X}_k\}$  is exponentially tight.*

*Proof:* Denote  $\mathbf{S}_n = \sum_{k=1}^n \mathbf{X}_k$ ,  $\mathbf{S}_n^M = \sum_{k=1}^n \mathbf{Y}_k^M$ ,  $\mathbf{T}_n^M = \mathbf{S}_n - \mathbf{S}_n^M$ . By assumption (and Varadhan's Integral Theorem, see e.g. [13, Theorem 4.3.1]), for every bounded continuous  $F : E \rightarrow R$ , the limit

$$L_M(F) = \lim_{n \rightarrow \infty} n^{-1} \log E(\exp(nF(n^{-1} \mathbf{S}_n^M)))$$

exists. We shall show that for bounded above Lipschitz  $F(\cdot)$  the limit

$$L(F) = \lim_{n \rightarrow \infty} n^{-1} \log E(\exp(nF(n^{-1}\mathbf{S}_n)))$$

exists and  $L(F) = \lim_{M \rightarrow \infty} L_M(F)$ . This will end the proof by [13, Theorem 4.4.10].

Since  $F(\cdot)$  is Lipschitz, by Hölder's inequality we have for  $\delta > 0$

$$\begin{aligned} E(\exp(nF(n^{-1}\mathbf{S}_n))) &\leq E(\exp(nF(n^{-1}\mathbf{S}_n^M) + C_1\|\mathbf{T}_n^M\|)) \\ &\leq (E(\exp(n(1+\delta)F(n^{-1}\mathbf{S}_n^M))))^{1/(1+\delta)} (E(\exp((1+\delta)C_1/\delta\|\mathbf{T}_n^M\|)))^{\delta/(1+\delta)} \end{aligned}$$

Therefore

$$\begin{aligned} &\limsup_{n \rightarrow \infty} n^{-1} \log E(\exp(nF(n^{-1}\mathbf{S}_n))) \\ &\leq \frac{1}{1+\delta} L_M((1+\delta)F) + \frac{\delta}{1+\delta} \limsup_{n \rightarrow \infty} \frac{1}{n} \log E(\exp((1+\delta)C_1/\delta \sum_{k=1}^n \|\mathbf{X}_k - \mathbf{Y}_k^M\|)). \end{aligned}$$

Passing to the limit as  $M \rightarrow \infty$  we get by (34) with  $\theta = (1+\delta)C_1/\delta$ ,

$$\limsup_{n \rightarrow \infty} n^{-1} \log E(\exp(nF(n^{-1}\mathbf{S}_n))) \leq \frac{1}{1+\delta} \liminf_{M \rightarrow \infty} L_M((1+\delta)F) + \frac{\delta}{1+\delta} K((1+\delta)C_1/\delta)$$

Since  $L_M((1+\delta)F) \leq \delta \sup_x F(x) + L_M(F)$  we now pass to the limit as  $\delta \rightarrow 0$ , proving that

$$\limsup_{n \rightarrow \infty} n^{-1} \log E(\exp(nF(n^{-1}\mathbf{S}_n))) \leq \liminf_{M \rightarrow \infty} L_M(F).$$

The opposite bound

$$\liminf_{n \rightarrow \infty} n^{-1} \log E(\exp(nF(n^{-1}\mathbf{S}_n))) \geq \limsup_{M \rightarrow \infty} L_M(F)$$

is produced analogously, starting with

$$E \exp\left(\frac{1}{1+\delta} nF(n^{-1}\mathbf{S}_n^M)\right) \leq E \exp\left(\frac{1}{1+\delta} nF(n^{-1}\mathbf{S}_n) + \frac{C_1}{1+\delta} \|\mathbf{T}_n^M\|\right).$$

□

*Proof of Proposition 2.1 :* For truncated  $F(X_j)$ , Assumption (a) of Lemma 4.8 follows from [15, Theorem 2.3].

With  $b_r = \sup_{\{x: |F(x)| \geq r\}} x^{-2}|F(x)|$ , by Lemma 4.6 we have for all  $r \geq r_0(\theta)$

$$\limsup_{n \rightarrow \infty} \frac{1}{n} \log E(\exp(\theta \sum_{j=1}^n (|F(X_j)| - r) I_{\{|F(X_j)| \geq r\}})) \leq L(\theta b_r),$$

where  $L(\cdot)$  is given by (4). Passing to the limit as  $r \rightarrow \infty$  we see that inequality (34) of Lemma 4.8 holds with  $K(\theta) = 1$ , and by Remark 4.7 the sequence  $\{n^{-1} \sum_{j=1}^n F(X_j)\}$  is exponentially tight. □

*Proof of Proposition 2.2 :* For  $\mathbf{Y}_k^M = \mathbf{X}_k I_{\{\|\mathbf{X}_k\| \leq M\}}$ , Assumption (a) of Lemma 4.8 follows from [13, proof of Lemma 6.4.6]. Indeed, it suffices to show that if  $g : E \rightarrow R$  is concave and Lipschitz (with constant denoted  $\|g\|$ ) and  $g(0) = 0$  then  $\lim_{n \rightarrow \infty} n^{-1} \log E(\exp(n g(n^{-1}\mathbf{S}_n^M)))$  exists. Denote  $x_n = -\log E(\exp(n g(n^{-1}\mathbf{S}_n^M)))$ . The argument [13, proof of Lemma 6.4.6] shows that

$$x_{n+m} \leq x_n + x_m + 2\ell M \|g\| - \log \left( 1 - \gamma e^{M \|g\| (n+m) \frac{\beta(\ell)}{1+\beta(\ell)}} \right).$$

Also,  $|x_n| \leq M\|g\|n < \infty$ . Therefore, choosing  $\ell = \ell(n+m) = (n+m)\log^{1+\delta/2}(n+m)$  by [19],  $\{n^{-1}x_n\}$  has the finite limit.

Inequality (34) of Lemma 4.8 with  $K(\theta) = C$  follows from the proof of [13, Lemma 6.4.18(a)], see also [10, (1.1)]. This can be seen as follows. First, notice that by the standard approximation argument used for each fixed  $k \geq 1$ , inequality (9) holds true for any  $\alpha$ -integrable r.v.  $Y_j$ . Indeed, suppose  $Y_j > 0$  are  $\alpha$ -integrable and take non-negative bounded  $Y_j^N \uparrow Y_j$  as  $N \rightarrow \infty$ . Then by monotone convergence theorem,  $E(Y_1 \cdots Y_k) = \lim_{N \rightarrow \infty} E(Y_1^N \cdots Y_k^N) \leq C^k \lim_{N \rightarrow \infty} \prod_{j=1}^k \|Y_j^N\|_\alpha = C^k \prod_{j=1}^k \|Y_j\|_\alpha$ .

Let  $\mathbf{Z}_j^M = \mathbf{X}_j - \mathbf{Y}_j^M$ . Then by Hölder's inequality

$$\begin{aligned} E(\exp(\theta \sum_{j=1}^n \|\mathbf{Z}_j^M\|)) &\leq E(\exp(\theta \sum_{j=1}^{n\ell} \|\mathbf{Z}_j^M\|)) \\ &\leq E(\exp(\ell\theta \sum_{j=1}^n \|\mathbf{Z}_{j\ell}^M\|)) \leq C^n (E(\exp(\ell\theta\alpha \|\mathbf{Z}_1^M\|)))^n \end{aligned} \quad (35)$$

and since  $E(\exp(\ell\theta\alpha \|\mathbf{X}_1\|)) < \infty$  we have  $E(\exp(\ell\theta\alpha \|\mathbf{Z}_1^M\|)) \rightarrow 1$  as  $M \rightarrow \infty$ .

The inequality (35) holds also for arbitrary seminorm in place of  $\|\cdot\|$ . Therefore, exponential tightness follows from inequality (35) applied to  $M = 0$  and with the  $\|\cdot\|$  replaced by a seminorm  $q: E \rightarrow [0, \infty)$ , where  $q(\cdot)$  is such that  $q^{-1}[0, 1]$  is compact and  $E \exp(q(\mathbf{X}_1)) < \infty$  (such seminorm exists by [12, Theorem 3.1]). Indeed, for  $\theta = (\ell\alpha)^{-1}$  we have

$$n^{-1} \log P(q(n^{-1}\mathbf{S}_n) > N) \leq \log C + \log(E(\exp(\ell\theta\alpha q(\mathbf{X}_1)))) - \theta N,$$

which can be made arbitrarily small by choosing  $N$  large enough.  $\square$

## 4.5 Proof of Theorem 2.4

We shall need the following elementary Lemma.

**Lemma 4.9** *Suppose  $a$  is a complex number*

(i) *If  $|a| > 1$  then  $\frac{1}{2\pi i} \int_{|z|=1} \frac{\log(|a-z|)}{z} dz = \log |a|$ .*

(ii) *If  $|a| < 1$  then  $\frac{1}{2\pi i} \int_{|z|=1} \frac{\log(|1-a/z|)}{z} dz = 0$ .*

*Proof:* Take a suitable analytic branch of the logarithm. Then

(i') *If  $|a| > 1$  then  $\frac{1}{2\pi i} \int_{|z|=1} \frac{\log(a-z)}{z} dz = \log a$ .*

(ii') *If  $|a| < 1$  then  $\frac{1}{2\pi i} \int_{|z|=1} \frac{\log(1-a/z)}{z} dz = 0$ .*

Indeed, since in case (i')  $z \mapsto \log(a-z)$  is analytic in the unit disc, (i') follows from the Cauchy integral formula; (ii') follows from the fact that for  $|a| < 1$  the mapping  $z \mapsto 1/z \log(1-a/z)$  is analytic at  $\infty$ .

The formulas given in the statement of the lemma come from taking the real part of the above complex identities; it is obvious that the integrals in the statement of the lemma are real, since for  $z = e^{is}$ ,  $dz/(iz) = ds$ .  $\square$

The following lemma connects formulas (33) and (13).

**Lemma 4.10** *Under the assumption of Theorem 2.4, for  $y < 1/(2M)$ ,  $y \neq 0$ ,  $y \neq \frac{\alpha_0 \alpha_p}{\beta_0 \beta_q}$  let*

$$\Phi_1(y) = \frac{1}{2\pi i} \int_{|z|=1} \frac{\log(g_y(z))}{z} dz. \quad (36)$$

and let  $\Phi(y)$  be given by (13). Then  $\Phi(y) - \Phi_1(y) = C$  is a constant.

*Proof:* Since  $y \neq 0$  and  $y \neq \frac{\alpha_0 \alpha_p}{\beta_0 \beta_q}$ , therefore  $g_y(z)$  can be re-written as the polynomial (with real coefficients) of degree  $k = \max\{p, q\}$  in variable  $z + 1/z$ ; hence equation  $g_y(z) = 0$  has exactly  $2k$  solutions. For  $y < \frac{1}{2M}$ , equation  $g_y(z) = 0$  has no solution of modulus one; therefore there are exactly  $k$  roots (counting multiple roots with their multiplicities) of  $g_y(z) = 0$  that have modulus larger than 1.

Write  $g_y(z)$  as the polynomial (with real coefficients) in variable  $z + 1/z$  and let  $a_1 = a_1(y), \dots, a_k = a_k(y)$  be the (complex) roots of this polynomial. Thus, noting that  $g_y(z) > 0$  when  $|z| = 1$ ,

$$\Phi_1(y) = C + \log |U_0| + \sum_{j=1}^k \frac{1}{2\pi i} \int_{|z|=1} \frac{\log(|z + 1/z - a_j|)}{z} dz$$

For each  $1 \leq j \leq k$ , factor the quadratic expression  $z^2 + 1 - a_j z$  to get  $z + 1/z - a_j = \frac{(z - U_j)(z - V_j)}{z}$ . Since  $U_j V_j = 1$  and  $g_y(z)$  has no roots of modulus one, exactly one of the numbers  $U_j, V_j$  has modulus larger than one and we may assume that the labeling is such that  $|U_j| > 1, 1 \leq j \leq k$ . Clearly,

$$\Phi_1(y) = C + \log |U_0| + \sum_{j=1}^k \frac{1}{2\pi i} \int_{|z|=1} \frac{\log(|z - U_j|)}{z} dz + \sum_{j=1}^k \frac{1}{2\pi i} \int_{|z|=1} \frac{\log(|1 - V_j/z|)}{z} dz,$$

which ends the proof by Lemma 4.9.  $\square$

*Proof of Theorem 2.4:* Since polynomial  $p(z)$  has no complex zeros of modulus one, therefore, see e.g. [23, page 38 formula (9)],  $\{X_j\}$  has differentiable spectral density

$$f(s) = \left| \frac{q(e^{is})}{p(e^{is})} \right|^2$$

and the Large Deviation Principle follows from Corollary 2.1.

The rate function identification goes as follows. Write

$$f(s) = \frac{q(z)q(1/z)}{p(z)p(1/z)},$$

where  $z = e^{is}$ . Clearly,  $1 - 2yf(s) = \frac{g_y(z)}{p(z)p(1/z)}$  and from (4) we have for  $y < 1/(2M)$

$$L(y) = -\frac{1}{4\pi i} \int_{|z|=1} \frac{1}{z} \log\left(\frac{g_y(z)}{p(z)p(1/z)}\right) dz.$$

Therefore for  $y < 1/(2M)$ ,  $y \neq 0$ ,  $y \neq \frac{\alpha_0 \alpha_p}{\beta_0 \beta_q}$  (if  $q = p$ )

$$L(y) = -\frac{1}{2} \Phi_1(y) + C$$

with  $\Phi_1(\cdot)$  defined by (36).

By Lemma 4.10

$$L(y) = -\frac{1}{2} \Phi(y) + \text{const}, \quad (37)$$

where  $\Phi(\cdot)$  is defined by (13). Since  $L(y)$  is continuous, therefore  $\Phi(y)$  extends to a continuous function. Since  $L(0) = 0$ , we get  $\text{const} = \frac{1}{2} \Phi(0)$  and the result follows.  $\square$

## 4.6 Proof of Proposition 2.3

For  $f(s)$  or  $\|\mathbf{F}(s)\|$  bounded, the CLT follows immediately from Lemmas 4.3 and 4.6 by a simple complex analysis argument given in [6, Proposition 1]. In general, for every  $M < \infty$ , we let  $X_t = Y_t + Z_t$  in the continuous time setup and  $\mathbf{X}_k = \mathbf{Y}_k + \mathbf{Z}_k$  in the discrete time setup; in the former case  $Y_t$  and  $Z_t$  are independent, real-valued, centered, separable stationary Gaussian processes with spectral densities  $f_y(s) = \min(f(s), M)$  and  $f_z(s) = f(s) - f_y(s)$ , while in the latter  $\mathbf{Y}_k$  and  $\mathbf{Z}_k$  are independent,  $R^d$ -valued, centered, stationary Gaussian sequences, with the spectral densities  $\mathbf{F}_y(s)$  and  $\mathbf{F}_z(s)$  having the same eigenvectors as  $\mathbf{F}(s)$  but with eigenvalues  $\min(\lambda_j(s), M)$  and  $\max(\lambda_j(s) - M, 0)$  respectively. Then, in the continuous time setup,

$$W_M := \frac{1}{\sqrt{T}} \int_0^T (X_t^2 - Y_t^2 - E(X_0^2 - Y_0^2)) dt = \frac{1}{\sqrt{T}} \int_0^T (Z_t^2 - E(Z_0^2)) dt + \frac{2}{\sqrt{T}} \int_0^T Y_t Z_t dt,$$

has mean zero and variance bounded above by  $\epsilon_M := 4\sigma(4\pi \int_{-\infty}^{\infty} f_z(s)^2 ds)^{1/2}$ , while in the discrete time setup,

$$W_M := \frac{1}{\sqrt{n}} \sum_{i=1}^n (\langle \mathbf{X}_i | \mathbf{X}_i \rangle - \langle \mathbf{Y}_i | \mathbf{Y}_i \rangle - E(\langle \mathbf{X}_0 | \mathbf{X}_0 \rangle - \langle \mathbf{Y}_0 | \mathbf{Y}_0 \rangle)),$$

has zero mean and variance bounded by  $\epsilon_M := 4\sigma(\pi^{-1} \int_0^{2\pi} \text{tr}(\mathbf{F}_z(s)^2) ds)^{1/2}$ . Note that in both cases  $\epsilon_M \rightarrow 0$  as  $M \rightarrow \infty$ , hence for every  $\delta > 0$ , by Chebyshev's inequality  $P(|W_M| > \delta) < \epsilon_M/\delta^2 \rightarrow 0$  as  $M \rightarrow \infty$  uniformly in  $T$  ( $n$ ). Since  $f_y(s)$  is bounded,  $\frac{1}{\sqrt{T}} \int_0^T (Y_t^2 - E(Y_0^2)) dt$  is asymptotically normal  $N(0, \sigma_M)$  as  $T \rightarrow \infty$ , with  $\sigma_M := (4\pi \int_{-\infty}^{\infty} f_y(s)^2 ds)^{1/2}$  monotonically increasing to  $\sigma$  as  $M \rightarrow \infty$ . Similarly, in the discrete time setup,  $\|\mathbf{F}_y(s)\|$  is bounded and hence  $\frac{1}{\sqrt{n}} \sum_{i=1}^n (\langle \mathbf{Y}_i | \mathbf{Y}_i \rangle - E(\langle \mathbf{Y}_0 | \mathbf{Y}_0 \rangle))$  is asymptotically normal  $N(0, \sigma_M)$  as  $n \rightarrow \infty$ , with  $\sigma_M := (\pi^{-1} \int_0^{2\pi} \text{tr}(\mathbf{F}_y(s)^2) ds)^{1/2} \nearrow \sigma$  as  $M \rightarrow \infty$ . The required CLT then follows by the continuity of the normal distribution function.  $\square$

## 4.7 Proof of Proposition 2.4

For  $\mathbf{y} = [y_1, y_2]$  define  $L_n(\mathbf{y}) = \log E \exp(\langle \mathbf{y} | \mathbf{S}_n \rangle)$ . Let  $\mathbf{R}_n$  be the covariance matrix of  $\mathbf{X} = [X_1, \dots, X_n]'$  with  $\lambda_1(n)$  denoting the maximal eigenvalue of  $\mathbf{R}_n$ ,  $\mathbf{I}_n$  denoting the identity matrix, and  $\mathbf{e}_n = [1, 1, \dots, 1]'$ . By adapting the calculations of Lemma 4.1 we have for  $y_2 < 1/(2\lambda_1(n))$

$$L_n(\mathbf{y}) = L_n([0, y_2]) + \frac{1}{2} y_1^2 \langle \mathbf{e}_n | \mathbf{R}_n^{1/2} (\mathbf{I}_n - 2y_2 \mathbf{R}_n)^{-1} \mathbf{R}_n^{1/2} \mathbf{e}_n \rangle$$

(and  $L_n(\mathbf{y}) = \infty$  for all other values of  $\mathbf{y}$ ).

**Lemma 4.11** *If  $y_2 < 1/(2M)$  then*

$$L(\mathbf{y}) = \lim_{n \rightarrow \infty} n^{-1} L_n(\mathbf{y}) = L(y_2) + \frac{y_1^2 f(0)}{2(1 - 2y_2 f(0))},$$

with  $L(y)$  given by (4), and  $n^{-1} L_n(\mathbf{y}) \rightarrow \infty$  when  $y_2 > 1/(2M)$ .

*Proof:* We have by [18, page 65] that  $n^{-1} L_n([0, y_2]) \rightarrow L(y_2)$  for all  $y_2 < 1/(2M)$  and  $n^{-1} L_n(\mathbf{y}) \rightarrow \infty$  for all  $y_2 > 1/(2M)$ . Taking  $y_2 < 1/(2M)$  we have by [18, pages 27, 53, 209] that

$$n^{-1} \langle \mathbf{e}_n | (\mathbf{I}_n - 2y_2 \mathbf{R}_n)^{-1} \mathbf{e}_n \rangle \rightarrow 1/(1 - 2y_2 f(0)),$$

and the proof is completed by noting that  $2y_2 \mathbf{R}_n^{1/2} (\mathbf{I}_n - 2y_2 \mathbf{R}_n)^{-1} \mathbf{R}_n^{1/2} + \mathbf{I}_n = (\mathbf{I}_n - 2y_2 \mathbf{R}_n)^{-1}$ .  $\square$



*Proof of Proposition 2.4:* Defining  $L([y_1, 1/(2M)]) = \lim_{y_2 \nearrow 1/(2M)} L(\mathbf{y})$  and  $L(\mathbf{y}) = \infty$  for  $y_2 > 1/(2M)$ , it is easy to check that  $J(x_1, x_2)$  of (14) is the Fenchel-Legendre transform of  $L(\mathbf{y})$ . The proof of Gärtner-Ellis Theorem in [13, Theorem 2.3.6] then yields the Large Deviation Principle provided  $L(\mathbf{y})$  is steep. To that end, note that for  $y_2 < 1/(2M)$

$$\frac{\partial L(\mathbf{y})}{\partial y_2} \geq \frac{1}{2\pi} \int_0^{2\pi} \frac{f(s)}{1 - 2y_2 f(s)} ds.$$

Hence, by the differentiability of  $f(s)$  we have  $\frac{\partial L(\mathbf{y})}{\partial y_2} \rightarrow \infty$  as  $y_2 \nearrow 1/(2M)$  implying that  $L(\mathbf{y})$  is steep (for more details, see the proof of Proposition 2.5).  $\square$

## 4.8 Proof of Proposition 2.5

Let  $L_n(\mathbf{y}) = \log E(\exp(\langle \mathbf{y} | \mathbf{S}_n \rangle))$  and define  $n \times n$ -matrix

$$\mathbf{Y}_n = \begin{bmatrix} y_0 & \frac{1}{2}y_1 & \dots & \frac{1}{2}y_d & 0 & \dots & 0 \\ \frac{1}{2}y_1 & y_0 & \dots & & \frac{1}{2}y_d & \ddots & \vdots \\ \vdots & & \ddots & & & \ddots & 0 \\ \frac{1}{2}y_d & \dots & & y_0 & \dots & & \frac{1}{2}y_d \\ 0 & & & & & & \\ \vdots & \ddots & & & & \ddots & \vdots \\ 0 & \dots & 0 & \frac{1}{2}y_d & \dots & & y_0 \end{bmatrix}$$

Let  $\mathbf{R}_n$  be the covariance matrix of  $\mathbf{X} = [X_1, \dots, X_n]'$ . Since  $\langle \mathbf{y} | \mathbf{S}_n \rangle = \mathbf{X}' \mathbf{Y}_n \mathbf{X}$ , by Lemma 4.1 we have

$$L_n(\mathbf{y}) = -1/2 \sum_{j=1}^n \log(1 - 2\lambda_j(\mathbf{y})), \quad (38)$$

where  $\lambda_j(\mathbf{y})$  are the eigenvalues of the matrix  $\mathbf{M}_n = \mathbf{Y}_n \mathbf{R}_n$  and  $\mathbf{y}$  is such that  $\max_j \{\lambda_j(\mathbf{y})\} < 1/2$ .

For i.i.d.  $X_j$  we have that  $\mathbf{R}_n$  is the identity matrix, hence  $\mathbf{M}_n = \mathbf{Y}_n$  is the symmetric Toeplitz matrix corresponding to the "signed" bounded spectral density  $\langle \mathbf{y} | \mathbf{f}(s) \rangle$ . In particular, by [18, page 65] for  $\mathbf{y} \in D$

$$\frac{1}{n} L_n(\mathbf{y}) \rightarrow L(\mathbf{y}) = -\frac{1}{4\pi} \int_0^{2\pi} \log(1 - 2\langle \mathbf{y} | \mathbf{f}(s) \rangle) ds.$$

By [18, pages 38, 44] this relation holds also for  $\mathbf{y} \in \partial D$ , i.e. when  $\sup_s \langle \mathbf{y} | \mathbf{f}(s) \rangle = \frac{1}{2}$ , while  $n^{-1} L_n(\mathbf{y}) \rightarrow \infty$  for all other values of  $\mathbf{y}$ .

Notice that if  $\|\mathbf{y}\| < 1/(2(d+1))$  then  $\mathbf{y} \in D$ . Therefore, in order to establish the Large Deviation Principle, we need only to verify the steepness condition, i.e.,

$$\lim_{\mathbf{y} \rightarrow \mathbf{y}_0, \mathbf{y} \in D} \|L'(\mathbf{y})\| = \infty$$

for all  $\mathbf{y}_0 \in \partial D$ , see [13, Theorem 2.3.6]. To this end, fix  $\mathbf{y}_0 \in \partial D$  and let  $0 \leq s_0 \leq 2\pi$  be such that  $\langle \mathbf{y}_0 | \mathbf{f}(s_0) \rangle = 1/2$ . It suffices to show that

$$|\langle \mathbf{y}_0 | L'(\mathbf{y}) \rangle| \rightarrow \infty$$

as  $\mathbf{y} \rightarrow \mathbf{y}_0, \mathbf{y} \in D$ . Clearly,

$$\langle \mathbf{y}_0 | L'(\mathbf{y}) \rangle = \frac{1}{2\pi} \int_0^{2\pi} \frac{\langle \mathbf{y}_0 | \mathbf{f}(s) \rangle}{1 - 2\langle \mathbf{y} | \mathbf{f}(s) \rangle} ds.$$

Let

$$I_+ = \{s : \langle \mathbf{y}_0 | \mathbf{f}(s) \rangle \geq 0\},$$

$$I_- = \{s : \langle \mathbf{y}_0 | \mathbf{f}(s) \rangle < 0\}.$$

We have

$$\limsup_{\mathbf{y} \rightarrow \mathbf{y}_0} \left| \frac{1}{2\pi} \int_{I_-} \frac{\langle \mathbf{y}_0 | \mathbf{f}(s) \rangle}{1 - 2\langle \mathbf{y} | \mathbf{f}(s) \rangle} ds \right| \leq \|\mathbf{y}_0\|.$$

Since  $\mathbf{f}(s)$  is differentiable, for each  $\epsilon > 0$  there is  $\delta > 0$  such that for  $|s - s_0| < \delta$  we have  $|\langle \mathbf{y}_0 | \mathbf{f}(s) \rangle - \langle \mathbf{y}_0 | \mathbf{f}(s_0) \rangle| < \epsilon\delta$  and  $\langle \mathbf{y}_0 | \mathbf{f}(s) \rangle \geq m > 0$ ; i.e.  $(s_0 - \delta, s_0) \subset I_+$  (if  $s_0 = 0$  replace  $(s_0 - \delta, s_0)$  by  $(s_0, s_0 + \delta)$ ). Then

$$\begin{aligned} & \int_{I_+} \frac{\langle \mathbf{y}_0 | \mathbf{f}(s) \rangle}{1 - 2\langle \mathbf{y} | \mathbf{f}(s) \rangle} ds \geq \int_{s_0 - \delta}^{s_0} \frac{\langle \mathbf{y}_0 | \mathbf{f}(s) \rangle}{1 - 2\langle \mathbf{y} | \mathbf{f}(s) \rangle} ds \\ & \geq m \int_{s_0 - \delta}^{s_0} \frac{1}{2\langle \mathbf{y}_0 | \mathbf{f}(s_0) - \mathbf{f}(s) \rangle + 2\langle \mathbf{y}_0 - \mathbf{y} | \mathbf{f}(s) \rangle} ds \geq m \frac{\delta}{2\epsilon\delta + 2\|\mathbf{y}_0 - \mathbf{y}\|} \end{aligned}$$

Therefore  $\liminf_{\mathbf{y} \rightarrow \mathbf{y}_0} \langle \mathbf{y}_0 | L'(\mathbf{y}) \rangle \geq m/(4\pi\epsilon) - \|\mathbf{y}_0\|$ . Since  $\epsilon > 0$  is arbitrary, this ends the proof.  $\square$

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