On the stability problem for conditional expectation

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Abstract: The behavior of the conditional expectation $E\{X|Y\}$ under a small perturbation Z of the conditioning random variable Y is analyzed. We show that if Y and Z are independent then $E\{X|Y+\varepsilon Z\}$ converges to $E\{X|Y\}$ in mean as $\varepsilon \to 0$ for all integrable X, provided the distribution of Y is absolutely continuous. We also show that the limit is $E\{X|Y,Z\}$ rather than $E\{X|Y\}$, i.e., there is no stability, when Y is a discrete (i.e., countably valued) random variable. Finally, we show that in general $E\{X|Y+\varepsilon Z\}$ might have no limit in distribution as $\varepsilon \to 0$.

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1. Introduction

In this note we investigate the stability of the conditional expectation $E\{X \mid Y\}$ under a small additive perturbation by a random variable Z, i.e., we study the behavior of $E\{X \mid Y + \varepsilon Z\}$ as $\varepsilon \to 0$. In the case when Y and Z are independent real valued and the convergence is understood in probability it is possible to answer the stability problem completely. Namely, we show that $E\{X \mid Y + \varepsilon Z\}$ converges to $E\{X \mid Y\}$ as $\varepsilon \to 0$, if Y is absolutely continuous, $E\{X \mid Y + \varepsilon Z\}$ converges to $E\{X \mid Y, Z\}$ as $\varepsilon \to 0$, if Y is discrete, and there is no limit in general. The same question can be asked also when Y and Z are \mathbb{R}^d -valued random variables or when Y and Z are replaced by stochastic processes $\{Y_t\}$; however, the information that we have in those cases is far less complete and is left out of this note.

The following result gives a sufficient condition for the stability of the conditional expectation under independent perturbations.

Theorem 1.1. If Y, Z are independent, the distribution of Y is absolutely continuous with respect to the Lebesgue measure on \mathbb{R} and $E\{|X|\} < \infty$, then

$$E\{X | Y + \varepsilon Z\} \to E\{X | Y\} \quad in \ L_1\{dP\} \ as \ \varepsilon \to 0. \tag{1.1}$$

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Remark 1. Suitable modifications in the proof of Theorem 1 show that the conclusion holds also for dependent random variables Y, Z, provided they have joint density.

Remark 2. In general, to show that (1.1) holds for all integrable X, it is enough to consider all X of the form $X = \phi(Z)$, where $\phi: \mathbb{R} \to \mathbb{R}$ is bounded and uniformly continuous. Indeed, by the standard approximation argument, the finite sums $\Sigma \phi_i(Z) \Psi_i(Y)$, where ϕ_i , Ψ_i are bounded and uniformly continuous, are dense in $L_1(\sigma(Y, Z), dP)$. For a single term of the sum we have

$$E\{\phi(Z)\Psi(Y) | Y + \varepsilon Z\} = \Psi(Y + \varepsilon Z)E\{\phi(Z) | Y + \varepsilon Z\} + \theta_{\varepsilon},$$

where

$$\theta_{\varepsilon} = E\{\phi(Z)(\Psi(Y) - \Psi(Y + \varepsilon Z)) | Y + \varepsilon Z\} \to 0 \text{ in } L_1(dP) \text{ as } \varepsilon \to 0.$$

Therefore $E\{\phi(Z)\Psi(Y)|Y + \varepsilon Z\} \rightarrow \Psi(Y)E\{\phi(Z)|Y\}$, provided (1.1) is established for $X = \phi(Z)$.

Remark 3. The argument given above also shows that (1.1) holds for all $\sigma(Y)$ -measurable integrable X. As a consequence, one easily gets that if Y has a density, $E\{|Z|^p\} < \infty$ for some $p \ge 1$ and $E\{Z\} = 0$, then $E\{Y | Y + \varepsilon Z\} = Y + \varepsilon Z + o(\varepsilon)$, where $||o(\varepsilon)||_p / |\varepsilon| \to 0$ as $\varepsilon \to 0$. Indeed,

 $E\{Y \mid Y + \varepsilon Z\} = Y + \varepsilon Z - \varepsilon E\{Z \mid Y + \varepsilon Z\} \text{ and } E\{Z \mid Y + \varepsilon Z\} \rightarrow E\{Z\}$

by Theorem 1.1. Szablowski (1986, Corollary 1.2) shows that if one conditions Z rather than Y, then the representation

$$E\{Z \mid Y + \varepsilon Z\} = E\{Z\} + \varepsilon \beta(Y + \varepsilon Z) + o(\varepsilon)$$

permits one to determine the distribution of Y, which then has to be absolutely continuous.

The following result shows that there is no stability in the discrete case. No assumption of independence is needed.

Theorem 1.2. If Y is discrete and $E\{|X|\} < \infty$ then

$$E\{X \mid Y + \varepsilon Z\} \to E\{X \mid Y, Z\} \quad in \ L_1(dP) \text{ as } \varepsilon \to 0 \text{ for all } Z.$$

$$(1.2)$$

Remark 4. If in addition either

(i) Z is discrete, or

(ii) Z is bounded and Y takes a finite number of values,

then the convergence in (1.2) is also in the most sure sense.

The following observation complements Theorems 1.1 and 1.2.

Proposition 1.3. Suppose Y, Z are independent and identically distributed. If (1.1) holds for all integrable X, then either Y is continuous (i.e., atomless), or Y is deterministic.

Proposition 1.3 gives a necessary condition for (1.1). The following example shows that there are i.i.d. continuous (but not absolutely continuous) Y, Z such that $E\{X | Y + \varepsilon Z\}$ has no limit in distribution as $\varepsilon \to 0$.

Example. Let ξ_k , η_k be independent {0, 1}-valued random variables. Suppose

 $P(\xi_{2k}=1) = P(\eta_{2k}=1) = p_1$ and $P(\xi_{2k+1}=1) = P(\eta_{2k+1}=1) = p_2$ for all k,

where $0 < p_1 < p_2 < 1$ are fixed. Let

$$X = \xi_1, \qquad Y = \sum_{k=1}^{\infty} \eta_k / 3^k, \qquad Z = \sum_{k=1}^{\infty} \xi_i / 3^k.$$

Clearly, Y and Z are independent and identically distributed. Moreover, from the uniqueness of the ternary expansion in this case we see that

$$\sigma\left\{Y+\frac{1}{3^{n}}Z\right\}=\sigma\{\eta_{1}, \eta_{2}, \ldots, \eta_{n}, \eta_{n+1}+\xi_{1}, \eta_{n+2}+\xi_{2}, \ldots\}.$$

Therefore

$$E\left\{X \mid Y + \frac{1}{3^{n}}Z\right\} = E\left\{\xi_{1} \mid \eta_{n+1} + \xi_{1}\right\} = \begin{cases} 0 & \text{if } \eta_{n+1} + \xi_{1} = 0, \\ \frac{p_{1}(1-q)}{p_{1}+q-2p_{1}q} & \text{if } \eta_{n+1} + \xi_{1} = 1, \\ 1 & \text{if } \eta_{n+1} + \xi_{1} = 2, \end{cases}$$

where $q = p_1$ if n is even and $q = p_2$ if n is odd. In particular, $E\{X | Y + (1/3^n)Z\}$ has no limit in distribution as $n \to \infty$.

2. Proofs

Proof of Theorem 1.1. As it was pointed out in Remark 2, it is enough to establish (1.1) for X = f(Z), where $f(\cdot)$ is a bounded continuous function. Furthermore, since the family of conditional expectations of a fixed random variable X with respect to different σ -fields is uniformly integrable, we need only to show convergence in probability.

Let g(y) be the density of Y. By F_X , F_Y , F_Z we denote the cumulative distribution functions of X, Y, Z; $\lambda(\cdot)$ denotes the Lebesgue measure and $\mu \otimes F$ is the product of measures μ and dF. Clearly,

$$E\{X \mid Y\} = \int f(z) F_Z(\mathrm{d}z)$$

and

$$E\{X | Y + \varepsilon Z\} = \frac{\int f(z)g(Y + \varepsilon Z - \varepsilon z)F_Z(dz)}{\int g(Y + \varepsilon Z - \varepsilon z)F_Z(dz)}$$

Since f is bounded, we have

$$\begin{split} \left| \int f(z)g(Y+\varepsilon Z-\varepsilon z)F_Z(\mathrm{d} z) - \int f(z)g(Y)F_Z(\mathrm{d} z) \right| \\ &\leq \|f\|_{\infty} \int |g(Y+\varepsilon Z-\varepsilon z)-g(Y)|F_Z(\mathrm{d} z). \end{split}$$

Therefore, to show that $E\{X | Y + \varepsilon Z\} \rightarrow E\{X | Y\}$ in probability, it is enough to show that

$$\int |g(Y + \varepsilon Z - \varepsilon z) - g(Y)| F_Z(\mathrm{d}z) \to 0 \quad \text{in probability.}$$
(2.1)

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To this end, for each $\delta > 0$ let $h(\cdot) = h_{\delta}(\cdot)$ be a uniformly continuous probability density function such that

$$\int |g(y) - h(y)| \, \mathrm{d} y < \delta^2.$$

For real ε let

$$A_{\varepsilon} = \left\{ (y, z) \colon \int |g(y + \varepsilon z - \varepsilon u) - h_{\delta}(y + \varepsilon z - \varepsilon u)| F_{Z}(\mathrm{d}u) > \delta \right\}.$$

By Chebyshev's inequality,

$$(\lambda \otimes F_Z)(A_{\varepsilon}) \leq \frac{1}{\delta} \iiint |g(y + \varepsilon z - \varepsilon u) - h(y + \varepsilon z - \varepsilon u)| dy F_Z(du) F_Z(dz) < \delta.$$

Let $\rho > 0$ be arbitrary. Since $F_Y \otimes F_Z$ is absolutely continuous with respect to $\lambda \otimes F_Z$, therefore

$$P((Y, Z) \in A_{\varepsilon}) < \rho \tag{2.2}$$

for all ε and for all δ small enough, say for all $\delta < \delta_0(\rho)$. Let

$$B = B(\varepsilon, \delta) = \left\{ (y, z) \colon \int |g(y + \varepsilon z - \varepsilon u) - g(y)| F_Z(du) > 4\delta \right\},$$

$$C = C(\varepsilon, \delta) = \left\{ (y, z) \colon \int |h(y + \varepsilon z - \varepsilon u) - h(y)| F_Z(du) > 2\delta \right\}.$$

To prove (2.1) it is enough to show that $P((Y, Z) \in B) < 3\rho$ as $\varepsilon \to 0$. Clearly

$$P((Y, Z) \in B) \leq P((Y, Z) \in A_0) + P((Y, Z) \in A_{\varepsilon}) + F_Y \otimes F_Z(C).$$

To end the proof it is enough to show that for every fixed $\delta > 0$,

$$F_Y \otimes F_Z(C) \to 0 \quad \text{as } \varepsilon \to 0.$$
 (2.3)

Notice that since $h(\cdot)$ is uniformly continuous, there is an $\eta = \eta(\delta)$ such that $|h(y + \varepsilon z - \varepsilon u) - h(y)| < \delta$ except when either $|z| > \eta/|\varepsilon|$ or $|u| > \eta/|\varepsilon|$. In particular, if $(y, z) \in C$ then either $|z| > \eta/|\varepsilon|$ or

$$(y, z) \in D = \left\{ (y, z) \colon \int_{|u| > \eta/|\varepsilon|} |h(y + \varepsilon z - \varepsilon u) - h(y)|F_Z(\mathrm{d} u) > \delta \right\}.$$

Therefore,

$$F_Y \otimes F_Z(C) \leq P(|Z| > \eta / |\varepsilon|) + F_Y \otimes F_Z(D)$$

Clearly, $P(|Z| > \eta / |\varepsilon|) \to 0$ as $\varepsilon \to 0$. The second term tends to zero, too. Indeed, $F_Y \otimes F_Z$ is absolutely continuous with respect to $\lambda \otimes F_Z$ and by Chebyshev's inequality,

$$\lambda \otimes F_Z(D) \leq \frac{1}{\delta} \int_{\mathbb{R}} \int_{\mathbb{R}} \int_{\|u\| > \eta/|\varepsilon|} \left[h(y + \varepsilon z - \varepsilon u) + h(y) \right] F_Z(du) \lambda(dy) F_Z(dz)$$
$$= \frac{2}{\delta} P(|Z| > \eta/|\varepsilon|) \to 0 \quad \text{as } \varepsilon \to 0.$$

This establishes (2.3) and ends the proof of Theorem 1.1. \Box

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Lemma 2.1. Suppose $\mathscr{F}_{\varepsilon} \subset \mathscr{F}$ are σ -fields. Assume for every $\delta > 0$ there is an $\varepsilon_0(\delta) > 0$ and an $A \in \mathscr{F}$ with $P(A) > 1 - \delta$ such that for all $0 < |\varepsilon| < \varepsilon_0(\delta)$,

$$\{S: S = B \cap A, B \in \mathscr{F}_{\varepsilon}\} \supset \{S: S = C \cap A, C \in \mathscr{F}\}.$$
(2.4)

Then $E\{X \mid \mathscr{F}_{\varepsilon}\} \to X$ in L_1 as $\varepsilon \to 0$ for each integrable X.

Remark. Obviously, (2.4) implies the equality of sets.

Proof of Lemma 2.1. By the standard approximation argument it is enough to consider $X = I_C$, the indicator function of an event $C \in \mathcal{F}$. Clearly,

$$E\{|E(I_C | \mathscr{F}_{\varepsilon}) - I_C|\} \leq E\{|E\{I_{C \cap A} | \mathscr{F}_{\varepsilon}\} - I_{C \cap A}|\} + 2P(A^{c}).$$

By (2.4) we have $E\{I_{C \cap A} \mid \mathscr{F}_{\varepsilon}\} - I_{C \cap A} = I_B\{E(I_A \mid \mathscr{F}_{\varepsilon}) - I_A\}$, for some $B \in \mathscr{F}_{\varepsilon}$. Therefore

$$E\left\{\left|E\left\{I_{C\cap A} \mid \mathscr{F}_{\varepsilon}\right\} - I_{C\cap A}\right|\right\} \leq 2E\left\{\left|I_{A} - P(A)\right|\right\} = 4P(A)P(A^{c}) < 4\delta.$$

This shows that $|| E\{I_C | \mathscr{F}_{\varepsilon}\} - I_C ||_1 \leq 6\delta$ for all $|\varepsilon| < \varepsilon_0(\delta)$. \Box

Proof of Theorem 1.2. Given $\delta > 0$ pick a measurable set $A \subset \Omega$ with $P(A) > 1 - \delta$ and such that the following two conditions are satisfied:

(i) YI_A has finitely many values;

(ii) ZI_A is bounded.

Put $\mathscr{F}_{\varepsilon} = \sigma(Y + \varepsilon Z)$. Clearly, for all $\varepsilon \neq 0$ such that $|\varepsilon|$ is small enough we have

$$\sigma(YI_A + \varepsilon ZI_A) = \sigma(YI_A, ZI_A). \tag{2.5}$$

This shows that condition (2.4) holds with $\mathscr{F}_{\varepsilon}$ given above and $\mathscr{F} = \sigma(Y, Z)$. Indeed, let $C \in \mathscr{F}$, i.e., $C = \{\omega: (Y, Z) \in \mathscr{V}\}$ for some Borel set $\mathscr{V} \subset \mathbb{R}^2$. Then

$$C \cap A = \{ \omega \colon (YI_A, ZI_A) \in \mathscr{V} \} \cap A.$$

$$(2.6)$$

However, $\{\omega: (YI_A, ZI_A) \in \mathscr{V}\} = \{\omega: YI_A + \varepsilon ZI_A \in \mathscr{U}_{\varepsilon}\}$ for some Borel set $\mathscr{U}_{\varepsilon} \subset \mathbb{R}$, provided $\varepsilon \neq 0$ and $|\varepsilon|$ is small enough so that (2.5) holds. Therefore by (2.6),

$$C \cap A = A \cap \{\omega \colon YI_A + \varepsilon ZI_A \in \mathscr{U}_{\varepsilon}\} = A \cap \{\omega \colon Y + \varepsilon Z \in \mathscr{U}_{\varepsilon}\} = A \cap B$$

for some $B \in \mathscr{F}_{\epsilon}$. By Lemma 2.1 the proof is complete. \Box

Proof of Proposition 1.3. Suppose $a \in \mathbb{R}$ is such that P(Y = a) > 0. Let $C = \{\omega: Y = a, Z = a\}$, $A_{\varepsilon} = \{\omega: Y + \varepsilon Z = (1 + \varepsilon)a\}$. Clearly, P(C) > 0 and $A_{\varepsilon} = C$ almost surely except for at most a countable number of ε 's. In particular, $P(A_{\varepsilon}) = P(C) > 0$.

Let $X = \phi(Z)$ for some ϕ . By assumption $E\{\phi(Z) | Y + \varepsilon Z\} \rightarrow E\{\phi(Z)\}$ in probability. On A_{ε} we have

$$E\{\phi(Z) \mid Y + \varepsilon Z\} = \frac{1}{P(A_{\varepsilon})} \int_{A_{\varepsilon}} \phi(Z) \, \mathrm{d}P = \frac{1}{P(C)} \int_{C} \phi(Z) \, \mathrm{d}P = \phi(a).$$

Therefore $E\{\phi(Z) | Y + \varepsilon Z\} = \phi(a)$ on a set of probability P(C) > 0. Since for each $\delta > 0$ we have

$$P\{|E\{\phi(Z)|Y+\varepsilon Z\}-E\{\phi(Z)\}|>\delta\} < P(C)$$

for all $\varepsilon \neq 0$ with $|\varepsilon|$ small enough, therefore $\phi(a) = E\{\phi(Z)\}$. The last equality, however, cannot be true for all functions ϕ unless Z = a with probability one. This shows that if Y is not continuous, then it is deterministic and ends the proof. \Box

References

- Szablowski, P.J. (1990), Expansions of $E\{X | Y + eZ\}$ and their applications to the analysis of elliptically contoured measures, *Comput. Math. Appl.* **19**, 75–83.
- Szablowski, P.J., (1986), On distributive relations involving conditional moments and the probability distributions of the conditioning random variable, *Demonstratio Math.* 19, 721-746.